

Status of copper rockfish (*Sebastes caurinus*) in U.S. waters off  
the coast of Washington in 2021 using catch and length data

by  
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# 1 Introduction

## 1.1 Basic Information

This assessment reports the status of copper rockfish (*Sebastes caurinus*) off the Washington coast using data through 2020.

Copper rockfish is a medium- to large-sized nearshore rockfish found from Mexico to Alaska. The core range is comparatively large, from northern Baja Mexico to the Gulf of Alaska, as well as in Puget Sound. Copper rockfish have historically been a part of both commercial and recreational fisheries throughout its range.

Copper rockfish are commonly found in waters less than 130 meters in depth in nearshore kelp forests and rocky habitat (Love 1996). The diets of copper rockfish consist primarily of crustaceans, mollusks, and fish (Lea et al. 1999; Bizzarro et al. 2017). The body coloring of copper rockfish varies across the West Coast with northern fish often exhibiting dark brown to olive with southern fish exhibiting yellow to olive-pink variations in color (Miller and Lea 1972), which initially led to them being designated as two separate species (*S. caurinus* and *S. vexillaris*).

Numerous genetic studies have been performed looking for genetic variation in copper rockfish, with variable outcomes. Genetic work has revealed significant differences between Puget Sound and coastal stocks (Dick et al. 2014). Stocks along the West Coast have not been determined to be genetically distinct populations, but significant population subdivision has been detected, indicating limited oceanographic exchange among geographically proximate locations (Buonaccorsi et al. 2002; Johansson et al. 2008). A specific study examining copper rockfish populations off the coast of Santa Barbara and Monterey California identified a genetic break between the north and south, with moderate differentiation (Sivasundar and Palumbi 2010).

Copper rockfish are a relatively long-lived rockfish, estimated to live at least 50 years (Love 1996). Copper rockfish was determined to have the highest vulnerability ( $V = 2.27$ ) of any West Coast groundfish stock evaluated in a productivity susceptibility analysis (Cope et al. 2011). This analysis calculated species-specific vulnerability scores based on two dimensions: productivity characterized by the life history and susceptibility that characterized how the stock could be impacted by fisheries and other activities.

## 1.2 Historical and Current Fishery Information

Off the coast of Washington State copper rockfish is primarily caught in the recreational/sport fishery with very little mortality from commercial fishing (Table 1 and Figure 1). Copper

rockfish has been a target of recreational fishing starting as early as 1935, with catches stabilizing around 2,500 - 3,000 fish per year starting around 1980, with the exception of select years with high (2005) or low catches (2015).

Copper rockfish has not been targeted by commercial fisheries in Washington waters. Washington banned commercial fixed gears in 1995 and trawl gear in 1999 in state waters, which encompass the vast majority of depths preferred by copper rockfish. In response to the development of the live-fish fishery in California and Oregon, Washington took preemptive action in 1999 to prevent the fishery from developing by prohibiting the landing of live-fish. There are four treaty tribes that fish under separate rules and are not subject to the state water closure. Copper rockfish are usually landed in the Nearshore Rockfish group, a mixed-species market category. Species composition samples were taken from sampled landings, and proportions of copper rockfish reported in the Nearshore market category are estimated by port, quarter, gear, and year. In 2020, COVID-19 closures of tribal lands prevented samplers from accessing all commercial catch, so an average proportion of copper rockfish in 2017 - 2019 was applied to all associated tribal landings.

The primary region of recreational fishing off the Washington coast occurs in the central to northern regions. These areas have rocky habitat within which rockfish species such as copper rockfish are associated, whereas to the southern coast of Washington consists primarily of soft and sandy substrate. The stock off the Washington coast was assessed as a separate stock (Figure 2) from other populations off the West Coast based on three factors: 1) suspected limited movement of adult fish between Washington and Oregon given the preferred substrate separation, 2) the different exploitation patterns within Washington waters compared to Oregon and California, and 3) the quantity of length data in Washington compared to other areas.

Analysis that summarized current research to inform stock structure in copper rockfish off the West Coast and evaluated the available information to guide the selection of the management area relative to the assessment area is presented in Wetzal et al. (2021).

### **1.3 Summary of Management History and Performance**

Copper rockfish is managed by the Pacific Fishery Management Council (PFMC) as a part of the Nearshore Rockfish North and Nearshore Rockfish South complexes, split at 40° 10' N. lat. off the West Coast. Each complex, comprised of nearshore rockfish species, is managed based on a complex level overfishing limit (OFL) and annual catch limit (ACL) that are determined by summing the species-specific OFLs and ACLs (ACLs set equal to the Acceptable Biological Catches) contributions for all stocks managed in the complex (North or South). Removals for species within the Nearshore Rockfish North and South complexes are managed and tracked against the complex total OFL and ACL, rather than on a species by species basis.

Table 3 show the Nearshore Rockfish North complex level OFLs and ACLs, the copper rockfish OFL and ACL contribution amounts, the state-specific allocations (26 percent for Washington, Groundfish Management Team, personal communication) applied to the copper rockfish ACL contribution, and the total removals in Washington.

## **2 Data**

A description of each data source is provided below (Figure 3).

### **2.1 Fishery-Dependent Data**

#### **2.1.1 Commercial Data**

There are very limited commercial fishery removals of copper rockfish off the Washington coast (Table 1 and Figure 1). Across all model years there were less than 2 mt removed by the commercial fishery. The commercial catches were provided directly by Washington Department of Fish and Wildlife (WDFW). Given the limited observed removals by the commercial fleet, the historical discards (discards prior to 2002) were assumed to be nominal and were not accounted for in the model. In recent years, 2002-2019, the coastwide discards observed by the West Coast Groundfish Observer Program (WCGOP) were attributed to each state based on the proportion of commercial removals by state. The commercial discards allocated to Washington were relatively minor (less than 0.02 mt total since 2002). There were no length data available for copper rockfish in Pacific Fisheries Information Network (PacFIN) for use in this assessment.

#### **2.1.2 Recreational / Sport Data**

Recreational removals in the model begin in 1935 and are the primary source of fishing mortality for copper rockfish (Figure 1). Removals were specified in numbers of fish (1,000s) and were converted to metric tons internally to the model (Table 2). Annual catches (in numbers) from the recreational fishery (1967, 1975-1980) were obtained from historical reports, and landings from 1990-2020 were obtained from WDFW Ocean Sampling Program (OSP) and Puget Sound Baseline Sampling Program (Puget Sound copper rockfish samples not included in the data). To fill in the missing years, linear interpolations were used to find landed values between 1968 and 1974, and to bring catch down to zero in year 1934. Discard estimates are not available prior to 2002. Historical discards were estimated based on a discard to retained catch relationship from 2002 - 2020. Discard mortality by depth was applied to post-2001 discards estimate. Prior to 2002, a 31 percent mortality rate is applied

to all discarded fish. The sum of retained and dead released copper rockfish made up the total removal (in numbers) from the recreational fishery. The recreational removals generally increased over time, spiked in 2005 to an all-time high, and since have been roughly between 1,000 and 3,000 fish per year.

Length compositions for the recreational fleet were available in 1979, 1981, 1982, 1983, and then each year from 1995 - 2020 (Table 4). The number of length observations by year were quite variable ranging between 3 - 463 samples per year (Figure 4). The size of sexed and unsexed fish observed by the recreational fleet were primarily between 30 - 45 cm (Figure 5). The mean length observed by year had limited variation with year with a marginally smaller mean length between 35 - 40 cm until 2010 after which the mean length increased slightly to range around 40 cm (Figure 6).

## 2.2 Fishery-Independent Data

There were no fishery-independent data sources that are commonly incorporated in West Coast groundfish stock assessments, as required by the Terms of Reference for Stock Synthesis catch and length (SS-CL) assessments, available for copper rockfish off the Washington coast.

## 2.3 Biological Data

### 2.3.1 Natural Mortality

The current method for developing a prior on natural mortality for West Coast groundfish stock assessments is based on Hamel (2015), a method for combining meta-analytic approaches relating the  $M$  rate to other life-history parameters such as longevity, size, growth rate, and reproductive effort to provide a prior for  $M$ . This approach modifies work done by Then et al. (2015) who estimated  $M$  and related life history parameters across a large number of fish species from which to develop an  $M$  estimator for fish species in general. They concluded by recommending  $M$  estimates be based on maximum age alone, based on an updated Hoenig non-linear least squares estimator  $M = 4.899A_{\max}^{-0.916}$ . Hamel (personal communication) re-evaluated the data used by Then et al. (2015) by fitting the one-parameter  $A_{\max}$  model under a log-log transformation (such that the slope is forced to be -1 in the transformed space (Hamel 2015), the point estimate and median of the prior for  $M$  is:

$$M = \frac{5.4}{A_{\max}}$$

where  $A_{\max}$  is the maximum age. The prior is defined as a lognormal distribution with mean  $\ln(5.4/A_{\max})$  and standard error = 0.438. Using a maximum age of 50, the point estimate



and median of the prior is  $0.108 \text{ yr}^{-1}$ . The maximum age was selected based on available age data from all West Coast data sources and literature values. The oldest aged copper rockfish was 51 years with two observations, one each off of the coast of Washington and Oregon in 2019. The maximum age in the model was set at 50 years. This selection was consistent with the literature examining the longevity of copper rockfish (Love 1996) and was supported by the observed ages that had multiple observations of fish between 44 and 51 years of age.

### 2.3.2 Length-Weight Relationship

The length-weight relationship for copper rockfish was estimated outside the model using all coastwide biological data available from fishery-independent data from the NWFSC West Coast Groundfish Bottom Trawl Survey (WCGBTS) and the NWFSC Hook and Line survey (Figure 7). The estimated length-weight relationship for female fish was  $W = 9.56e-06L^{3.19}$  and males  $1.08e-05L^{3.15}$  where  $L$  is length in cm and  $W$  is weight in kilograms (Figure 8).

### 2.3.3 Growth (Length-at-Age)

Length-at-age was estimated for male and female copper rockfish using data collected from fishery-dependent data sources off the coasts of Oregon and Washington, collected between 1998-2019 (Table 5). The available fishery-dependent data from Oregon and Washington included limited observations of young fish (less than 4 years of age), which presented challenges for estimating growth. Attempting to estimate growth in the absence of data to inform the rate of growth ( $k$ ) and the size-at-age 0 ( $t_0$ ) could result in biased estimates of all parameters including the size-at-maximum length ( $L_\infty$ ). A published growth study for copper rockfish by Lea et al. (1999) had numerous observations of young fish and also reported the mean length, the number of observations, and the standard deviation of the length observations by age. These pieces of information were used to simulate length-at-age data that would be representative of the study's data for fish younger than 5 years of age. The simulated data for young fish appeared consistent with the data for older fish observed off the Oregon and Washington coast (Figure 9). This combined data set was used to estimate growth curves for male and female copper rockfish that were used in this assessment. Ideally, growth would be estimated using data collected from similar sources. However, the bias from using data from different sources was considered to be less than the bias that may arise from estimating growth from observations that did not cover the range of ages.

The estimated growth used in this assessment had females reach marginally larger asymptotic sizes compared to males. Sex-specific growth parameters were estimated at the following values:

$$\text{Females } L_\infty = 48.4 \text{ cm; } k = 0.206$$

$$\text{Males } L_\infty = 47.2 \text{ cm; } k = 0.231$$

These values were fixed within the base model for male and female copper rockfish. While the growth differences between sexes was limited for copper rockfish, sex-specific parameterization was used in the hopes that it would allow the length data to be the most informative within the assessment. The coefficient of variation (CV) around young and old fish was fixed at a value of 0.10 for both sexes, a value that was based on values observed across other groundfish stocks. The length-at-age curve with the CV around length-at-age by sex is shown in Figure 10.

In contrast to the current approach, the length-at-age values cited in the 2013 data-moderate assessment (Cope et al. 2013) for copper rockfish (although not directly used by the data-moderate model) were from Lea et al. (1999). The  $L_{\infty}$  from Lea et al. (1999) by sex were quite a bit larger than those estimated for this assessment using recent length and age data off the coast of Oregon and Washington. In the Lea et al. (1999) young fish were well sampled, however, there were very few observations of fish older than 12 years of age (fewer than 5 total), which appears to have led to a poorly informed estimate of  $L_{\infty}$ .

For the sake of parsimony, the length-age samples were pooled across sources to estimate a single length-at-age curve for copper rockfish in California north of Point Conception, Oregon, and Washington. In the future, if adequate area-based length-age samples across a range of fishery-dependent and -independent source are available, copper rockfish growth should be re-evaluated for possible area-specific variation.

#### 2.3.4 Maturation and Fecundity

Maturity-at-length is based upon the work of Hannah (2014) who estimated the 50 percent size-at-maturity of 34.8 cm and slope of -0.6 for copper rockfish off the coast of Oregon, with maturity reaching the asymptote of 1.0 for larger fish (Figure 11).

The fecundity-at-length was based on research from Dick et al. (2017). The fecundity relationship for copper rockfish was estimated equal to  $3.362e-07L^{3.68}$  in millions of eggs where  $L$  is length in cm. Fecundity-at-length is shown in Figure 12.

Table 6 shows the length-at-age, weight-at-age, maturity-at-age, and spawning output (the product of fecundity and maturity) assumed in the base model.

#### 2.3.5 Sex Ratio

There were limited sex-specific observations by length or age across biological data sources. The sex ratio of copper rockfish by length and age across all available data sources off the

West Coast are shown in Figures 13 and 14. The sex ratio of young fish was assumed to be 1:1.

## 3 Assessment Model

### 3.1 Summary of Previous Assessments

Copper rockfish was last assessed in 2013 (Cope et al. 2013). The stock was assessed using extended depletion-based stock reduction analysis (XDB-SRA), a data-moderate approach that incorporated catch and index data with priors on select parameters: natural mortality, stock status in a specified year, productivity, and the relative status of maximum productivity. Copper rockfish was assessed as two separated stocks, the area south of Point Conception off the California coast and the area north of Point Conception to the Washington/Canadian border. The 2013 assessment estimated the stock south of Point Conception at 75 percent of unfished spawning biomass and the stock north of Point Conception at 48 percent of unfished spawning biomass.

#### 3.1.1 Bridging Analysis

A direct bridging analysis was not conducted because the previous assessment was structured to include the area from north of Point Conception to the Washington/Canadian border. The data types used in the 2013 assessment were catches and indices of abundance. Matching the 2013 data was not straightforward aside from the challenges already posed from the alternative model platform (XDB-SRA) and area grouping. First, the 2013 assessment document did not report catches on a state and source level (not atypical for grouped state or area assessments). Secondly, some of the recreational indices used in 2013 were calculated based on multi-state data. All of these items created significant challenges of how to conduct an effective, logical, and informative bridging analysis for the assessment north of Point Conception.

### 3.2 Model Structure and Assumptions

The assessment of copper rockfish in Washington waters was assessed using a two-sex model with sex specific life history parameters. The model assumed two fleets: 1) recreational and 2) commercial fleets with recreational removals starting in 1935. Selectivity was specified using the double normal parameterization within Stock Synthesis for the recreational fleet where selectivity was fixed to be asymptotic with the ascending slope and size of maximum selectivity parameters estimated. The commercial fleet selectivity was assumed to be the same as the recreational fleet due to a lack of length data to estimate a fleet-specific selectivity curve. The commercial and recreational fleets were kept separate in the model despite the limited commercial removals and composition data for two reasons: 1) clarity and 2) the recreational fleet removals were specified in terms of numbers while the commercial removals were in biomass. Recruitment was specified to be deterministic due to limited composition data.

### 3.2.1 Modeling Platform and Structure

The assessment was conducted using Stock Synthesis version 3.30.16 developed by Dr. Richard Methot at the NOAA, NWFSC (Methot and Wetzel 2013). This most recent version was used because it included improvements and corrections to older model versions. The R package r4ss, version 1.38.0, along with R version 4.0.1 were used to investigate and plot model fits.

### 3.2.2 Priors

Priors were used to determine fixed parameter values for natural mortality and steepness in the base model. The prior distribution for natural mortality was based on the Hamel (2015) meta-analytic approach with an assumed maximum age of 50 years. The prior assumed a log normal distribution for natural mortality. The log normal prior has a median of  $0.108 \text{ yr}^{-1}$  and a standard error of 0.438.

The prior for steepness assumed a beta distribution with mean of 0.72 and standard error of 0.15. The prior parameters are based on the Thorson-Dorn rockfish prior (commonly used in past West Coast rockfish assessments) conducted by James Thorson (personal communication, NWFSC, NOAA), which was reviewed and endorsed by the Scientific and Statistical Committee (SSC) in 2017. However, this approach was subsequently rejected for future analysis in 2019 when the new meta-analysis resulted in a mean value of approximately 0.95. In the absence of a new method for generating a prior for steepness the default approach reverts to the previously endorsed method, the 2017 value.

### 3.2.3 Data Weighting

Length compositions from the recreational fleet were the only composition data fit in the model. In the absence of index or commercial composition data, no data weighting was performed in the base model. Sensitivities were performed using the three data weighting approaches that are commonly applied for West Coast groundfish stock assessments: Francis method (Francis 2011), McAllister and Ianelli method, known as Harmonic Mean weighting (McAllister and Ianelli 1997), and the Dirichlet-Multinomial method (Thorson et al. 2017).

### 3.2.4 Estimated and Fixed Parameters

There were 3 estimated parameters in the base model. These included one parameter for  $R_0$  and 2 parameters for recreational selectivity (Table 7). Selectivity in the recreational fleet was fixed asymptotic with only the peak and the ascending slope estimated. Dome-shaped

selectivity was explored for the recreational fleet. Older and larger copper rockfish may be found deeper waters and may move into areas that limit their availability to fishing gear. However, limited support for dome-shaped selectivity for the recreational fleet was found and the selectivity was fixed to be asymptotic. The commercial selectivity was set equal to the recreational selectivity due to a lack of composition data to support fleet specific estimation.

Fixed parameters in the model were as follows. Steepness was fixed at 0.72, the mean of the prior. Natural mortality was fixed at  $0.108 \text{ yr}^{-1}$  for females and males, the median of the prior. The standard deviation of recruitment deviates was fixed at 0 and recruitment was assumed deterministic. Maturity-at-length was fixed as described above in Section 2.3.4. Length-weight parameters were fixed at estimates using all length-weight observations described above in Section 2.3.2. The length-at-age was fixed at sex-specific externally estimated values described above in Section 2.3.3.

### 3.3 Model Selection and Evaluation

The base assessment model for copper rockfish was developed to balance parsimony and realism, and the goal was to estimate a spawning output trajectory for the population of copper rockfish off the Washington coast. Several investigative model runs were done to achieve the final base model.

### 3.4 Base Model Results

The base model parameter estimates, along with approximate asymptotic standard errors, are shown in Table 7 and the likelihood components are shown in Table 8. Estimates of stock size and status over time are shown in Table 9.

#### 3.4.1 Parameter Estimates

Estimated parameter values are provided in Table 7. The model estimated 3 total parameters:  $R_0$  and two parameters associated with the recreational fleet selectivity. The  $R_0$  was estimated at 2.03. The selectivity curve was estimated for the recreational fleet (Figure 15). The selectivity was fixed to be asymptotic, reaching maximum selectivity for fish at 37 cm. The selectivity for the commercial fleet was assumed to be equal to the recreational fleet selectivity due to the lack of commercial length data.

### 3.4.2 Fits to the Data

Fits to the length data are shown based on the Pearson residuals-at-length, the annual mean lengths, and aggregated length composition data for the recreational fleet. The Pearson residuals show a pattern of greater observations of all sexes (unsexed, males, and females) that start around 1995 and appear to move through the length data by year, possibly indicating a strong or multiple strong recruitments entering the population (Figure 16). Estimating annual recruitment deviations would have allowed the model to fit the patterns in the length data. However, the base model did not estimate annual recruitment deviations due to limited length data, which resulted in extreme recruitment deviation estimates (large positive deviations in the 1990s followed by string of negative deviations from 2000 - 2020) resulting in a highly pessimistic stock status (see Section 3.5.2). The assessment of copper rockfish off the Oregon coast, which also did not include annual recruitment deviations had similar indications in the data of one or more strong year classes in the mid-1990s indicating that oceanographic forces driving recruitment success or failure may be shared across Oregon and Washington. The mean lengths across years with data was relatively stable ranging roughly between 35 and 42 cm by year (Figure 17).

Detailed fits to the length data by year are provided in the Appendix, Section 9. Aggregate fits by fleet are shown in Figure 18. There are a few things that stand out when examining the aggregated length composition data. The estimated fits to the data by sex matches the general shape of the aggregated lengths. However, the model expects a higher proportion of the largest fish relative to the data. This may indicate that the true selectivity of the recreational fleet may have some level of reduced selectivity of the largest fish (dome-shaped) but selectivity in the model was fixed to be asymptotic to follow the pre-specified terms of reference for a length-based assessment which specifies that at least one fleet in the model should have asymptotic selectivity. Sensitivities examining dome-shaped and estimating annual recruitment deviations were performed and presented in the Sensitivity Analyses section below.

### 3.4.3 Population Trajectory

The predicted spawning output (in millions of eggs) is given in Table 9 and plotted in Figure 19. The predicted spawning output from the base model generally showed a slow decline over the time series with the spawning output stabilizing in recent years. The total biomass shows a similar slow decline across the modeled years (Figure 20).

The 2020 spawning output relative to unfished equilibrium spawning output is above the target of 40 percent of unfished spawning output (0.42, Figure 21). Approximate confidence intervals based on the asymptotic variance estimates show that the uncertainty in the estimated spawning output is limited. The standard deviation of the log of the spawning output in 2020 is 0.1.

The stock-recruit curve resulting from a value of steepness fixed at is shown in Figure 22. The estimated annual recruitment is shown in Figure 23

## 3.5 Model Diagnostics

### 3.5.1 Convergence

Proper convergence was determined by starting the minimization process from dispersed values of the maximum likelihood estimates to determine if the model found a better minimum. Starting parameters were jittered by 10 percent. This was repeated 100 times, with 94 out of 100 runs returned to the base model likelihood. A better, lower negative log-likelihood, model fit was not found. The model did not experience convergence issues when provided reasonable starting values. Through the jittering done as explained and likelihood profiles, we are confident that the base model as presented represents the best fit to the data given the assumptions made. There were no difficulties in inverting the Hessian to obtain estimates of variability, although much of the early model investigation was done without attempting to estimate a Hessian.

### 3.5.2 Sensitivity Analyses

Several sensitivity analyses were conducted. The majority of the sensitivities conducted was a single exploration from the base model assumptions and/or data, and were not performed in a cumulative fashion.

1. Estimate female natural mortality ( $M$ ).
2. Estimate female growth coefficient ( $k$ ).
3. Estimate female  $L_\infty$ .
4. Estimate annual recruitment deviations.
5. Estimate annual recruitment deviations while allowing for a dome-shaped selectivity curve.
6. Data weighting according to the McAllister-Ianelli (MI DW) method using the weighting values shown in Table 10.
7. Data weighting according to the Francis method using the weighting values shown in Table 10.
8. Data weighting according to the Dirichlet-Multinomial (DM DW) method where the estimated parameters are shown in Table 10.



Likelihood values and estimates of key parameters from each sensitivity are available in Table 11. Plots of the estimated time-series of spawning biomass and relative spawning biomass are shown in Figures 24 and 25.

The two sensitivities that had the largest impact relative to the base model were those that estimated annual recruitment deviations alone or combined to also estimate dome-shaped selectivity (Figures 24 and 25). The time series of each of these sensitivities resulted in large scale swings in spawning output across time. The estimates of recruitment deviations had a string of average estimates during the late 1990s and early 2000s then switched to a period of low recruitment success in recent years (Figure 26). The recruitment deviation estimates were likely being driven by subtle shifts in the mean lengths being observed across time (Figure 6). The pattern of recruitment deviations estimated was extreme for the Washington area model but the general pattern showed some coherence with the recruitment deviations peaks estimated in the separate Oregon area model (shown as a sensitivity). The sensitivity in the Oregon model that estimated recruitment deviations also estimated above average recruitment in the late 1990s which led to an increase in spawning output during the 2000s similar to what was observed in this sensitivity for the Washington model. This may indicate that copper rockfish off the coast of Oregon and Washington experience similar drivers in recruitment.

A sensitivity run allowing dome-shaped selectivity with deterministic recruitment was also performed but not shown due to the model failing to estimate a reasonable biomass scale (hit the upper bound of  $\log(R_0)$  of 20).

The sensitivities exploring data weighting using the Francis or MI methods matched the base model as expected given that there was only one source of data used in the model. In contrast, the DM method resulted in a stock size and status less than the base model. It is unclear why this difference in results across data weighting arose.

### 3.5.3 Likelihood Profiles

Likelihood profiles were constructed for  $R_0$ , steepness, female natural mortality, female  $L_\infty$ , female growth coefficient ( $k$ ), and female coefficient of variation for older fish ( $CV_2$ ) values separately. These likelihood profiles were constructed by fixing the parameter at specific values and estimating the remaining parameters based on the fixed parameter value.

In regards to values of  $R_0$ , the negative log-likelihood was minimized at approximately  $\log(R_0)$  of 2.03 (Figure 27). Increasing the  $R_0$ , relative to the value estimated, results in an increase in stock scale and status (Figures 28 and 29).

For steepness, the negative log-likelihood supported values at the upper bound of 1.0 (Figure 30). Assuming higher or lower steepness values impacted the estimated unfished spawning

output but had less impact on the final spawning output (Figure 31). The estimated relative final stock status ranged between around the minimum threshold (0.25) to above the management target depending upon assuming a lower or higher steepness value (Figure 32).

The negative log-likelihood profile across female natural mortality supported values greater than the fixed value of  $0.108 \text{ yr}^{-1}$  (Figure 33). The estimate stock trajectories assuming lower or higher natural mortality values impacted the estimated unfished spawning output and resulted in stock statuses within the management precautionary zone (between 0.25 - 0.40) and above (Figures 34 and 35).

A profile across a range of female  $L_{\infty}$  values was also conducted (Figure 36). The negative log-likelihood showed support for values between 46 and 47 cm. The  $L_{\infty}$  value for female fish in the model was fixed at 48.43 cm based on length-at-age data collected off the Oregon and Washington coasts. The stock scale and status is quite variable across alternative  $L_{\infty}$  values where assuming lower values resulted in sharp increases in stock scale and status (Figures 37 and 38).

A profile across a range of female  $k$  values was also conducted (Figure 39). The negative log-likelihood showed support for values between 0.11 and  $0.13 \text{ yr}^{-1}$ . The  $k$  value for female fish in the model was fixed at  $0.206 \text{ yr}^{-1}$  based on length-at-age data collected off the Oregon and Washington coasts. The stock scale and status increases under lower  $k$  values where assuming higher values resulted in decreases in stock scale and status (Figures 40 and 41).

The profile across a range of  $CV_2$  for older females supported lower  $CV_2$  values (Figure 42). Assuming lower  $CV_2$  values increased the estimated spawning output but had limited impact in the estimate of fraction unfished (Figures 43 and 44).

### 3.5.4 Length-Based Spawner-per-Recruit Analysis

An exploratory length-based spawner-per-recruit analysis using the approach developed by Hordyk et al. (2015) was conducted. This approach assumes asymptotic selectivity and deterministic recruitment to produce independent estimates by year of selectivity and spawner-per-recruit (SPR) effort based on the observed recreational lengths. This analysis indicated that copper rockfish were 50 percent selected generally between 30 - 35 cm (excluding 2018) with full selection between 35 - 40 cm (Figure 45). The median estimates of SPR by year ranged between 0.60 - 0.75 between 2016 - 2019 with an average of the medians of 0.67. This type of analysis can provide insight on the fishing effort based on life history and observed length data in the absence of an integrated assessment model.

An additional analysis to estimate stock status based on length data alone was conducted within a length only version of Stock Synthesis. Within this approach the removal history is

removed but the same life history values, selectivities, and length compositions (both sexed and unsexed) are used. The underlying assumption is that the population has gone through an aggregate constant catch and constant recruitment to get an estimate of the current stock status. Length compositions are fit by estimating the parameter  $\log(R_0)$  (considered a nuisance parameter), which allows for best fits to the length comps and the selectivity by fleet. Using the recreational lengths, the estimated logistic selectivity, and life history from the Washington base model the implied stock status in 2020 was estimated to be approximately 0.44.

The estimates of the SPR harvest rate by year and the length only version of Stock Synthesis were used to provide external estimates of stock status in 2020 for two Simple Stock Synthesis (SSS) analysis.

### 3.5.5 Simple Stock Synthesis

A SSS was run to compare the results from the base model with a simpler modeling approach. SSS samples via Monte Carlo from three key parameter distributions: natural mortality, steepness, and stock status in a specific year. The mean and median of the priors assumed in the base model were used to create sampling distributions for natural mortality and steepness. Two alternative assumptions regarding the distribution of current stock status were explored. SSS applies parameter draws from each of the three parameters within the model and then estimates an  $R_0$  value based on the fixed removals and drawn parameters.

1. Current stock status based on LB-SPR estimates:
  - Number of draws = 1,000,
  - $M = \text{lognormal } (\mu = 0.108 \text{ yr}^{-1}, \sigma = 0.22)$ ,
  - $h = \text{truncated beta } (\alpha = 0.72, \beta = 0.15, a = 0.20, b = 1.0)$ , and
  - Fraction unfished in 2020 =  $\text{beta}(\alpha = 0.67, \beta = 0.20)$
2. Current stock status based on the estimate from length only Stock Synthesis:
  - Number of draws = 1,000,
  - $M = \text{lognormal distribution } (\mu = 0.108 \text{ yr}^{-1}, \sigma = 0.22)$ ,
  - $h = \text{truncated beta } (\alpha = 0.72, \beta = 0.15, a = 0.20, b = 1.0)$ , and
  - Fraction unfished in 2020 =  $\text{beta}(\alpha = 0.44, \beta = 0.20)$

The median of unfished spawning output, spawning output 2021, fraction unfished in 2021, the OFL in 2023, and the ABC in 2023 based on the 2021 fraction unfished of 67 percent is shown in in Table 12. The prior distribution for the parameters and the derived quantities with 95 percent uncertainty are shown in Figures 46 and 47. Assuming that the stock was less depleted relative to the base model resulted in higher estimates of the OFL and ABC in

2023, even when the category 3 buffer was applied (buffer = 0.778, based on a  $P^* = 0.45$  and  $\$4 = 2.0$ ).

The median of unfished spawning output, spawning output 2021, fraction unfished in 2021, the OFL in 2023, and the ABC in 2023 based on the 2021 fraction unfished of 44 percent is shown in Table 13. The prior distribution for the parameters and the derived quantities with 95 percent uncertainty are shown in Figures 48 and 49. Assuming a stock status similar to the base model, SSS resulted in an OFL and ABC values that were marginally lower due to the larger buffer applied in the SSS model (base model OFL in 2023 = 2.15, ABC in 2023 = 1.88).

### **3.5.6 Retrospective Analysis**

A five-year retrospective analysis was conducted by running the model using data only through 2015, 2016, 2017, 2018, 2019 and 2020. The estimated spawning output was consistent with the base model when recent data were removed up but estimated a lower stock size and status relative to the base model when the last 3-5 years of data were removed (Figures 50 and 51).

### **3.5.7 Comparison with Other West Coast Stocks**

Copper rockfish is assessed as four distinct stocks off the U.S. west coast: south of Point Conception in California; north of Point Conception in California; Oregon; and Washington. The area north of Point Conception off the coast of California was estimated to have the largest unfished spawning output of copper rockfish off the West Coast. The stocks off of the Oregon and Washington coast are smaller in size compared to the California stocks, with the stock off the coast of Washington estimated to have the smallest unfished spawning output. Comparison of the estimated spawning output trajectories for the California stocks are shown in Figure 52 with Oregon and Washington shown in Figure 53. The fraction unfished across all West Coast stocks shown in Figure 54. The California stocks are estimated to be the most depleted, with the stock south of Point Conception estimated below the management threshold of 25 percent of unfished and the stock north of Point Conception estimated to be in the precautionary zone (less than the management target of 40 percent but above the management threshold). The stock off the coast of Washington is estimated to be just above the management target and the Oregon stock well above the target.

## 4 Management

### 4.1 Reference Points

Reference points were calculated using the estimated selectivity and catch distributions among fleets in the most recent year of the model (2020, Table 14). The estimated sustainable total yields are estimated to be 2.24 mt when using an  $\text{SPR}_{50\%}$  reference harvest rate. The spawning output equivalent to 40 percent of unfished ( $\text{SB}_{40\%}$ ) was estimated to be 3.06 million eggs.

The spawning output relative to unfished equilibrium in 2021 is estimated to be above the management target of 40 percent of unfished spawning output (42 percent, Figure 21). The fishing intensity,  $1 - \text{SPR}$ , has been above and below the harvest rate limit ( $\text{SPR}_{50\%}$ ) in recent years (Table 9 and Figure 55). The base model estimates that the stock has not fallen below the biomass target and fishing intensity in 2020 is below the target (Figure 56). Table 14 shows the full suite of estimated reference points for the base model and Figure 57 shows the equilibrium curve based on a steepness value fixed at 0.72.

### 4.2 Harvest Projections and Decision Tables

A ten year projection of the base model with catches equal to the estimated Acceptable Biological Catch (ABC) based on the category 2 time-varying  $\sigma$  and  $P^* = 0.45$  for years 2023-2032 with all removals allocated to the recreational fleet (Table 15). The removals in 2021 and 2022 were set based on the adopted Annual Catch Limits (ACLs) for copper rockfish N. 40° 10' Lat. N. allocated to Washington state (26 percent, PFMC Groundfish Management Team, personal communication).

The axes of uncertainty in the decision table is based on the uncertainty around the spawning biomass in 2021 via the  $\log(R_0)$  parameter. The within model uncertainty estimated from the model was low (0.098) which resulted in little variation in the low and high state of nature relative to the base model if used. Meanwhile, the default category 2  $\sigma$  of 1.0 resulted in a very large range across the states of nature. As an alternative approach to determine a low and high state of nature from the base model the model uncertainties across the three other area based copper rockfish assessments were averaged to arise at a  $\sigma = 0.35$ . of This  $\sigma$  value was used to identify the 12.5 and 87.5 percentiles of the asymptotic standard deviation for the current year, 2021, spawning biomass from the base model to identify the low and high states of nature (i.e., 1.15 standard deviations corresponding to the 12.5 and 87.5 percentiles). Once the 2021 spawning biomass for the low and high states of nature were identified a search across  $\log(R_0)$  values were done to attain the current year spawning biomass values. The  $\log(R_0)$  values that corresponded with the lower and upper percentiles were 1.89 and 2.21.

Across the low and high states of nature and across alternative future harvest scenarios the fraction of unfished ranges between 0.36 - 0.58 by the end of the 10 year projection period (Table 16). The fraction unfished under the low state of nature declines below that management target by the end of the projection period.

### 4.3 Evaluation of Scientific Uncertainty

The estimated uncertainty in the base model around the 2021 spawning output is  $\sigma = 0.1$  and the uncertainty in the base model around the 2021 OFL is  $\sigma = 0.09$ . The estimated model uncertainty was less than the category 2 groundfish data-moderate assessment default value of  $\sigma = 1.0$ .

### 4.4 Research and Data Needs

The ability to estimate additional process and biological parameters for copper rockfish was limited by data. Collecting the following data would be beneficial to future assessments of the stock:

1. Continue collecting length and otolith samples from recreational catches with an emphasis on increasing annual samples collected.
2. Collect length and age data for the commercial fleet if catches occur.
3. The sub-stocks of copper rockfish off the coast of Washington and Oregon were assessed separately here. However, research on the connectivity of copper rockfish stocks off the coast of Oregon and Washington would be beneficial to inform future assessment area decisions.

## 5 Acknowledgments

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Melissa Head provided an area specific maturity estimate for copper rockfish and provided insight in the complex biological processes that govern maturity processes.

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## 7 Tables

**Table 1:** Catches (mt) by fleet for all years and total catches (mt) by year summed by year.

Year	WA Recreational	WA Commercial	Total Catch
1935	0.02	0.00	0.02
1936	0.05	0.00	0.05
1937	0.09	0.00	0.09
1938	0.12	0.00	0.12
1939	0.15	0.00	0.15
1940	0.19	0.00	0.19
1941	0.22	0.00	0.22
1942	0.26	0.00	0.26
1943	0.29	0.00	0.29
1944	0.33	0.00	0.33
1945	0.36	0.00	0.36
1946	0.39	0.00	0.39
1947	0.43	0.00	0.43
1948	0.46	0.00	0.46
1949	0.50	0.00	0.50
1950	0.53	0.00	0.53
1951	0.56	0.00	0.56
1952	0.60	0.00	0.60
1953	0.63	0.00	0.63
1954	0.67	0.00	0.67
1955	0.70	0.00	0.70
1956	0.73	0.00	0.73
1957	0.76	0.00	0.76
1958	0.80	0.00	0.80
1959	0.83	0.00	0.83
1960	0.87	0.00	0.87
1961	0.90	0.00	0.90
1962	0.93	0.00	0.93
1963	0.96	0.00	0.96
1964	1.00	0.00	1.00
1965	1.03	0.00	1.03
1966	1.06	0.00	1.06
1967	1.09	0.00	1.09
1968	1.12	0.00	1.12
1969	1.16	0.00	1.16
1970	1.19	0.00	1.19
1971	1.22	0.00	1.22
1972	1.25	0.00	1.25
1973	1.29	0.00	1.29
1974	1.32	0.00	1.32

**Table 1:** Catches (mt) by fleet for all years and total catches (mt) by year summed by year.  
(*continued*)

Year	WA Recreational	WA Commercial	Total Catch
1975	1.35	0.00	1.35
1976	0.97	0.00	0.97
1977	0.60	0.00	0.60
1978	1.11	0.00	1.11
1979	1.48	0.00	1.48
1980	0.87	0.00	0.87
1981	1.94	0.00	1.94
1982	2.03	0.00	2.03
1983	1.24	0.00	1.24
1984	1.97	0.00	1.97
1985	1.70	0.20	1.89
1986	2.04	0.19	2.23
1987	2.45	0.93	3.39
1988	2.29	0.25	2.53
1989	2.32	0.00	2.32
1990	2.96	0.03	2.98
1991	2.17	0.00	2.17
1992	3.51	0.00	3.51
1993	2.74	0.01	2.75
1994	1.91	0.00	1.91
1995	2.46	0.00	2.46
1996	2.85	0.00	2.85
1997	2.70	0.00	2.70
1998	2.76	0.00	2.76
1999	2.80	0.00	2.80
2000	2.92	0.00	2.92
2001	2.95	0.00	2.95
2002	1.90	0.00	1.90
2003	2.25	0.00	2.25
2004	2.21	0.00	2.21
2005	6.19	0.00	6.19
2006	2.87	0.00	2.87
2007	2.89	0.00	2.89
2008	3.03	0.00	3.03
2009	2.72	0.00	2.72
2010	2.13	0.00	2.13
2011	2.64	0.00	2.64
2012	1.76	0.00	1.76
2013	2.56	0.00	2.56
2014	2.34	0.00	2.34
2015	1.32	0.00	1.32
2016	1.86	0.00	1.86

**Table 1:** Catches (mt) by fleet for all years and total catches (mt) by year summed by year.  
*(continued)*

Year	WA Recreational	WA Commercial	Total Catch
2017	1.30	0.01	1.30
2018	3.03	0.00	3.03
2019	4.28	0.00	4.28
2020	1.55	0.00	1.55

**Table 2:** Input numbers of fish removals by year converted to metric tons (mt) within the model.

Year	Numbers of Fish	Model Estimated (mt)
1934	0	0.00
1935	10	0.02
1936	32	0.05
1937	53	0.09
1938	75	0.12
1939	96	0.15
1940	118	0.19
1941	139	0.22
1942	161	0.26
1943	182	0.29
1944	204	0.33
1945	225	0.36
1946	246	0.39
1947	268	0.43
1948	289	0.46
1949	311	0.50
1950	332	0.53
1951	354	0.56
1952	375	0.60
1953	397	0.63
1954	418	0.67
1955	440	0.70
1956	461	0.73
1957	482	0.76
1958	504	0.80
1959	525	0.83
1960	547	0.87
1961	568	0.90
1962	590	0.93
1963	611	0.96
1964	633	1.00
1965	654	1.03
1966	676	1.06
1967	696	1.09
1968	718	1.12
1969	740	1.16
1970	761	1.19
1971	783	1.22
1972	804	1.25
1973	826	1.29
1974	847	1.32

**Table 2:** Input numbers of fish removals by year converted to metric tons (mt) within the model. (*continued*)

Year	Numbers of Fish	Model Estimated (mt)
1975	868	1.35
1976	628	0.97
1977	387	0.60
1978	719	1.11
1979	957	1.48
1980	563	0.87
1981	1253	1.94
1982	1317	2.03
1983	805	1.24
1984	1280	1.97
1985	1105	1.70
1986	1335	2.04
1987	1608	2.45
1988	1506	2.29
1989	1534	2.32
1990	1966	2.96
1991	1449	2.17
1992	2359	3.51
1993	1850	2.74
1994	1296	1.91
1995	1675	2.46
1996	1948	2.85
1997	1853	2.70
1998	1897	2.76
1999	1932	2.80
2000	2027	2.92
2001	2053	2.95
2002	1327	1.90
2003	1573	2.25
2004	1551	2.21
2005	4359	6.19
2006	2038	2.87
2007	2066	2.89
2008	2183	3.03
2009	1972	2.72
2010	1544	2.13
2011	1916	2.64
2012	1277	1.76
2013	1858	2.56
2014	1699	2.34
2015	955	1.32
2016	1339	1.86

**Table 2:** Input numbers of fish removals by year converted to metric tons (mt) within the model. (*continued*)

Year	Numbers of Fish	Model Estimated (mt)
2017	932	1.30
2018	2173	3.03
2019	3073	4.28
2020	1115	1.55

**Table 3:** The OFL (mt) and ACL (mt) for north nearshore complex, the ACL allocated to Washington, and the total removals (mt).

Year	Complex OFL	Complex ACL	OFL - copper	ACL - copper	WA ACL	WA Removals
2011	-	-	28.61	23.88	6.21	2.64
2012	-	-	28.61	23.88	6.21	1.76
2013	-	-	25.96	21.65	5.63	2.56
2014	-	-	25.96	21.65	5.63	2.34
2015	-	69	10.64	9.71	2.52	1.32
2016	-	69	10.33	9.43	2.45	1.86
2017	118.39	105	11.24	10.26	2.67	1.30
2018	118.6	105	11.59	10.58	2.75	3.03
2019	91	81	11.91	10.88	2.83	4.28
2020	92	82	12.24	11.18	2.91	1.55



**Table 4:** Summary of the recreational length samples used in the stock assessment.

Year	All Fish	Sexed Fish	Unsexed Fish
1979	8	0	8
1981	4	0	4
1982	5	0	5
1983	3	0	3
1995	141	0	141
1996	221	0	221
1997	63	0	63
1998	202	46	156
1999	194	136	58
2000	26	26	0
2001	32	32	0
2002	83	61	22
2003	46	18	28
2004	244	201	43
2005	443	265	178
2006	169	96	73
2007	152	110	42
2008	91	71	20
2009	71	52	19
2010	57	38	19
2011	127	27	100
2012	81	37	44
2013	71	14	57
2014	136	130	6
2015	84	81	3
2016	179	155	24
2017	212	108	104
2018	315	188	127
2019	463	273	190
2020	59	58	1

**Table 5:** Summary of the number of samples by year from commercial (Com.) and recreational (Rec.) fisheries by state used to estimate length-at-age parameters.

	OR Com.	OR Rec.	WA Com.	WA Rec.
1998	0	0	0	46
1999	0	0	0	136
2000	0	0	0	26
2001	0	0	0	32
2002	1	0	0	19
2003	9	0	0	0
2004	26	0	0	188
2005	0	58	0	225
2006	1	150	0	65
2007	1	188	0	86
2008	1	217	0	65
2009	0	156	0	35
2010	6	273	0	24
2011	0	235	0	27
2012	11	216	0	35
2013	31	158	0	8
2014	25	121	0	123
2015	10	0	0	74
2016	25	0	0	169
2017	40	177	1	101
2018	44	175	0	176
2019	102	174	0	274

**Table 6:** Age, length, weight, maturity, and spawning output by age (product of maturity and fecundity) at the start of the year for female fish.

Age	Length (cm)	Weight (kg)	Maturity	Spawning Output
0	4.00	0.00	0.00	0.00
1	13.46	0.04	0.00	0.00
2	19.97	0.14	0.00	0.00
3	25.27	0.30	0.01	0.00
4	29.58	0.49	0.11	0.01
5	33.09	0.70	0.35	0.06
6	35.95	0.91	0.60	0.13
7	38.27	1.11	0.76	0.20
8	40.16	1.29	0.86	0.26
9	41.70	1.46	0.91	0.31
10	42.95	1.60	0.94	0.35
11	43.97	1.73	0.96	0.38
12	44.80	1.83	0.97	0.41
13	45.48	1.92	0.97	0.44
14	46.03	2.00	0.98	0.46
15	46.47	2.06	0.98	0.48
16	46.84	2.11	0.98	0.49
17	47.13	2.15	0.99	0.50
18	47.38	2.19	0.99	0.51
19	47.57	2.22	0.99	0.52
20	47.73	2.24	0.99	0.53
21	47.86	2.26	0.99	0.53
22	47.97	2.28	0.99	0.54
23	48.05	2.29	0.99	0.54
24	48.12	2.30	0.99	0.54
25	48.18	2.31	0.99	0.54
26	48.23	2.31	0.99	0.55
27	48.26	2.32	1.00	0.55
28	48.30	2.32	1.00	0.55
29	48.32	2.33	1.00	0.55
30	48.34	2.33	1.00	0.55
31	48.36	2.33	1.00	0.55
32	48.37	2.34	1.00	0.55
33	48.38	2.34	1.00	0.55
34	48.39	2.34	1.00	0.55
35	48.40	2.34	1.00	0.55
36	48.40	2.34	1.00	0.55
37	48.41	2.34	1.00	0.55
38	48.41	2.34	1.00	0.55
39	48.42	2.34	1.00	0.55
40	48.42	2.34	1.00	0.55

**Table 6:** Age, length, weight, maturity, and spawning output by age (product of maturity and fecundity) at the start of the year for female fish. (*continued*)

Age	Length (cm)	Weight (kg)	Maturity	Spawning Output
41	48.42	2.34	1.00	0.55
42	48.42	2.34	1.00	0.55
43	48.42	2.34	1.00	0.55
44	48.42	2.34	1.00	0.55
45	48.43	2.34	1.00	0.55
46	48.43	2.34	1.00	0.55
47	48.43	2.34	1.00	0.55
48	48.43	2.34	1.00	0.56
49	48.43	2.34	1.00	0.56
50	48.43	2.34	1.00	0.56

**Table 7:** List of parameters used in the base model, including estimated values and standard deviations (SD), bounds (minimum and maximum), estimation phase (negative values not estimated), status (indicates if parameters are near bounds), and prior type information (mean and SD).

Parameter	Value	Phase	Bounds	Status	SD	Prior (Exp.Val, SD)
NatM p 1 Fem GP 1	0.108	-2	(0.05, 0.4)	NA	NA	Log Norm (-2.2256, 0.48)
L at Amin Fem GP 1	13.460	-2	(3, 25)	NA	NA	None
L at Amax Fem GP 1	48.430	-2	(35, 60)	NA	NA	None
VonBert K Fem GP 1	0.206	-2	(0.03, 0.3)	NA	NA	None
CV young Fem GP 1	0.100	-2	(0.01, 0.3)	NA	NA	None
CV old Fem GP 1	0.100	-2	(0.01, 0.3)	NA	NA	None
Wtlen 1 Fem GP 1	0.000	-9	(0, 0.1)	NA	NA	None
Wtlen 2 Fem GP 1	3.190	-9	(2, 4)	NA	NA	None
Mat50% Fem GP 1	34.830	-9	(10, 60)	NA	NA	None
Mat slope Fem GP 1	-0.600	-9	(-1, 0)	NA	NA	None
Eggs scalar Fem GP 1	0.000	-9	(-3, 3)	NA	NA	None
Eggs exp len Fem GP 1	3.679	-9	(-3, 3)	NA	NA	None
NatM p 1 Mal GP 1	0.108	-2	(0.05, 0.4)	NA	NA	Log Norm (-2.2256, 0.48)
L at Amin Mal GP 1	8.500	-2	(3, 25)	NA	NA	None
L at Amax Mal GP 1	47.240	-2	(35, 60)	NA	NA	None
VonBert K Mal GP 1	0.231	-2	(0.03, 0.3)	NA	NA	None
CV young Mal GP 1	0.100	-2	(0.01, 0.3)	NA	NA	None
CV old Mal GP 1	0.100	-2	(0.01, 0.3)	NA	NA	None
Wtlen 1 Mal GP 1	0.000	-9	(0, 0.1)	NA	NA	None
Wtlen 2 Mal GP 1	3.150	-9	(2, 4)	NA	NA	None
FracFemale GP 1	0.500	-9	(0.01, 0.99)	NA	NA	None
SR LN(R0)	2.033	1	(1, 20)	OK	0.0390705	None
SR BH steep	0.720	-7	(0.22, 1)	NA	NA	Normal (0.72, 0.16)
SR sigmaR	0.900	-99	(0.15, 1)	NA	NA	None
SR regime	0.000	-99	(-2, 2)	NA	NA	None
SR autocorr	0.000	-99	(0, 0)	NA	NA	None
Size DbIN peak WA Recreational(1)	36.950	1	(15, 50)	OK	0.4347850	None
Size DbIN top logit WA Recreational(1)	-0.505	-2	(-7, 7)	NA	NA	None
Size DbIN ascend se WA Recreational(1)	3.653	3	(-10, 10)	OK	0.1059820	None
Size DbIN descend se WA Recreational(1)	-0.413	-4	(-10, 10)	NA	NA	None

**Table 7:** List of parameters used in the base model, including estimated values and standard deviations (SD), bounds (minimum and maximum), estimation phase (negative values not estimated), status (indicates if parameters are near bounds), and prior type information (mean and SD). (*continued*)

Parameter	Value	Phase	Bounds	Status	SD	Prior (Exp.Val, SD)
Size DblN start logit WA Recreational(1)	-20.000	-9	(-20, 30)	NA	NA	None
Size DblN end logit WA Recreational(1)	10.000	-3	(-10, 10)	NA	NA	None

**Table 8:** Likelihood components by source.

Label	Total
TOTAL	1132.78
Catch	0.00
Equil catch	0.00
Length comp	1132.78
Recruitment	0.00
InitEQ Regime	0.00
Forecast Recruitment	0.00
Parm priors	0.00
Parm softbounds	0.00
Parm devs	0.00
Crash Pen	0.00

**Table 9:** Time series of population estimates from the base model.

Year	Total Biomass (mt)	Spawn- ing Output	Total Biomass 3+ (mt)	Frac- tion Un- fished	Age-0 Re- cruits	Total Catch (mt)	1-SPR	Ex- ploita- tion Rate
1935	72.44	7.65	71.60	1.00	7.65	0.02	0.00	0.00
1936	72.43	7.65	71.58	1.00	7.65	0.05	0.01	0.00
1937	72.38	7.64	71.53	1.00	7.65	0.09	0.02	0.00
1938	72.30	7.63	71.45	1.00	7.65	0.12	0.02	0.00
1939	72.20	7.62	71.35	1.00	7.64	0.15	0.03	0.00
1940	72.07	7.61	71.22	0.99	7.64	0.19	0.03	0.00
1941	71.91	7.59	71.06	0.99	7.64	0.22	0.04	0.00
1942	71.73	7.57	70.89	0.99	7.64	0.26	0.05	0.00
1943	71.53	7.54	70.69	0.99	7.64	0.29	0.05	0.00
1944	71.32	7.52	70.47	0.98	7.63	0.33	0.06	0.00
1945	71.08	7.49	70.23	0.98	7.63	0.36	0.06	0.01
1946	70.83	7.46	69.98	0.97	7.63	0.39	0.07	0.01
1947	70.56	7.43	69.71	0.97	7.63	0.43	0.08	0.01
1948	70.28	7.39	69.43	0.97	7.62	0.46	0.08	0.01
1949	69.98	7.36	69.14	0.96	7.62	0.50	0.09	0.01
1950	69.68	7.32	68.83	0.96	7.61	0.53	0.09	0.01
1951	69.36	7.28	68.51	0.95	7.61	0.56	0.10	0.01
1952	69.03	7.24	68.19	0.95	7.61	0.60	0.11	0.01
1953	68.69	7.20	67.85	0.94	7.60	0.63	0.11	0.01
1954	68.35	7.16	67.50	0.94	7.60	0.67	0.12	0.01
1955	67.99	7.12	67.15	0.93	7.59	0.70	0.12	0.01
1956	67.63	7.07	66.79	0.92	7.59	0.73	0.13	0.01
1957	67.27	7.03	66.42	0.92	7.58	0.76	0.14	0.01
1958	66.89	6.98	66.05	0.91	7.58	0.80	0.14	0.01
1959	66.51	6.94	65.67	0.91	7.57	0.83	0.15	0.01
1960	66.13	6.89	65.29	0.90	7.57	0.87	0.15	0.01
1961	65.74	6.85	64.90	0.89	7.56	0.90	0.16	0.01
1962	65.35	6.80	64.51	0.89	7.56	0.93	0.17	0.01
1963	64.95	6.75	64.11	0.88	7.55	0.96	0.17	0.02
1964	64.55	6.70	63.72	0.88	7.55	1.00	0.18	0.02
1965	64.15	6.65	63.31	0.87	7.54	1.03	0.18	0.02
1966	63.74	6.61	62.91	0.86	7.53	1.06	0.19	0.02
1967	63.33	6.56	62.50	0.86	7.53	1.09	0.20	0.02
1968	62.92	6.51	62.09	0.85	7.52	1.12	0.20	0.02
1969	62.51	6.46	61.67	0.84	7.52	1.16	0.21	0.02
1970	62.09	6.41	61.26	0.84	7.51	1.19	0.21	0.02
1971	61.67	6.36	60.84	0.83	7.50	1.22	0.22	0.02
1972	61.25	6.31	60.42	0.82	7.50	1.25	0.23	0.02
1973	60.83	6.26	60.00	0.82	7.49	1.29	0.23	0.02
1974	60.40	6.21	59.57	0.81	7.48	1.32	0.24	0.02



**Table 9:** Time series of population estimates from the base model. (*continued*)

Year	Total Biomass (mt)	Spawn- ing Output	Total Biomass 3+ (mt)	Frac- tion Un- fished	Age-0 Re- cruits	Total Catch (mt)	1-SPR	Ex- ploita- tion Rate
1975	59.98	6.16	59.15	0.80	7.48	1.35	0.24	0.02
1976	59.55	6.10	58.72	0.80	7.47	0.97	0.19	0.02
1977	59.51	6.10	58.68	0.80	7.47	0.60	0.12	0.01
1978	59.85	6.13	59.02	0.80	7.47	1.11	0.21	0.02
1979	59.68	6.11	58.85	0.80	7.47	1.48	0.26	0.03
1980	59.16	6.06	58.33	0.79	7.46	0.87	0.17	0.01
1981	59.26	6.07	58.43	0.79	7.46	1.94	0.32	0.03
1982	58.33	5.96	57.50	0.78	7.45	2.03	0.34	0.04
1983	57.35	5.85	56.53	0.76	7.43	1.24	0.24	0.02
1984	57.19	5.82	56.36	0.76	7.43	1.97	0.34	0.03
1985	56.35	5.73	55.52	0.75	7.41	1.89	0.33	0.03
1986	55.62	5.64	54.80	0.74	7.40	2.23	0.38	0.04
1987	54.62	5.52	53.79	0.72	7.38	3.39	0.50	0.06
1988	52.55	5.29	51.73	0.69	7.34	2.53	0.43	0.05
1989	51.41	5.15	50.60	0.67	7.31	2.32	0.41	0.05
1990	50.56	5.05	49.74	0.66	7.29	2.98	0.48	0.06
1991	49.12	4.88	48.31	0.64	7.26	2.17	0.40	0.04
1992	48.55	4.81	47.74	0.63	7.24	3.51	0.55	0.07
1993	46.73	4.60	45.92	0.60	7.20	2.75	0.49	0.06
1994	45.73	4.48	44.93	0.59	7.17	1.91	0.39	0.04
1995	45.60	4.46	44.81	0.58	7.17	2.46	0.46	0.05
1996	44.98	4.38	44.18	0.57	7.15	2.85	0.51	0.06
1997	44.01	4.27	43.21	0.56	7.12	2.70	0.50	0.06
1998	43.23	4.18	42.43	0.55	7.10	2.76	0.51	0.06
1999	42.43	4.09	41.64	0.53	7.07	2.80	0.52	0.07
2000	41.64	4.00	40.85	0.52	7.04	2.92	0.54	0.07
2001	40.77	3.90	39.98	0.51	7.02	2.95	0.55	0.07
2002	39.91	3.80	39.13	0.50	6.99	1.90	0.43	0.05
2003	40.10	3.82	39.33	0.50	6.99	2.25	0.47	0.06
2004	39.97	3.80	39.20	0.50	6.99	2.21	0.47	0.06
2005	39.88	3.79	39.11	0.50	6.98	6.19	0.76	0.16
2006	35.96	3.37	35.18	0.44	6.84	2.87	0.58	0.08
2007	35.34	3.29	34.57	0.43	6.80	2.89	0.59	0.08
2008	34.76	3.21	34.00	0.42	6.78	3.03	0.61	0.09
2009	34.08	3.13	33.32	0.41	6.74	2.72	0.58	0.08
2010	33.73	3.09	32.97	0.40	6.72	2.13	0.51	0.06
2011	33.97	3.12	33.22	0.41	6.74	2.64	0.57	0.08
2012	33.72	3.09	32.97	0.40	6.72	1.76	0.45	0.05
2013	34.32	3.16	33.57	0.41	6.75	2.56	0.56	0.08
2014	34.13	3.14	33.38	0.41	6.74	2.34	0.53	0.07
2015	34.14	3.14	33.40	0.41	6.75	1.32	0.37	0.04

**Table 9:** Time series of population estimates from the base model. (*continued*)

Year	Total Biomass (mt)	Spawn- ing Output	Total Biomass 3+ (mt)	Frac- tion Un- fished	Age-0 Re- cruits	Total Catch (mt)	1-SPR	Ex- ploita- tion Rate
2016	35.15	3.25	34.40	0.42	6.79	1.86	0.46	0.05
2017	35.60	3.30	34.85	0.43	6.81	1.30	0.36	0.04
2018	36.57	3.41	35.82	0.45	6.85	3.03	0.59	0.08
2019	35.84	3.34	35.09	0.44	6.83	4.28	0.69	0.12
2020	33.90	3.13	33.14	0.41	6.74	1.55	0.42	0.05
2021	34.65	3.20	33.90	0.42	6.77	2.11	0.50	0.06
2022	34.86	3.22	34.11	0.42	6.78	2.10	0.49	0.06
2023	35.09	3.25	34.33	0.42	6.79	1.88	0.46	0.05
2024	35.51	3.29	34.76	0.43	6.81	1.89	0.46	0.05
2025	35.92	3.34	35.17	0.44	6.83	1.89	0.46	0.05
2026	36.31	3.39	35.56	0.44	6.84	1.89	0.45	0.05
2027	36.68	3.43	35.92	0.45	6.86	1.90	0.45	0.05
2028	37.03	3.47	36.27	0.45	6.87	1.90	0.45	0.05
2029	37.37	3.51	36.61	0.46	6.89	1.90	0.45	0.05
2030	37.69	3.54	36.93	0.46	6.90	1.90	0.44	0.05
2031	38.00	3.58	37.24	0.47	6.91	1.89	0.44	0.05
2032	38.30	3.61	37.53	0.47	6.92	1.89	0.44	0.05

**Table 10:** Data weights applied by each alternative data weighting method.

Method	Recreational Lengths
Francis	0.064
McAllister-Ianelli	0.163
Dirichlet Multinomial	0.360

**Table 11:** Sensitivities relative to the base model.[illegible]

**Table 12:** Derived quantities from SSS based on assuming fraction unfished of 67 percent in 2021 .

	Median	Lower Interval	Upper Interval
SSB Unfished	13.95	4.33	64.05
SSB 2021	9.45	1.66	58.96
Fraction Unfished 2021	0.70	0.28	0.93
OFL 2023	5.96	1.31	28.02
ABC 2023	4.65	0.86	21.86

**Table 13:** Derived quantities from SSS based on assuming fraction unfished of 44 percent in 2021 .

	Median	Lower Interval	Upper Interval
SSB Unfished	8.29	3.37	24.55
SSB 2021	3.42	0.71	19.13
Fraction Unfished 2021	0.44	0.13	0.81
OFL 2023	2.31	0.58	10.65
ABC 2023	1.80	0.12	8.31

**Table 14:** Summary of reference points and management quantities, including estimates of the 95 percent intervals.

	Estimate	Lower Interval	Upper Interval
Unfished Spawning Output	7.65	7.065	8.236
Unfished Age 3+ Biomass (mt)	71.596	66.113	77.079
Unfished Recruitment ( $R_0$ , thousands)	7.638	7.053	8.223
Spawning Output (2021 mt)	3.203	2.588	3.818
Fraction Unfished (2021)	0.419	0.37	0.467
Reference Points Based $SB_{40\%}$	-	-	-
Proxy Spawning Output $SB_{40\%}$	3.06	2.826	3.295
SPR Resulting in $SB_{40\%}$	0.458	0.458	0.458
Exploitation Rate Resulting in $SB_{40\%}$	0.072	0.072	0.072
Yield with SPR Based on $SB_{40\%}$ (mt)	2.347	2.175	2.518
Reference Points Based on $SPR_{50}$ for MSY	-	-	-
Proxy Spawning Output ( $SPR_{50}$ )	3.413	3.152	3.675
$SPR_{50}$	0.5	-	-
Exploitation Rate Corresponding to $SPR_{50}$	0.063	0.062	0.063
Yield with $SPR_{50}$ at SB SPR (mt)	2.239	2.075	2.403
Reference Points Based on Estimated MSY Values	-	-	-
Spawning Output at MSY ( $SB_{MSY}$ )	2.092	1.93	2.254
$SPR_{MSY}$	0.344	0.344	0.345
Exploitation Rate Corresponding to $SPR_{MSY}$	0.105	0.104	0.106
MSY (mt)	2.494	2.312	2.675

**Table 15:** Projections of potential OFLs (mt), ABCs (mt), estimated spawning output, and fraction unfished. The OFL, ACL, and Washington (WA) ACL for 2021 and 2022 reflect adopted management limits.

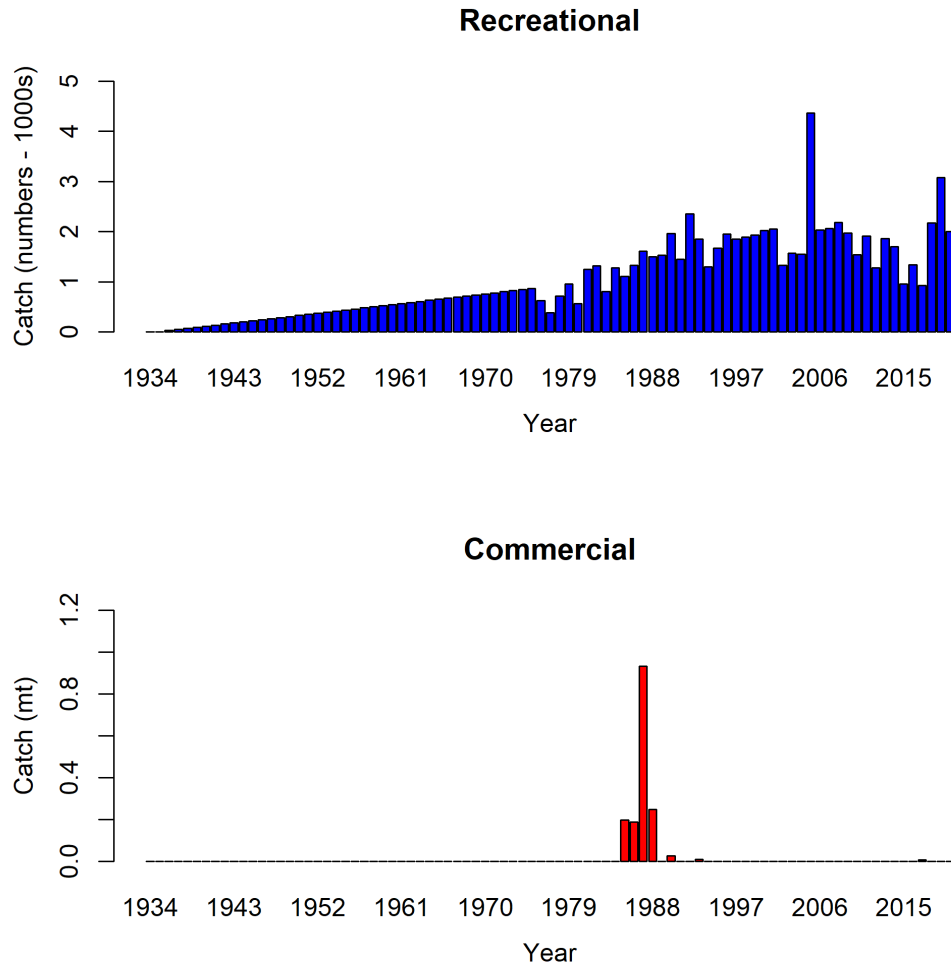
Year	Adopted OFL	Adopted ACL	ACL- WA	OFL	ABC	Buffer	Spaw- ing Output	Frac- tion Un- fished
2021	9.83	8.11	2.11	-	-	-	3.20	0.42
2022	9.86	8.06	2.1	-	-	-	3.22	0.42
2023	-	-	-	2.15	1.88	0.876	3.25	0.42
2024	-	-	-	2.18	1.89	0.867	3.29	0.43
2025	-	-	-	2.2	1.89	0.858	3.34	0.44
2026	-	-	-	2.23	1.89	0.85	3.39	0.44
2027	-	-	-	2.25	1.9	0.842	3.43	0.45
2028	-	-	-	2.28	1.9	0.834	3.47	0.45
2029	-	-	-	2.3	1.9	0.826	3.51	0.46
2030	-	-	-	2.32	1.9	0.818	3.54	0.46
2031	-	-	-	2.34	1.89	0.81	3.58	0.47
2032	-	-	-	2.36	1.89	0.803	3.61	0.47

**Table 16:** Decision table summary of 10 year projections beginning in 2023 for alternative states of nature based on an axis of uncertainty around initial stock size. Columns range over low, mid, and high states of nature and rows range over different catch level assumptions.

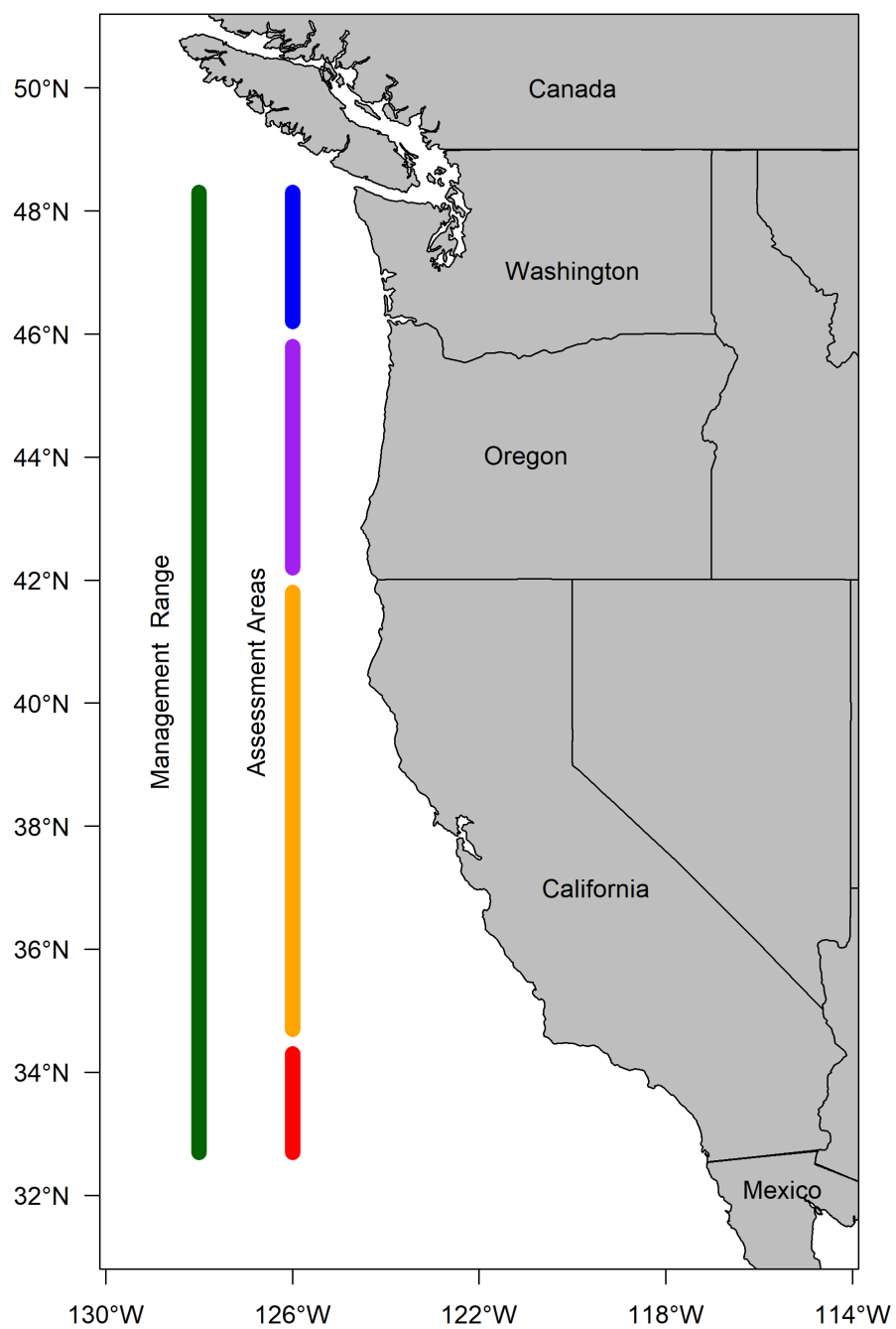
	Year	Catch	$\log(R_0)=1.89$		$\log(R_0)=2.03$		$\log(R_0)=2.21$	
			Spawning Biomass	Fraction Unfished	Spawning Biomass	Fraction Unfished	Spawning Biomass	Fraction Unfished
ACL P*= 0.45	2021	2.11	2.15	0.324	3.20	0.419	4.75	0.519
	2022	2.10	2.16	0.324	3.22	0.421	4.79	0.522
	2023	1.88	2.17	0.326	3.25	0.424	4.82	0.526
	2024	1.89	2.20	0.330	3.29	0.431	4.88	0.533
	2025	1.89	2.23	0.335	3.34	0.437	4.94	0.539
	2026	1.89	2.26	0.340	3.39	0.443	5.00	0.545
	2027	1.90	2.29	0.344	3.43	0.448	5.05	0.551
	2028	1.90	2.32	0.348	3.47	0.454	5.10	0.556
	2029	1.90	2.34	0.352	3.51	0.459	5.14	0.561
	2030	1.90	2.37	0.356	3.54	0.463	5.19	0.566
	2031	1.89	2.39	0.360	3.58	0.468	5.23	0.570
	2032	1.89	2.42	0.364	3.61	0.472	5.27	0.575



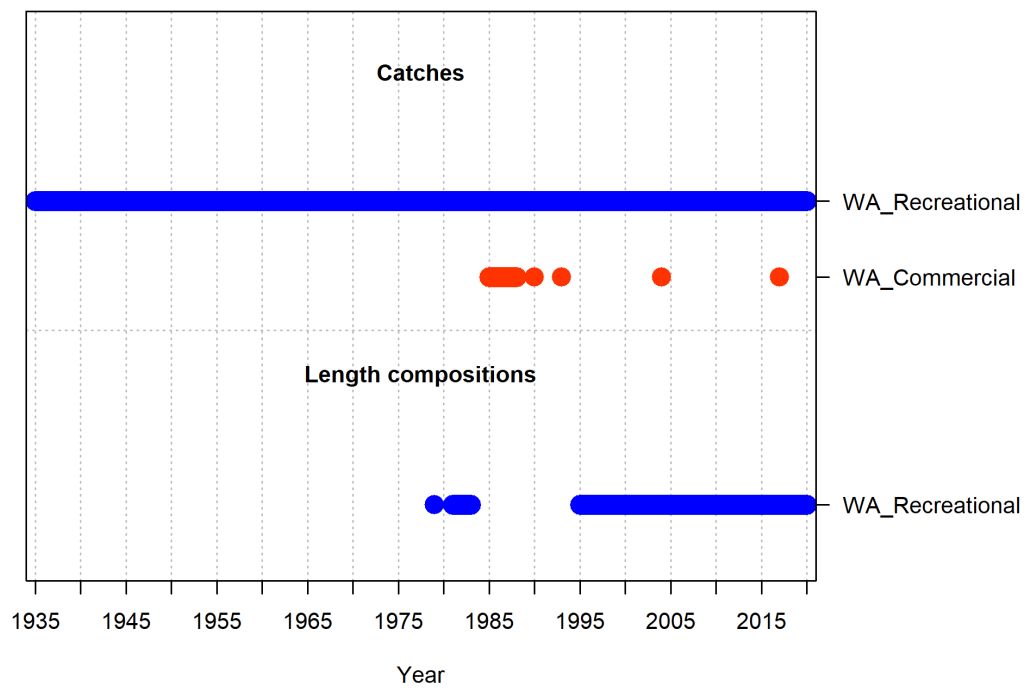
## 8 Figures



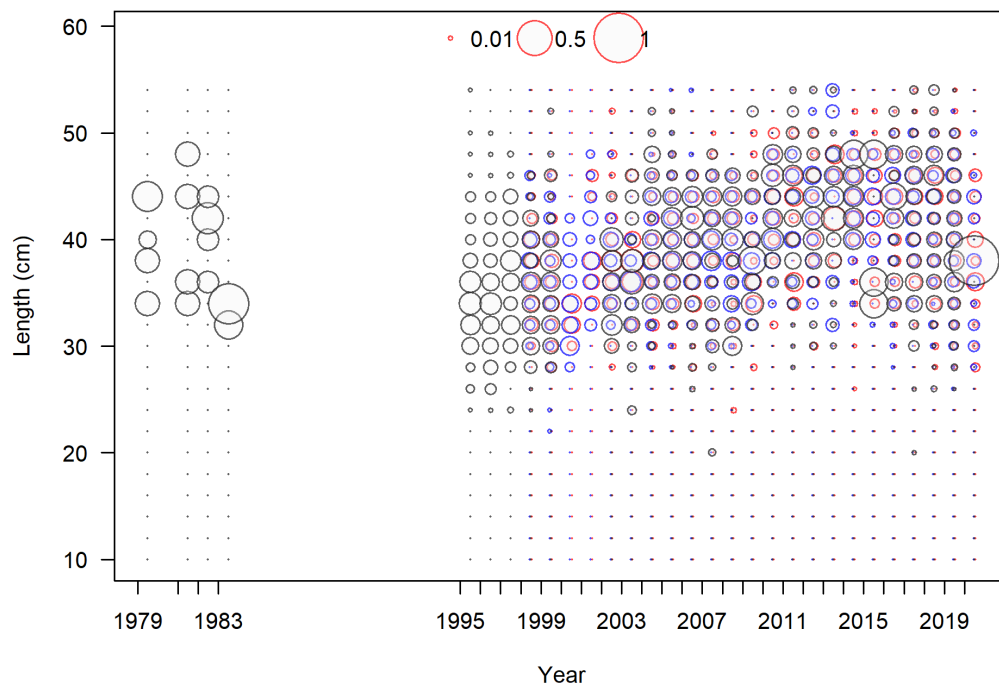
**Figure 1:** Catches by year for the recreational and commercial fleets in the model.



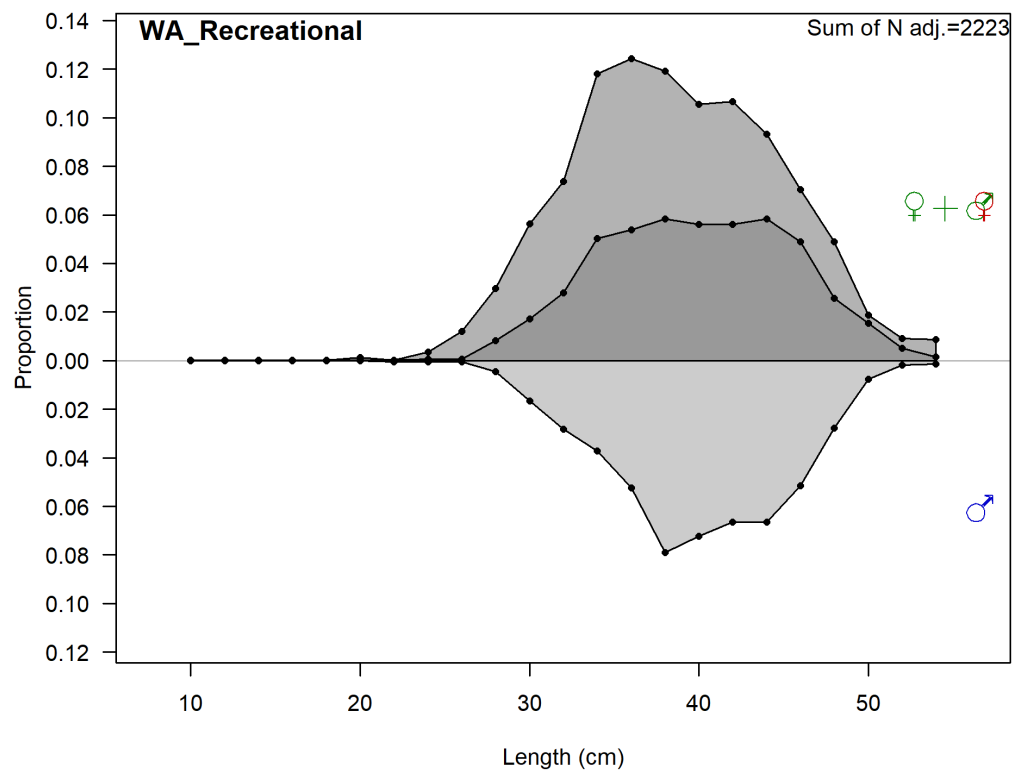
**Figure 2:** Map of management area and assessments areas for copper rockfish with the assessment area for Washington shown in blue.



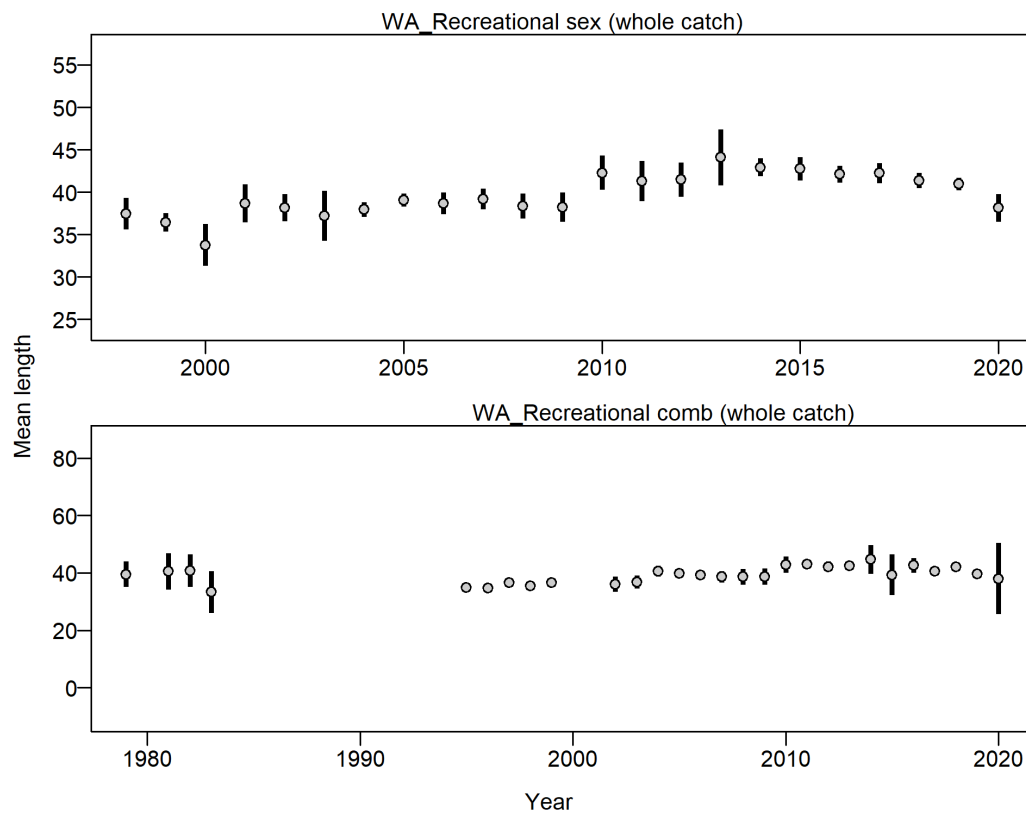
**Figure 3:** Summary of data sources used in the base model.



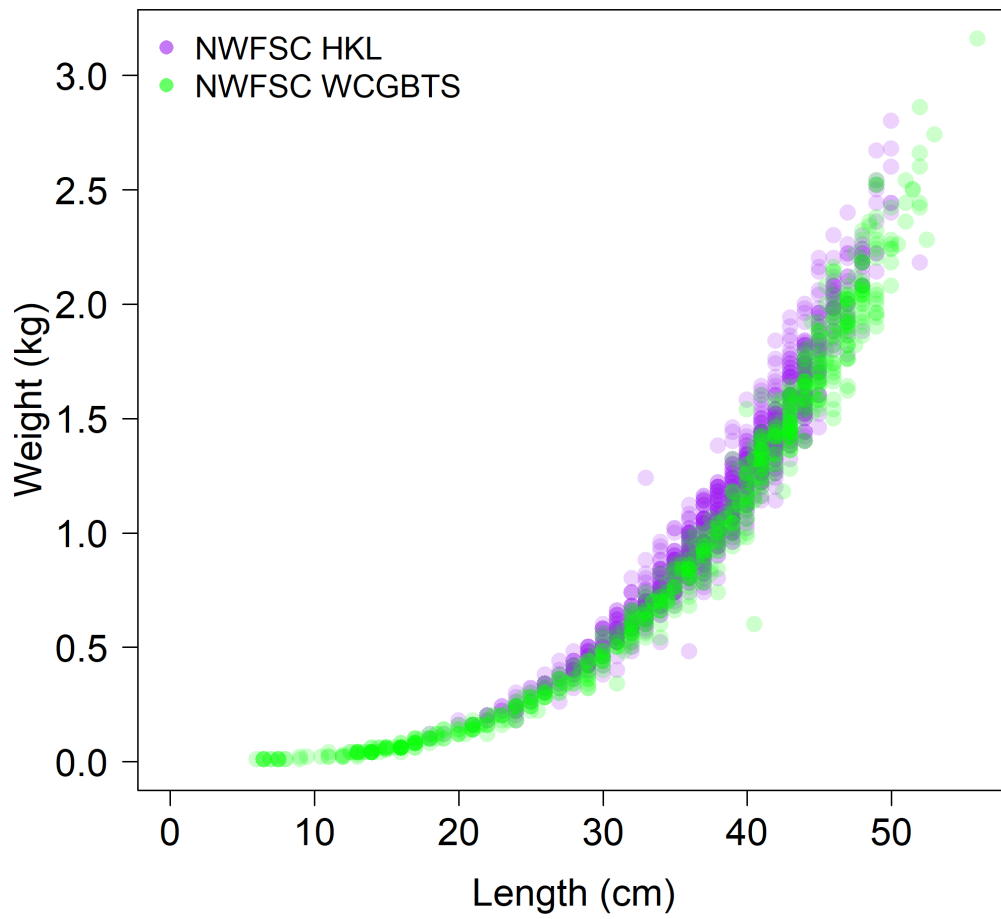
**Figure 4:** Length composition data from the recreational fleet.



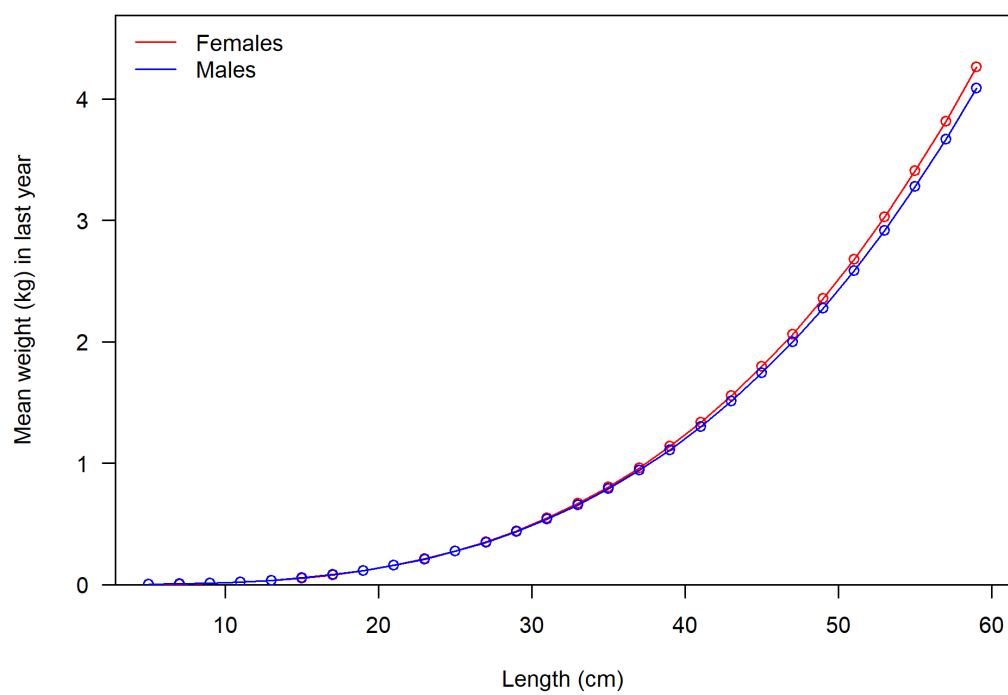
**Figure 5:** Aggregated length composition data across all years from the recreational fleet.



**Figure 6:** Mean length for the recreational fleet with 95 percent confidence intervals.

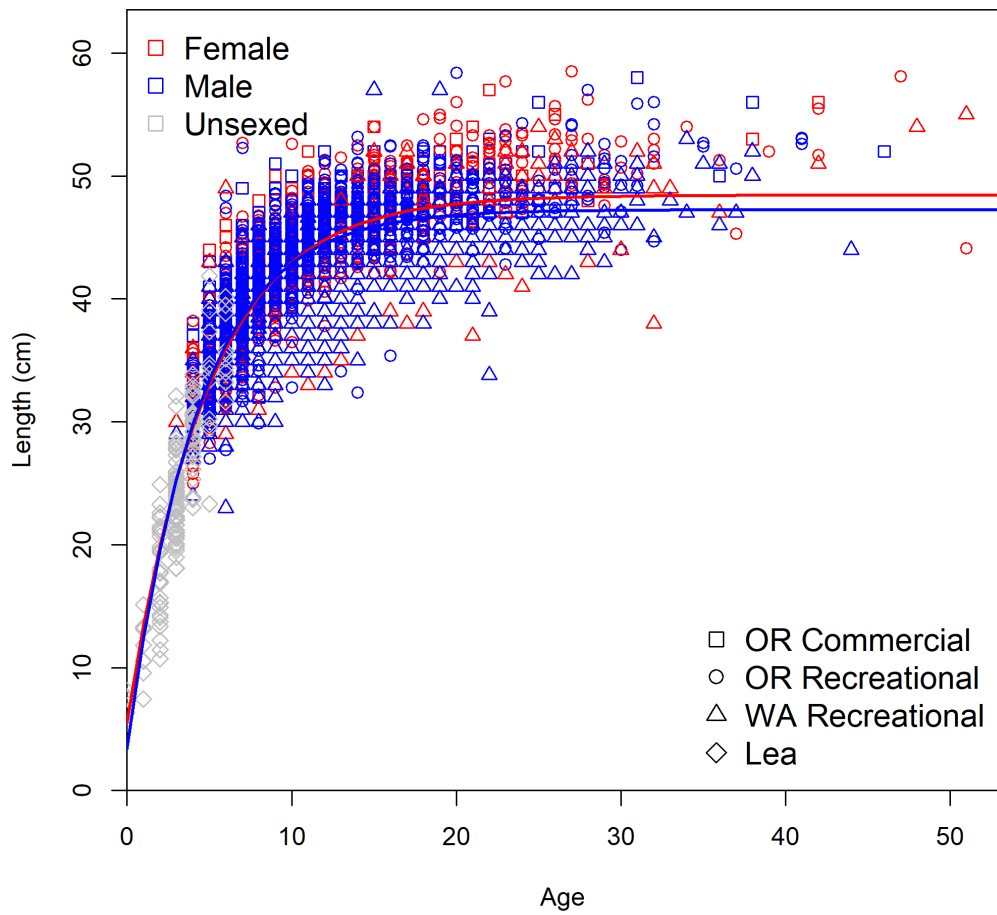


**Figure 7:** Comparison of the length-at-weight data from the NWFSC Hook and Line and the NWFSC WCGBT surveys.

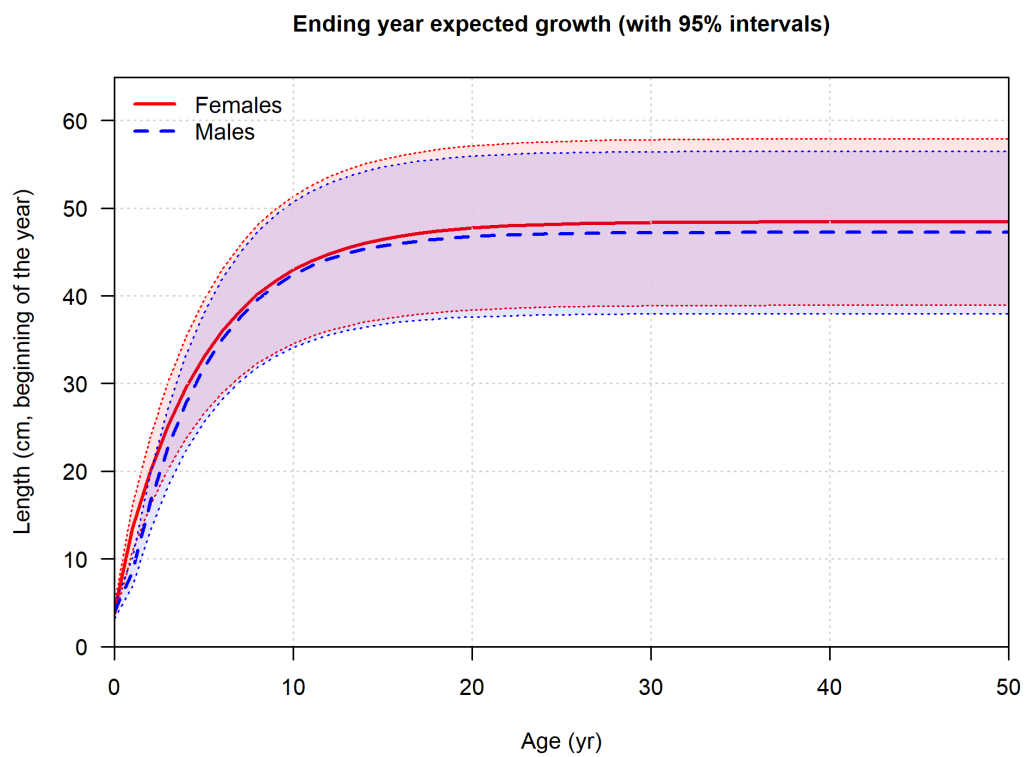


**Figure 8:** Weight-at-length by sex used in the model.

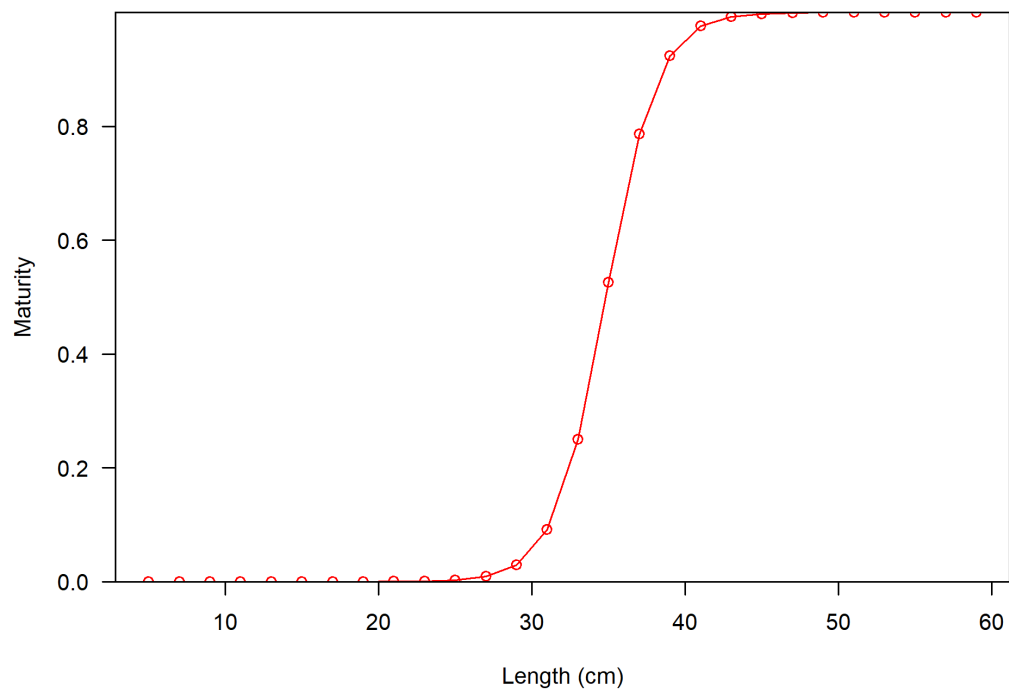




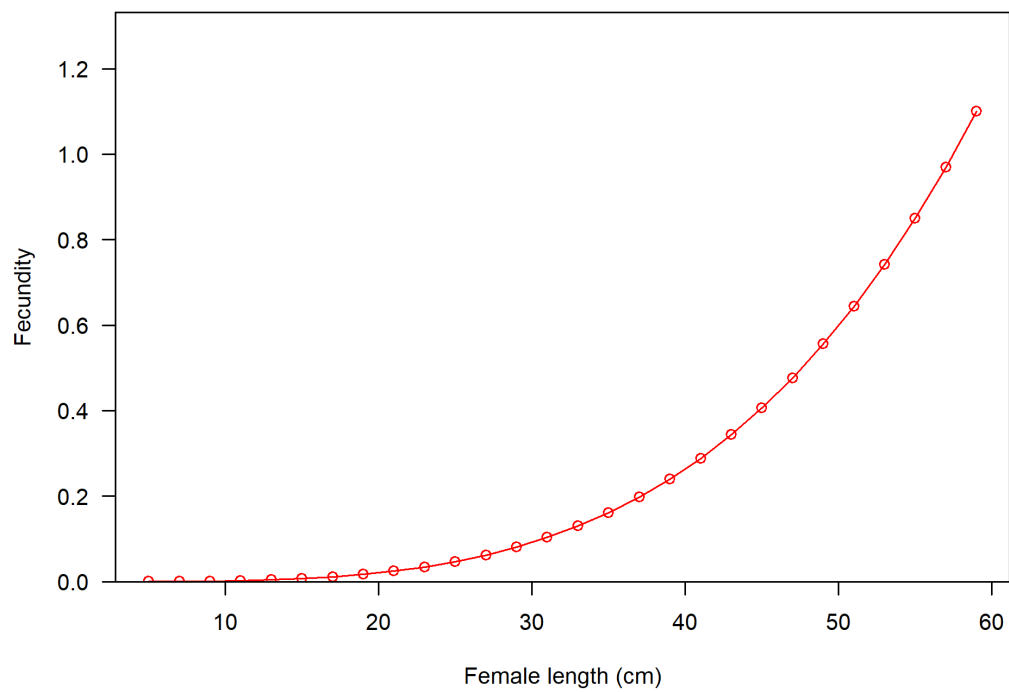
**Figure 9:** Observed sex-specific length-at-age by data source with the estimate length-at-age curve.



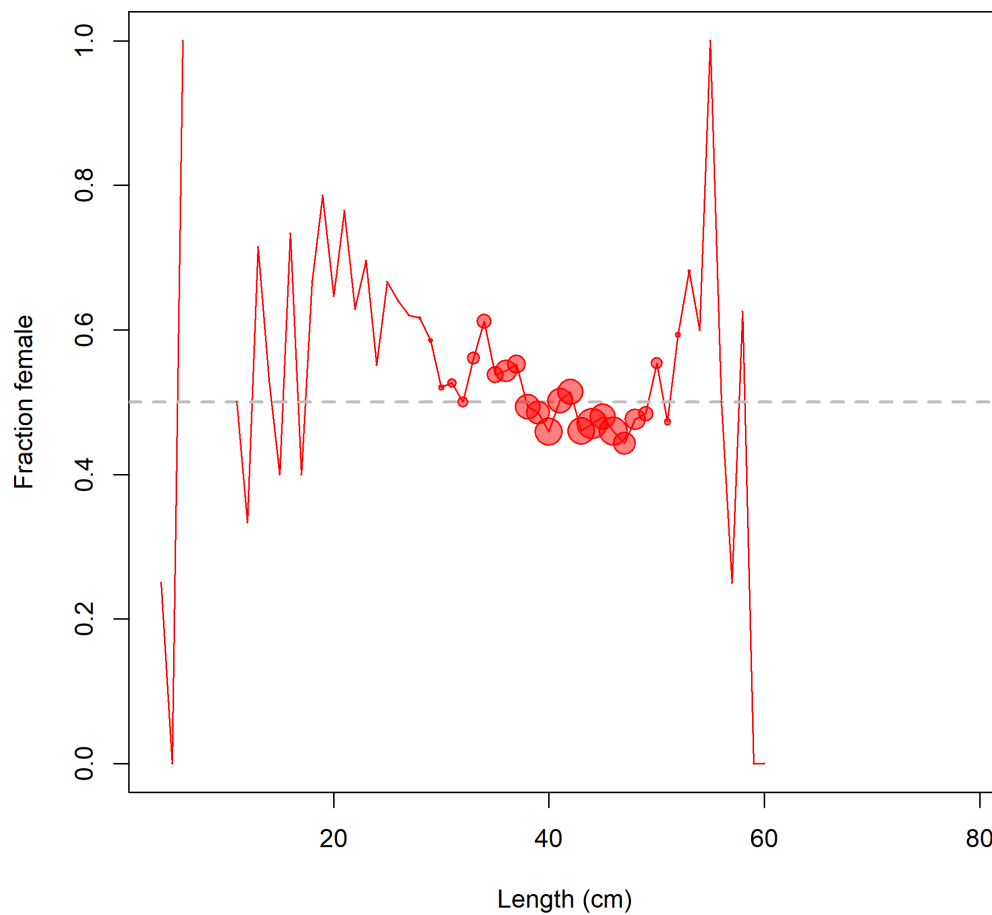
**Figure 10:** Length at age in the start of the year in the ending year of the model.



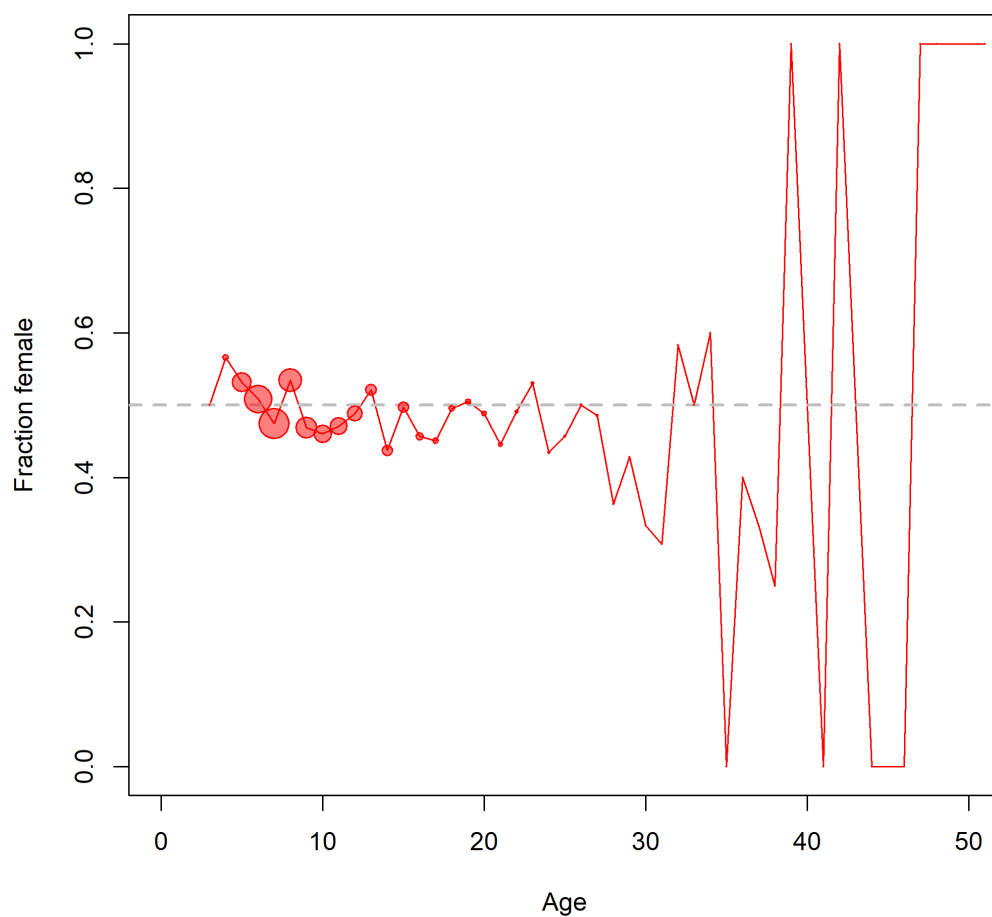
**Figure 11:** Maturity as a function of length.



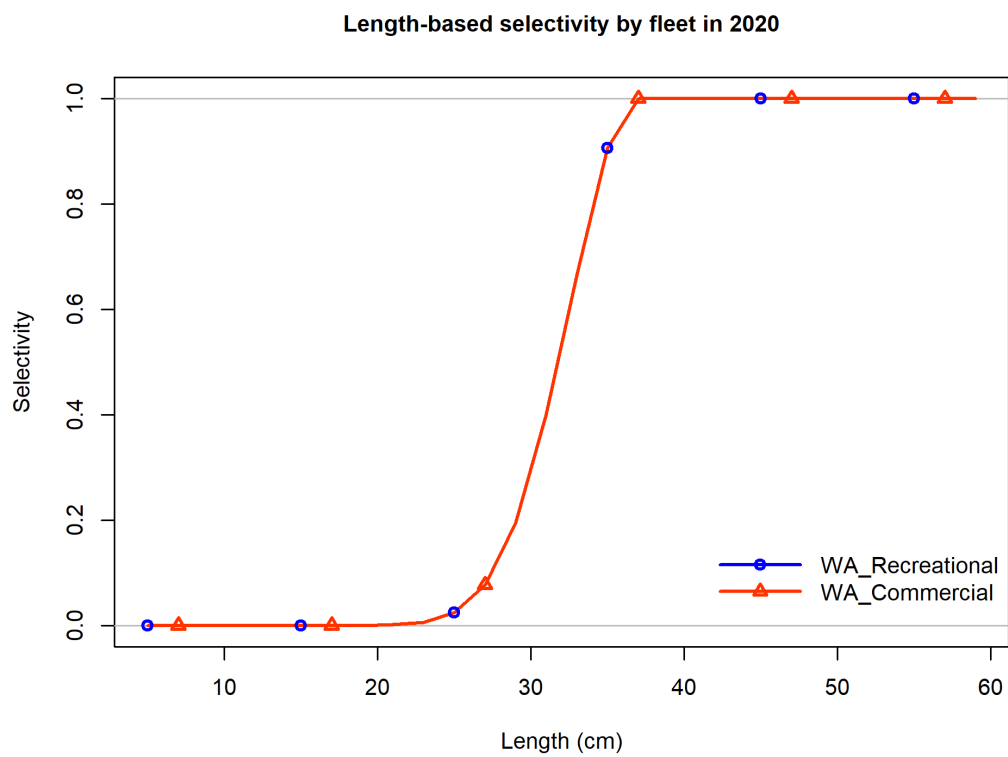
**Figure 12:** Fecundity as a function of length.



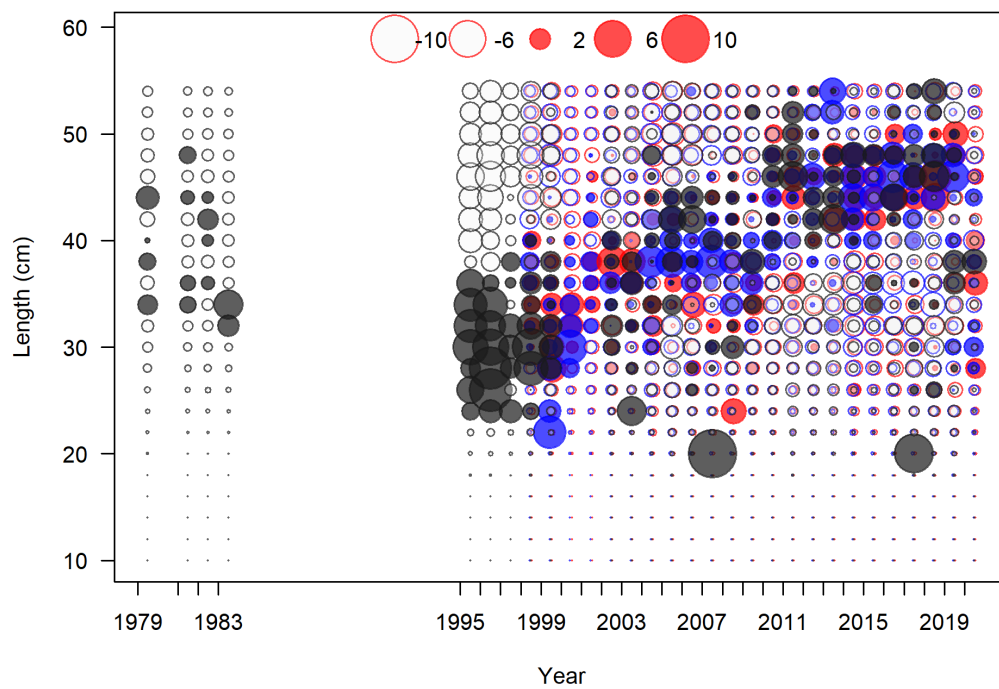
**Figure 13:** Fraction female by length across all available data sources where the size of red circles are based on the number of observations by length where larger circles indicate more observations.



**Figure 14:** Fraction female by age across all available data sources where the size of red circles are based on the number of observations by age where larger circles indicate more observations.

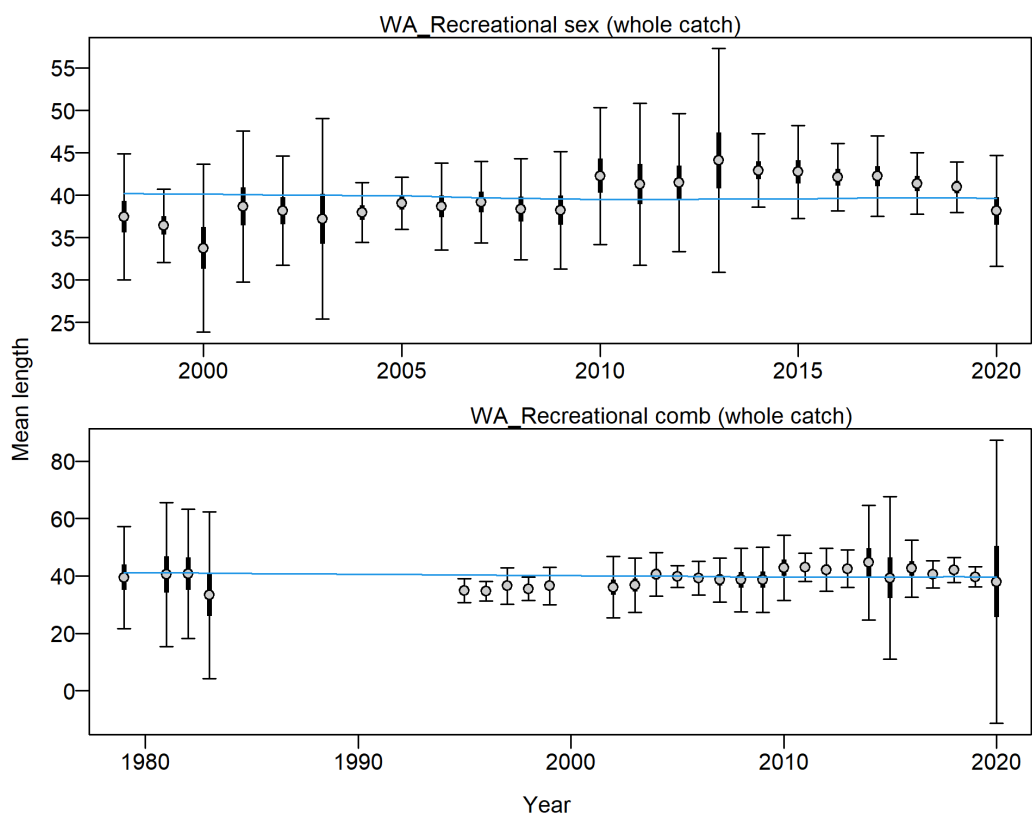


**Figure 15:** Selectivity at length by fleet.

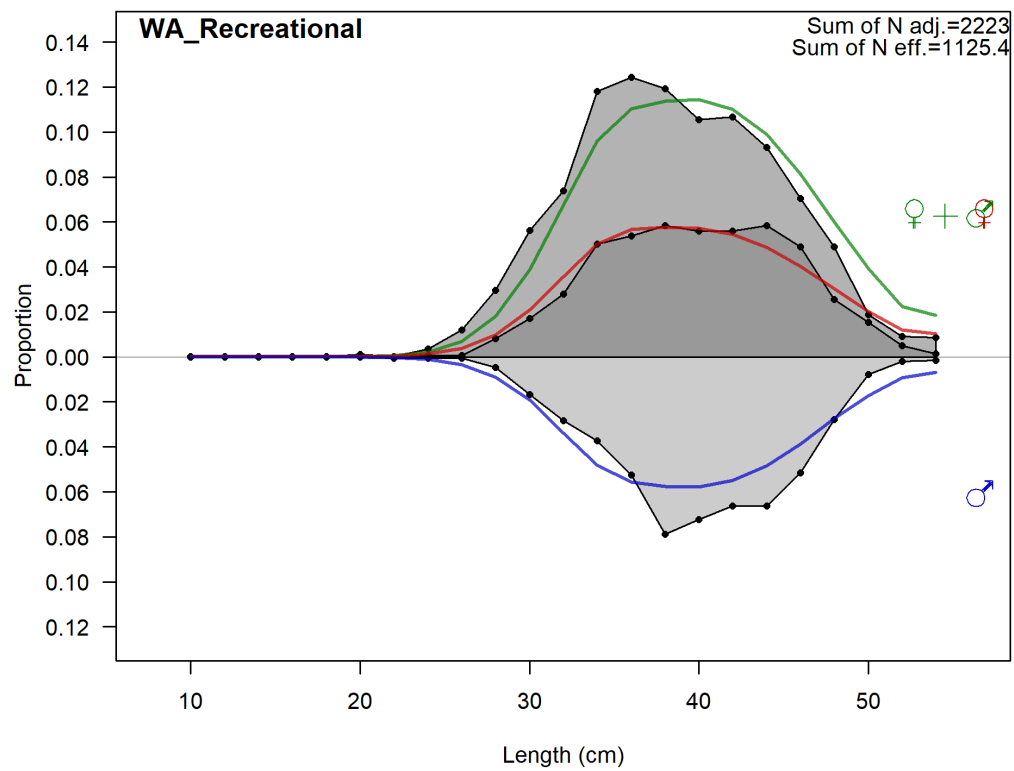


**Figure 16:** Pearson residuals for the recreational fleet. Closed bubble are positive residuals (observed > expected) and open bubbles are negative residuals (observed < expected).

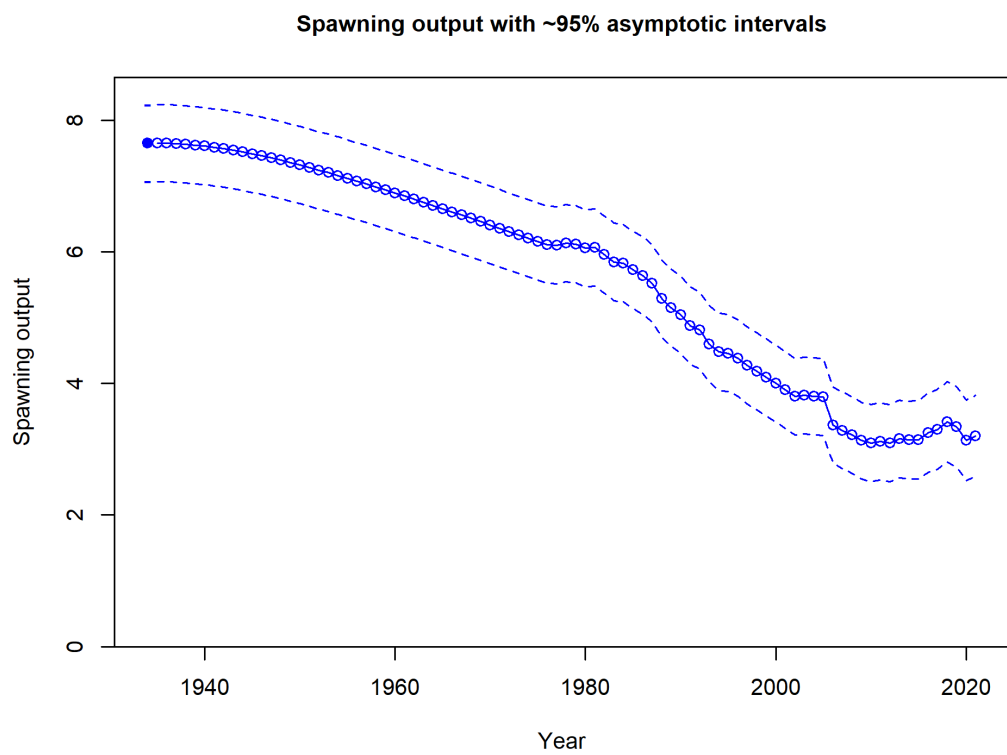




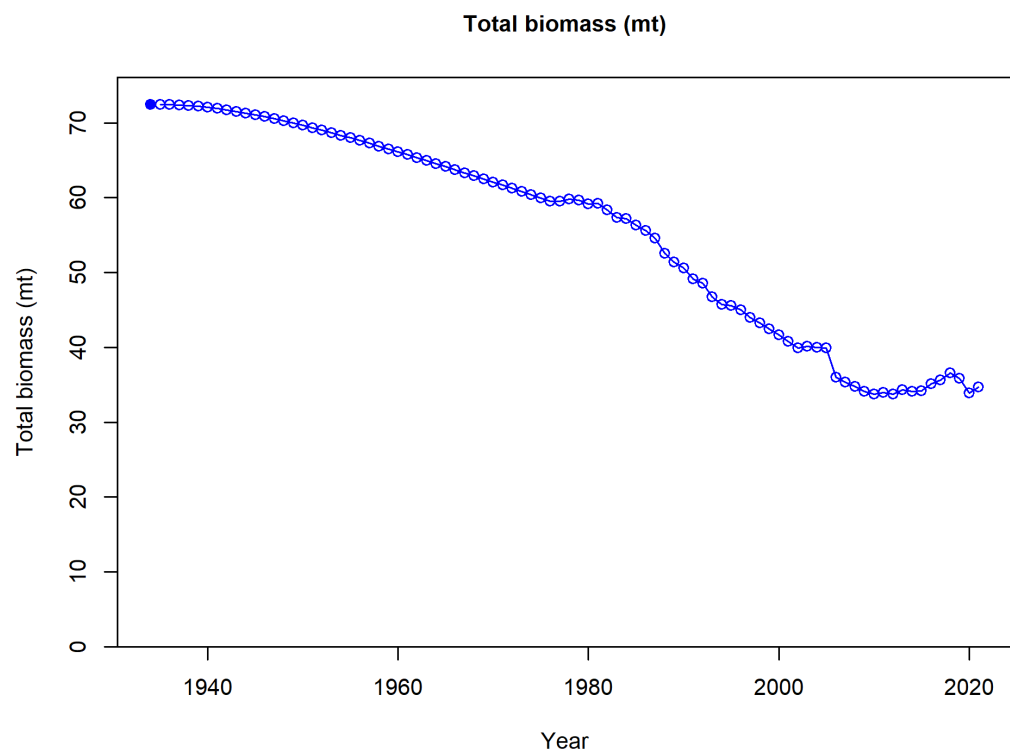
**Figure 17:** Mean length for recreational with 95 percent confidence intervals based on current samples sizes.



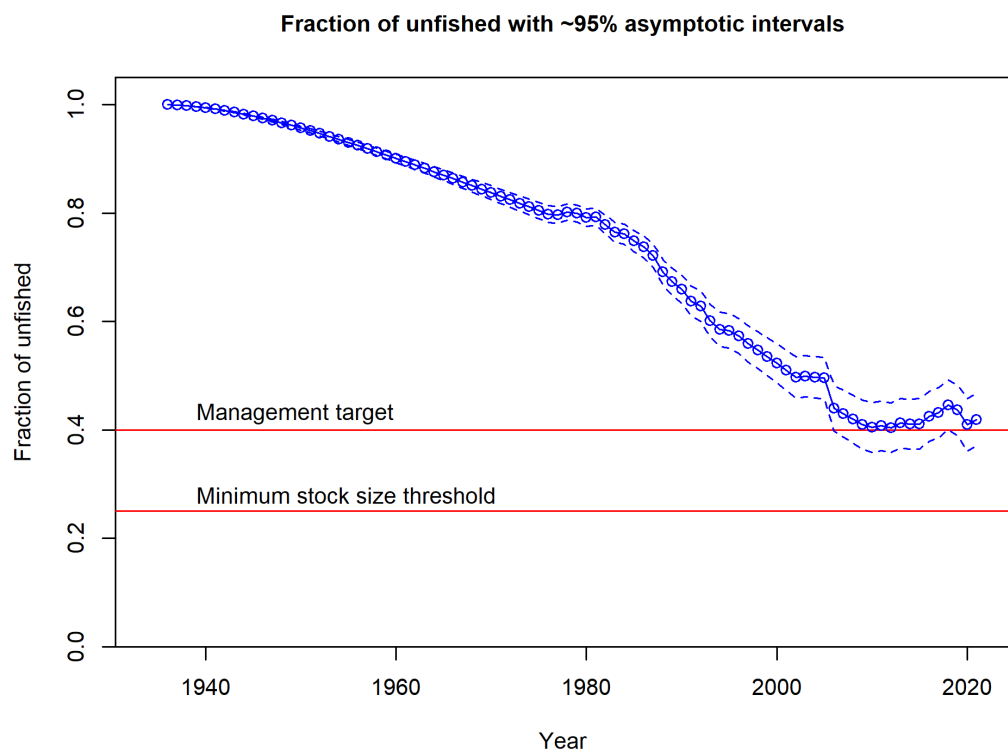
**Figure 18:** Aggregated length comps across all years.



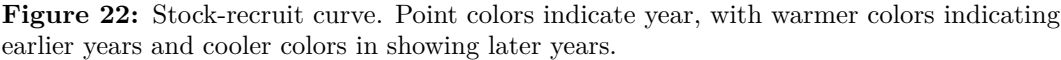
**Figure 19:** Estimated time series of spawning output.

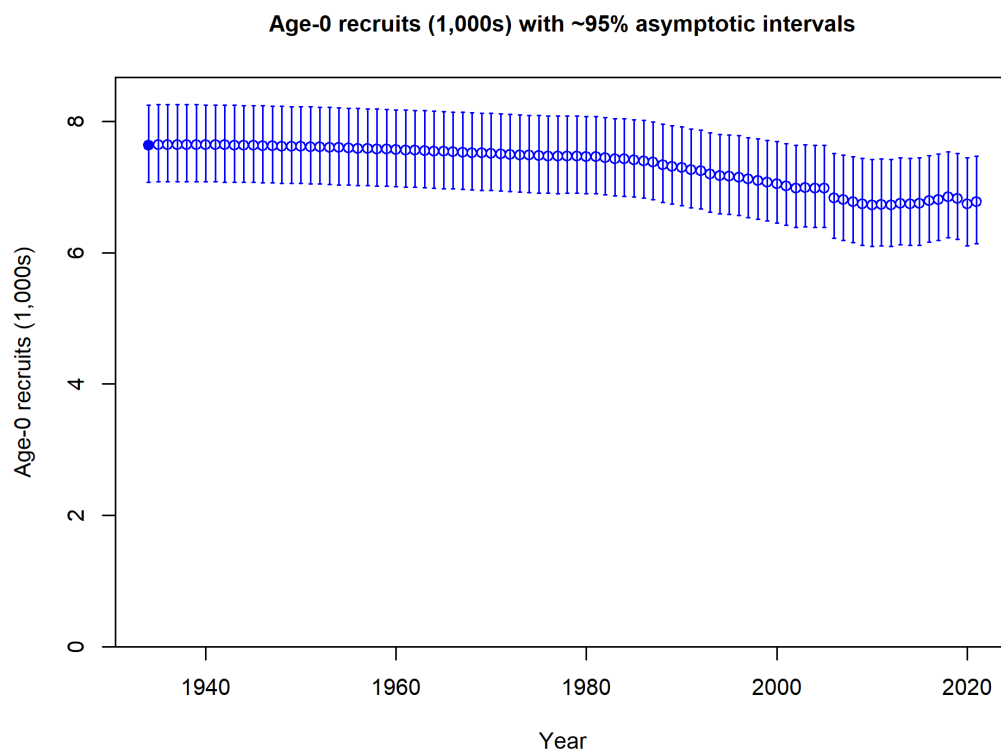


**Figure 20:** Estimated time series of total biomass.

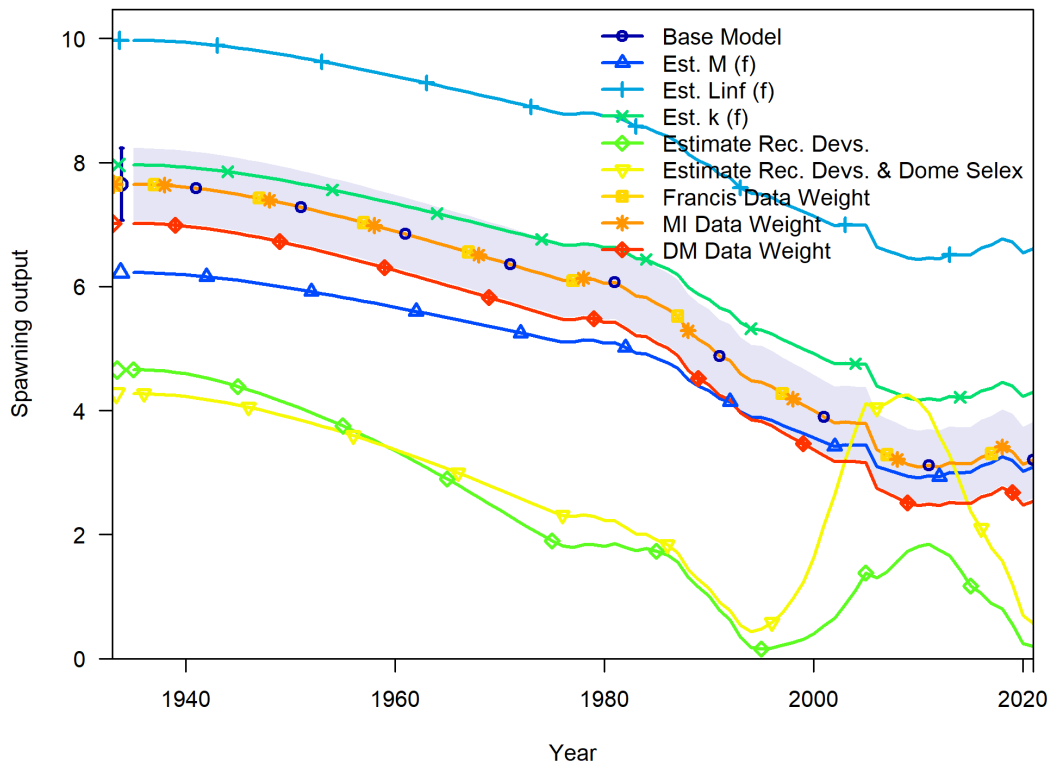


**Figure 21:** Estimated time series of fraction of unfished spawning output.



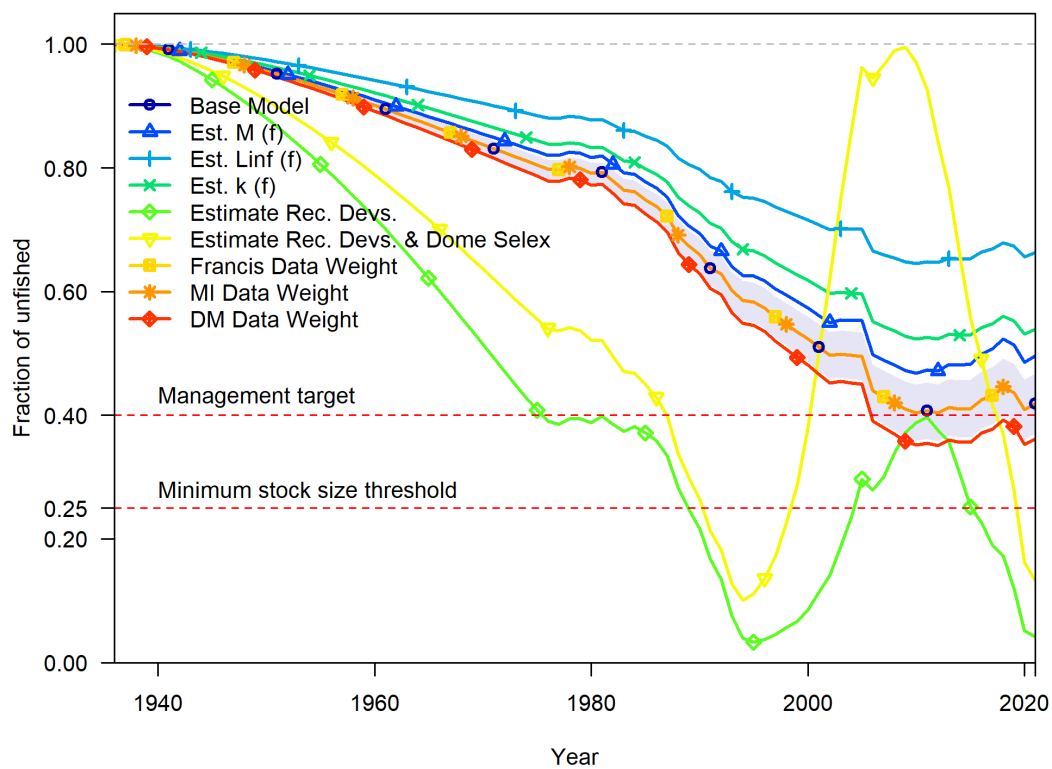


**Figure 23:** Estimated time series of age-0 recruits (1000s).

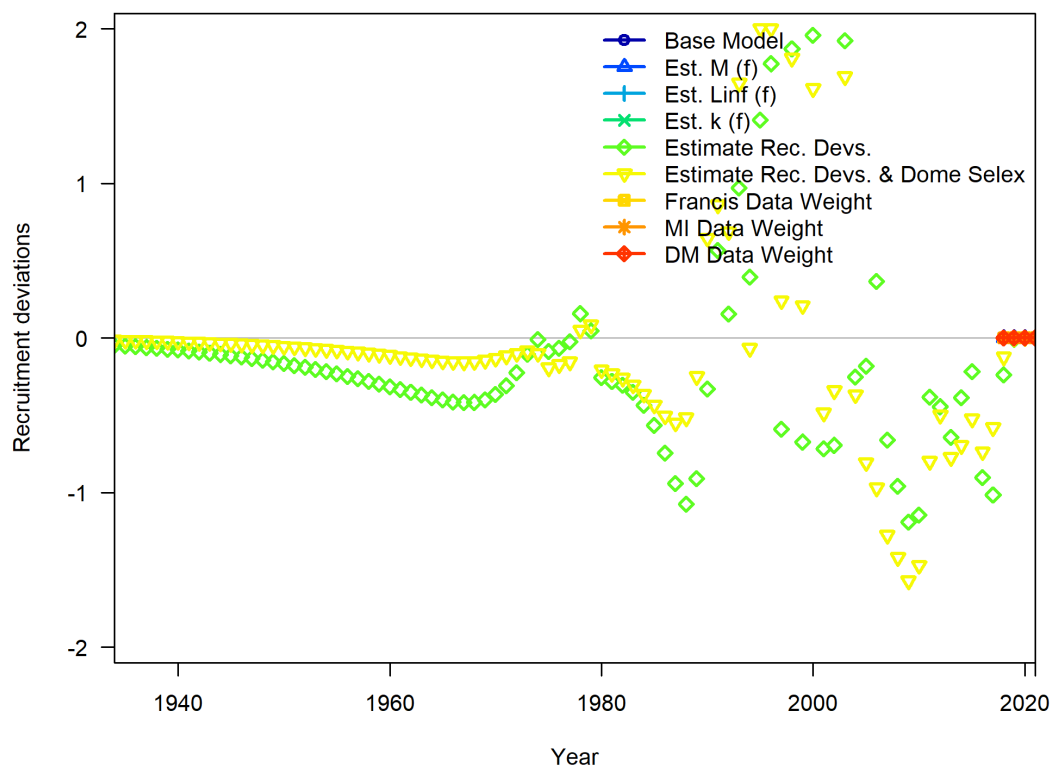


**Figure 24:** Change in estimated spawning output by sensitivity.

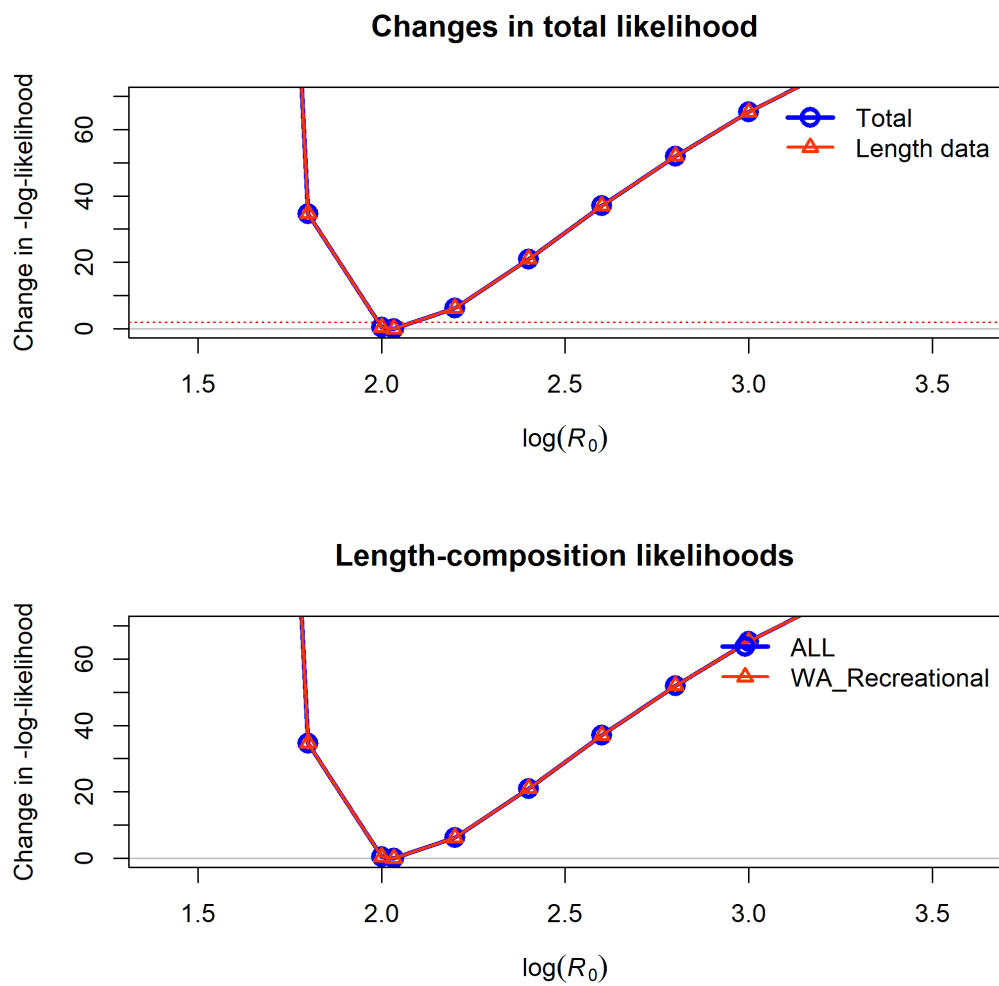




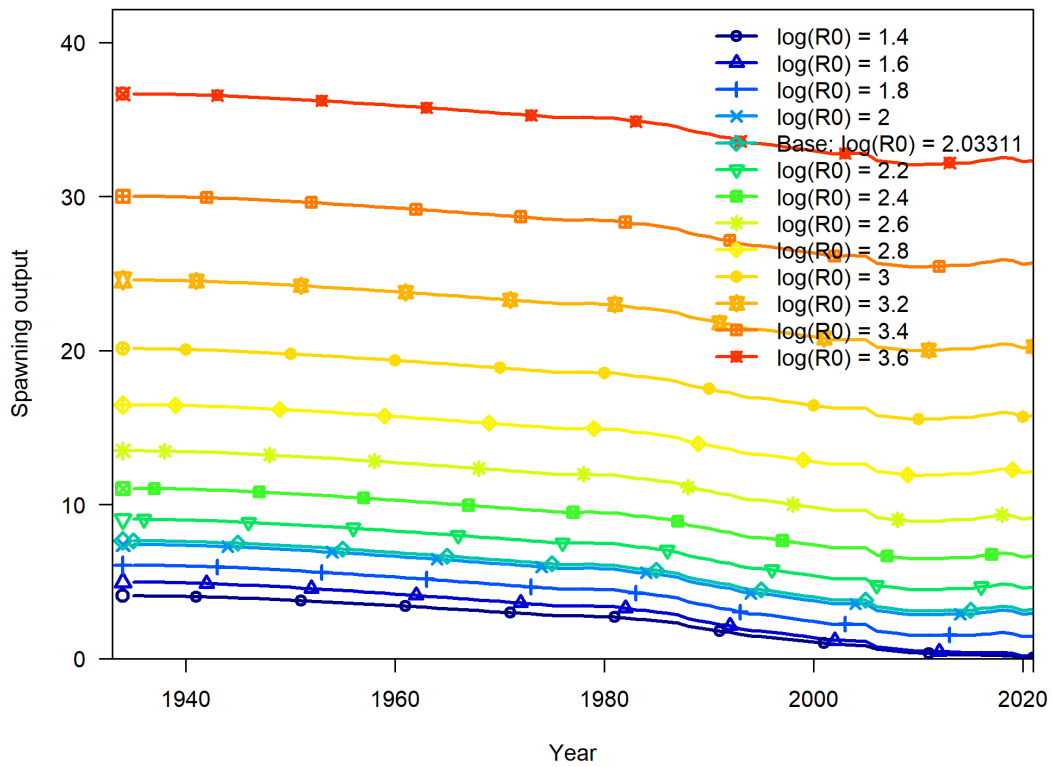
**Figure 25:** Change in estimated fraction unfished by sensitivity.



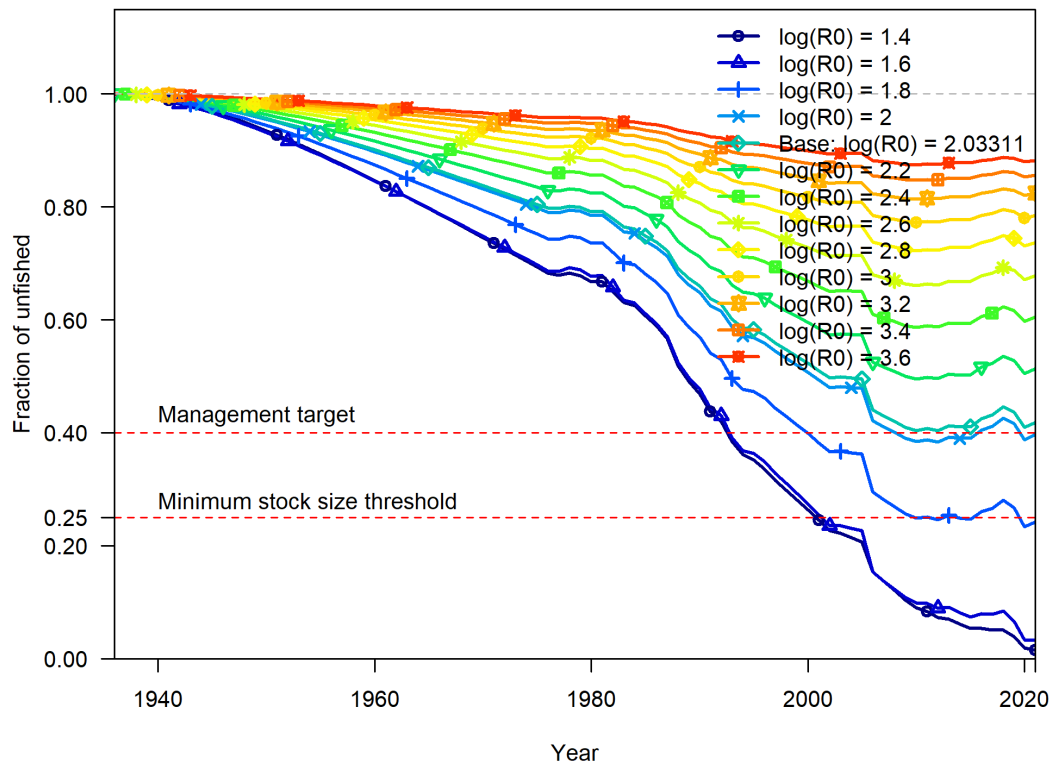
**Figure 26:** Change in estimated annual recruitment deviations by sensitivity.



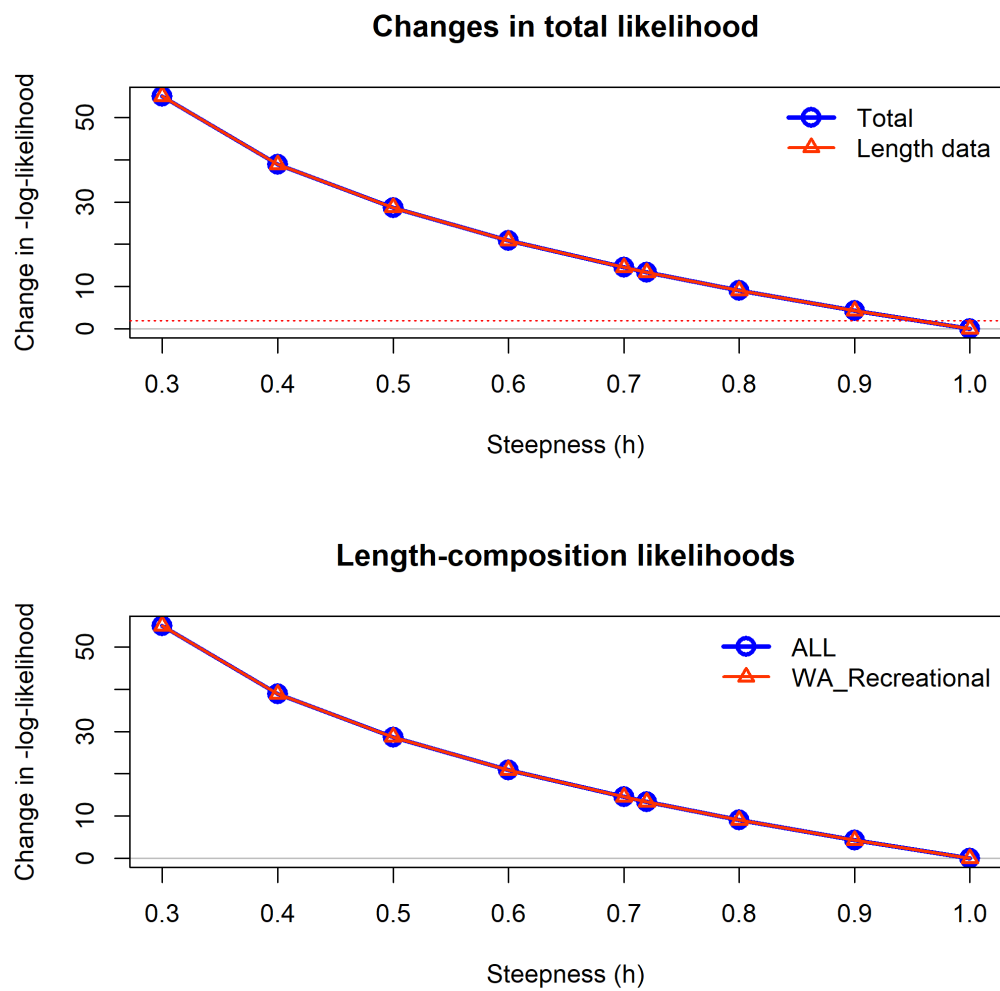
**Figure 27:** Change in the negative log-likelihood across a range of  $\log(R_0)$  values.



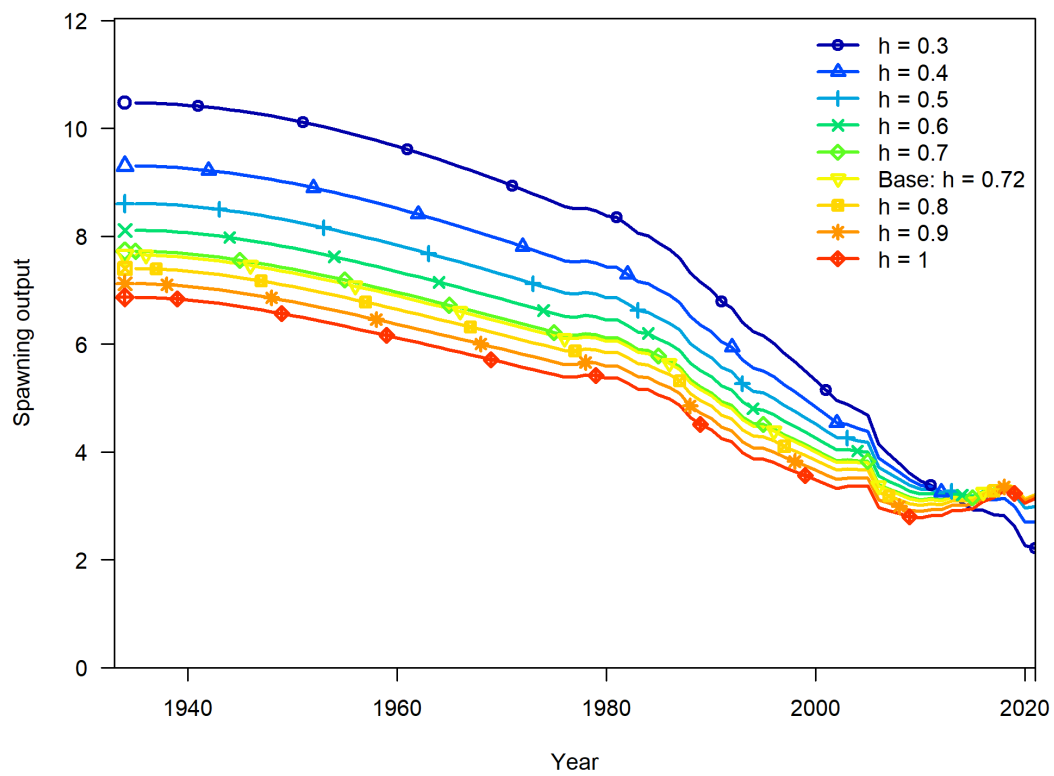
**Figure 28:** Change in the estimate of spawning output across a range of  $\log(R_0)$  values.



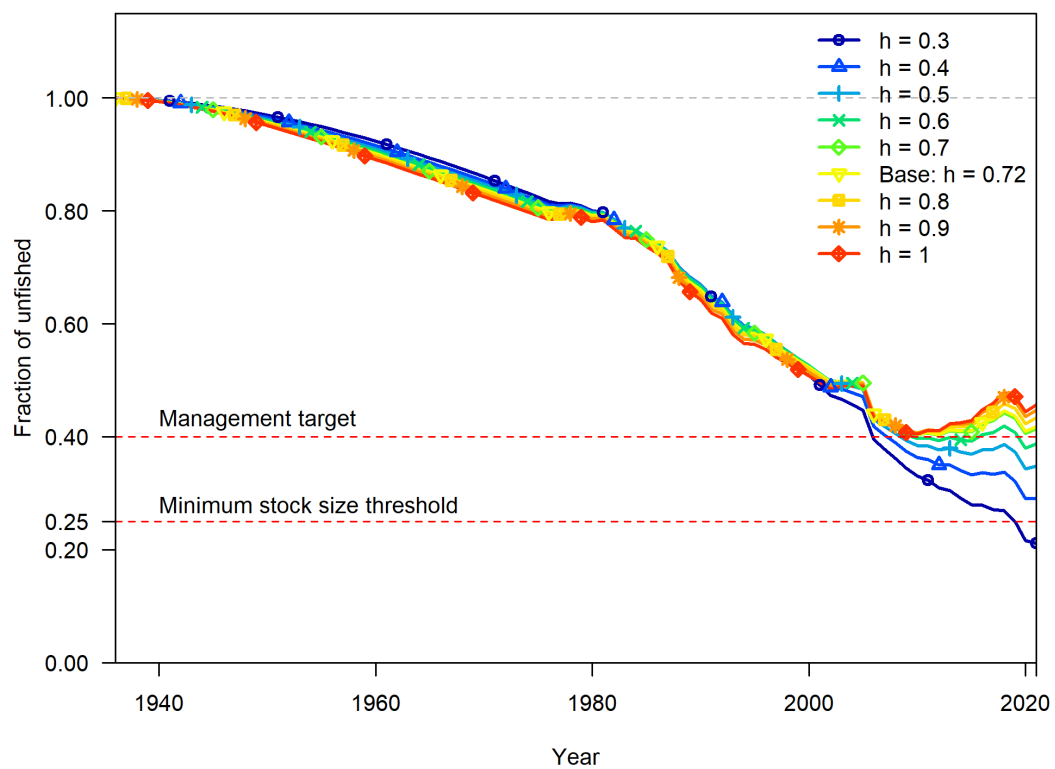
**Figure 29:** Change in the estimate of fraction unfished across a range of  $\log(R_0)$  values.



**Figure 30:** Change in the negative log-likelihood across a range of steepness values.

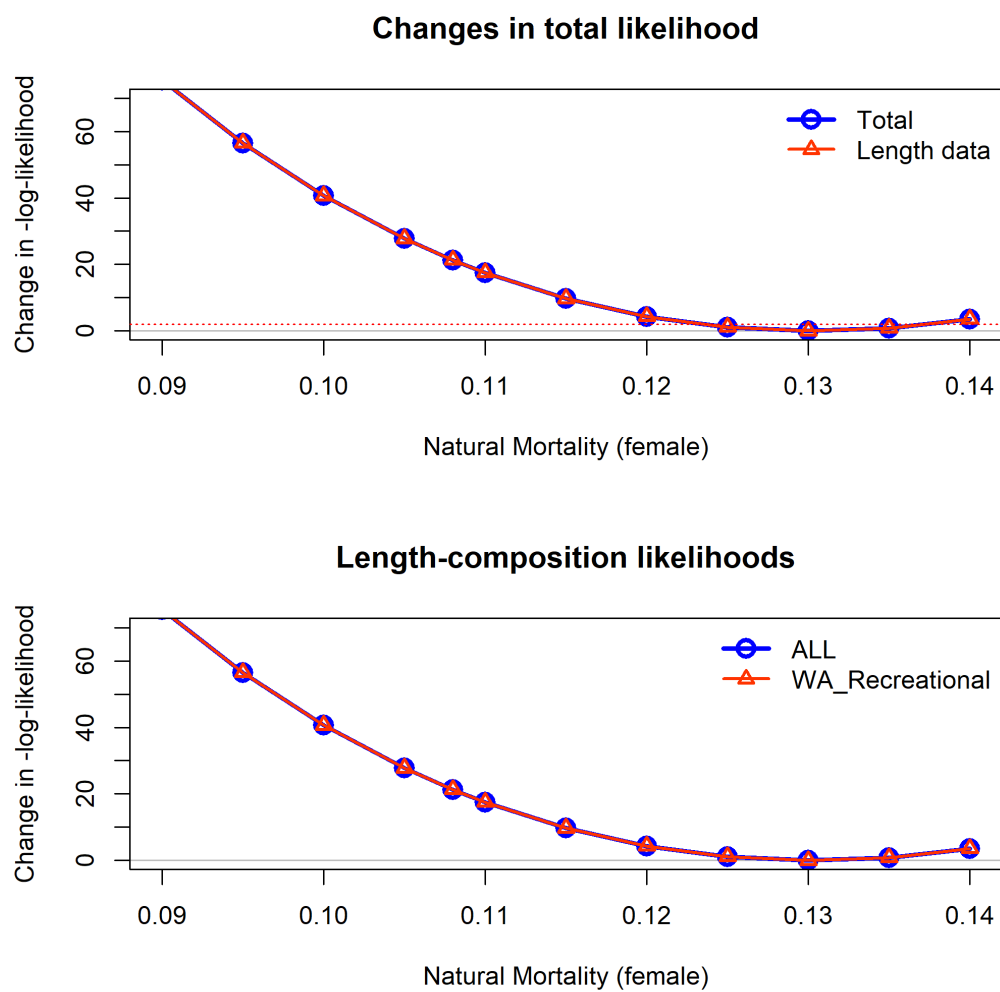


**Figure 31:** Change in the estimate of spawning output across a range of steepness values.

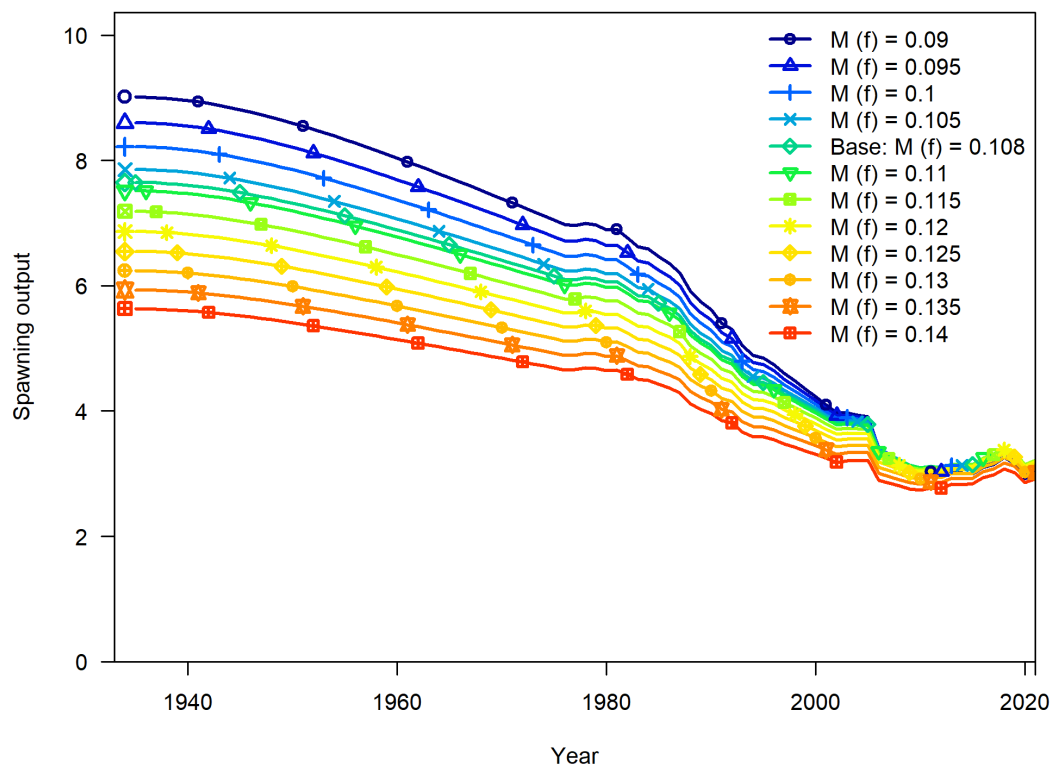


**Figure 32:** Change in the estimate of fraction unfished across a range of steepness values.

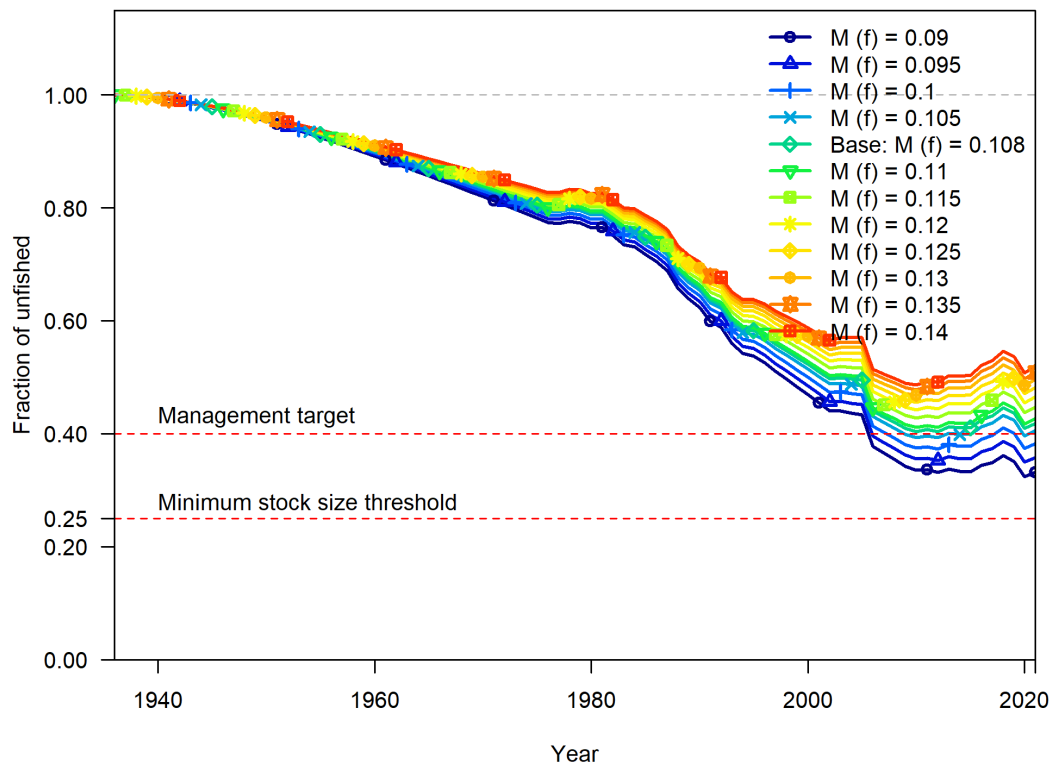




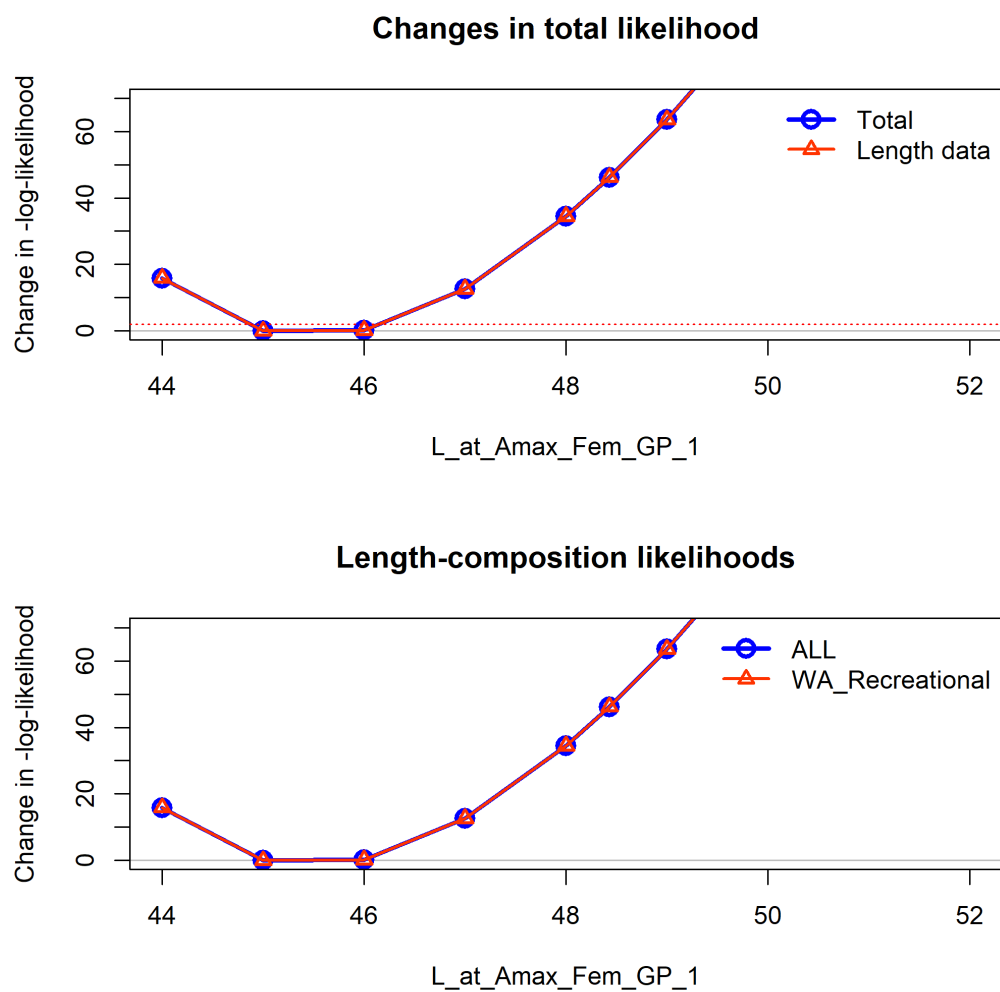
**Figure 33:** Change in the negative log-likelihood across a range of female natural mortality values.



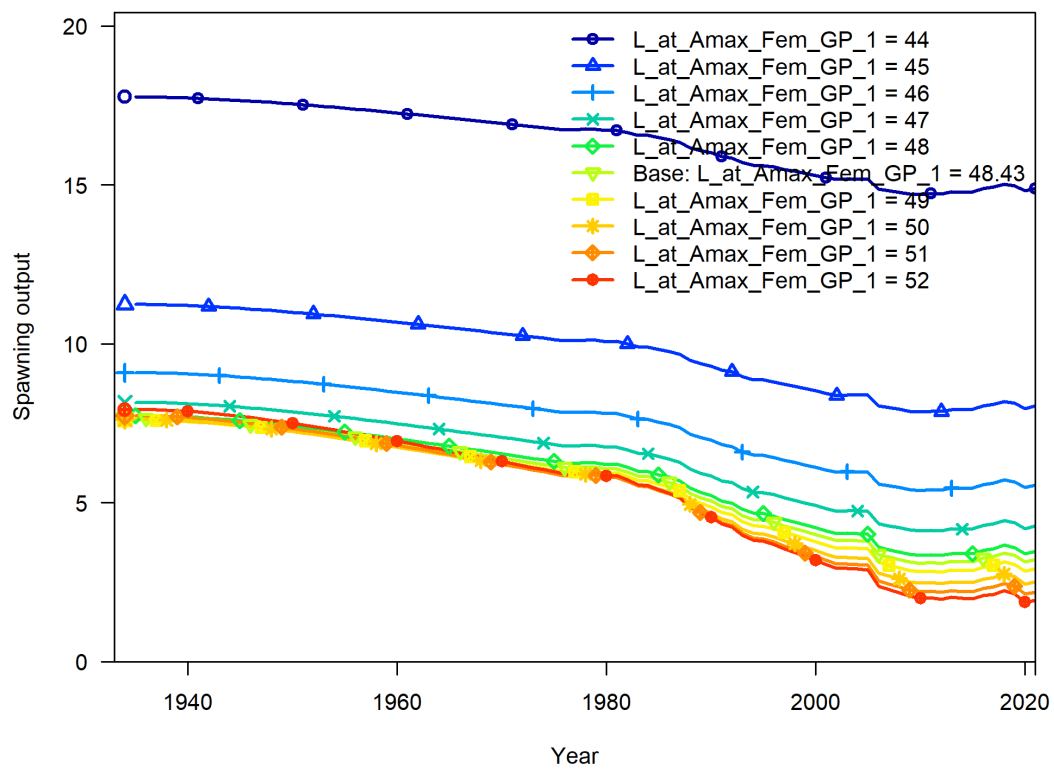
**Figure 34:** Change in the estimate of spawning output across a range of female natural mortality values.



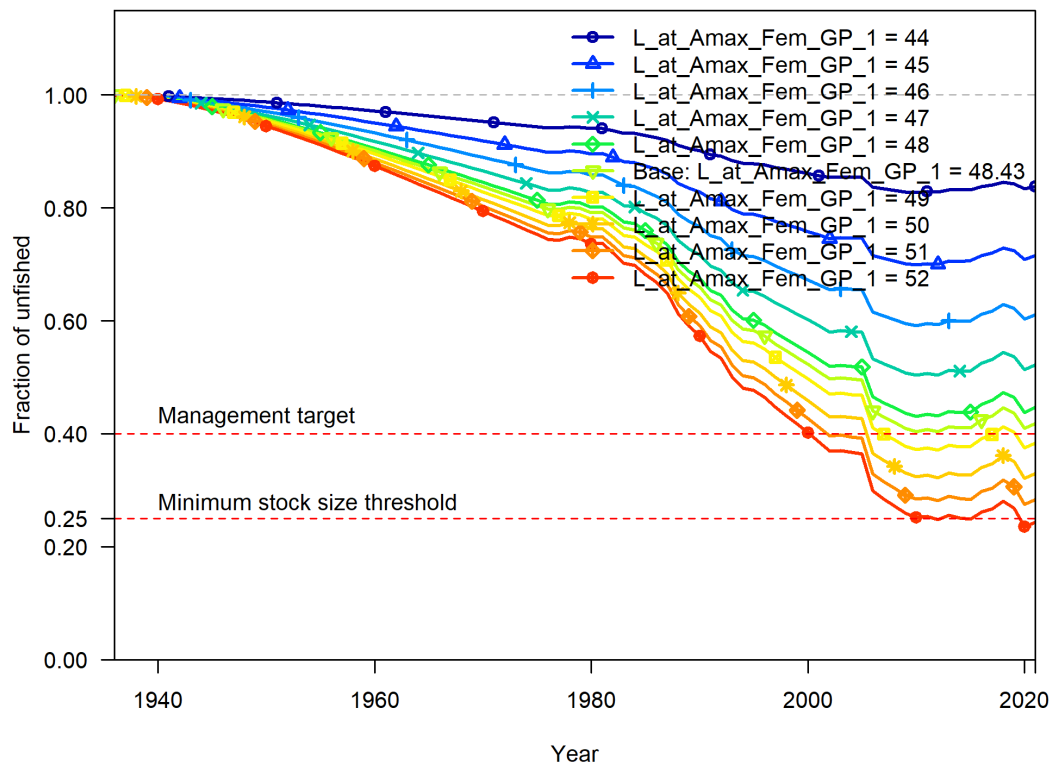
**Figure 35:** Change in the estimate of fraction unfished across a range of female natural values.



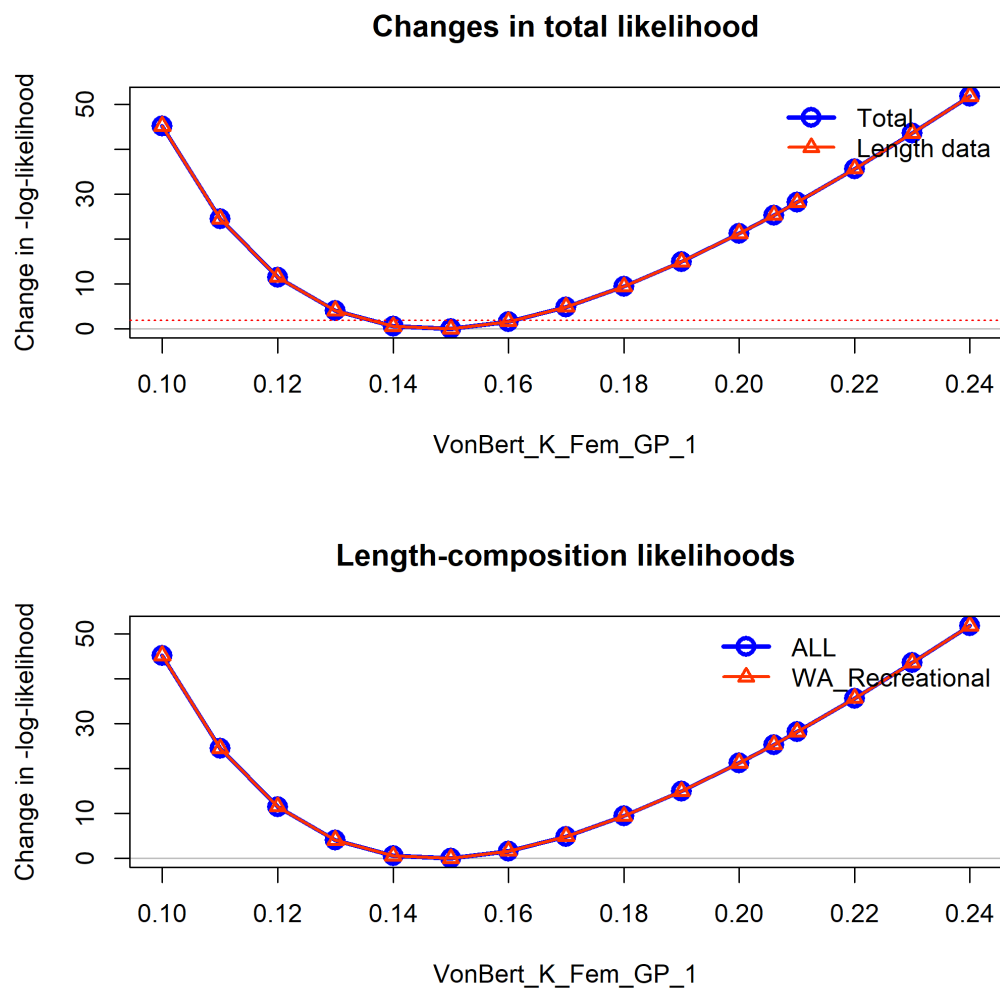
**Figure 36:** Change in the negative log-likelihood across a range of female maximum length values.



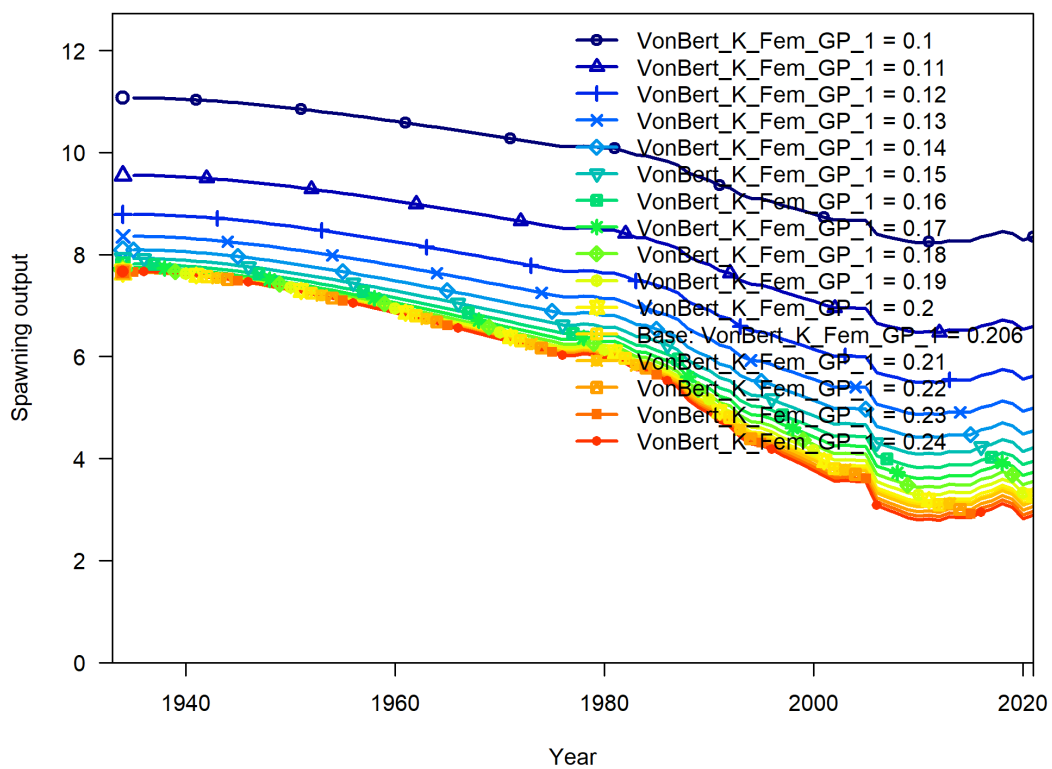
**Figure 37:** Change in the estimate of spawning output across a range of female maximum length values.



**Figure 38:** Change in the estimate of fraction unfished across a range of female maximum length values.

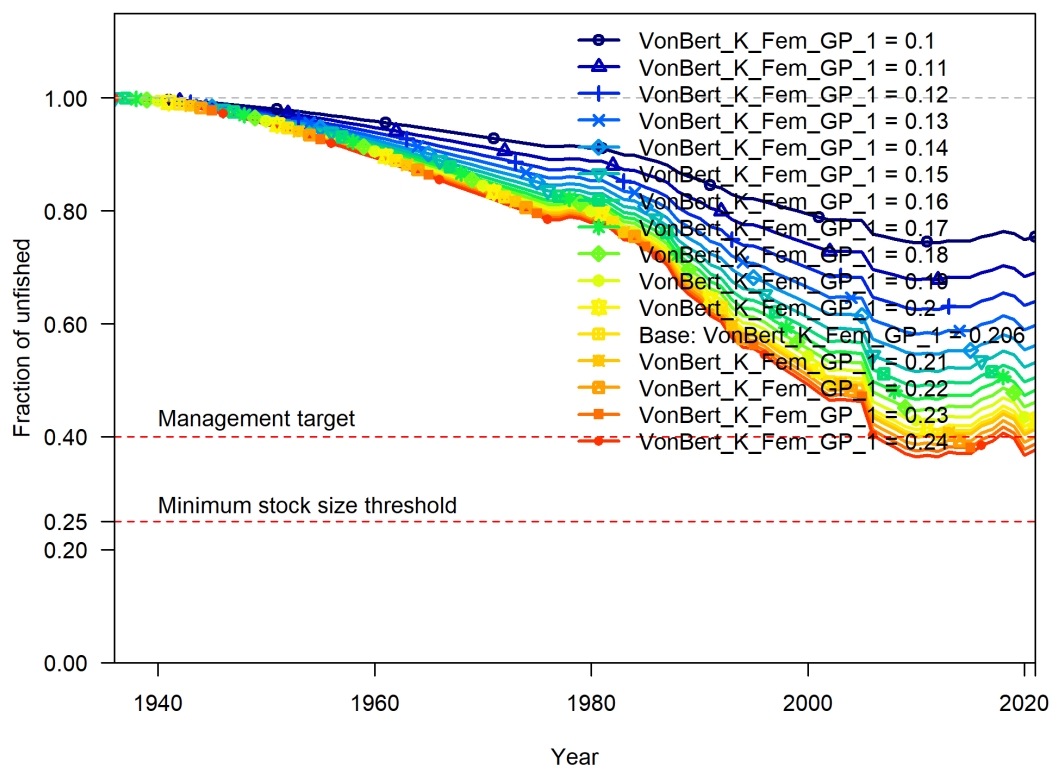


**Figure 39:** Change in the negative log-likelihood across a range of female k values.

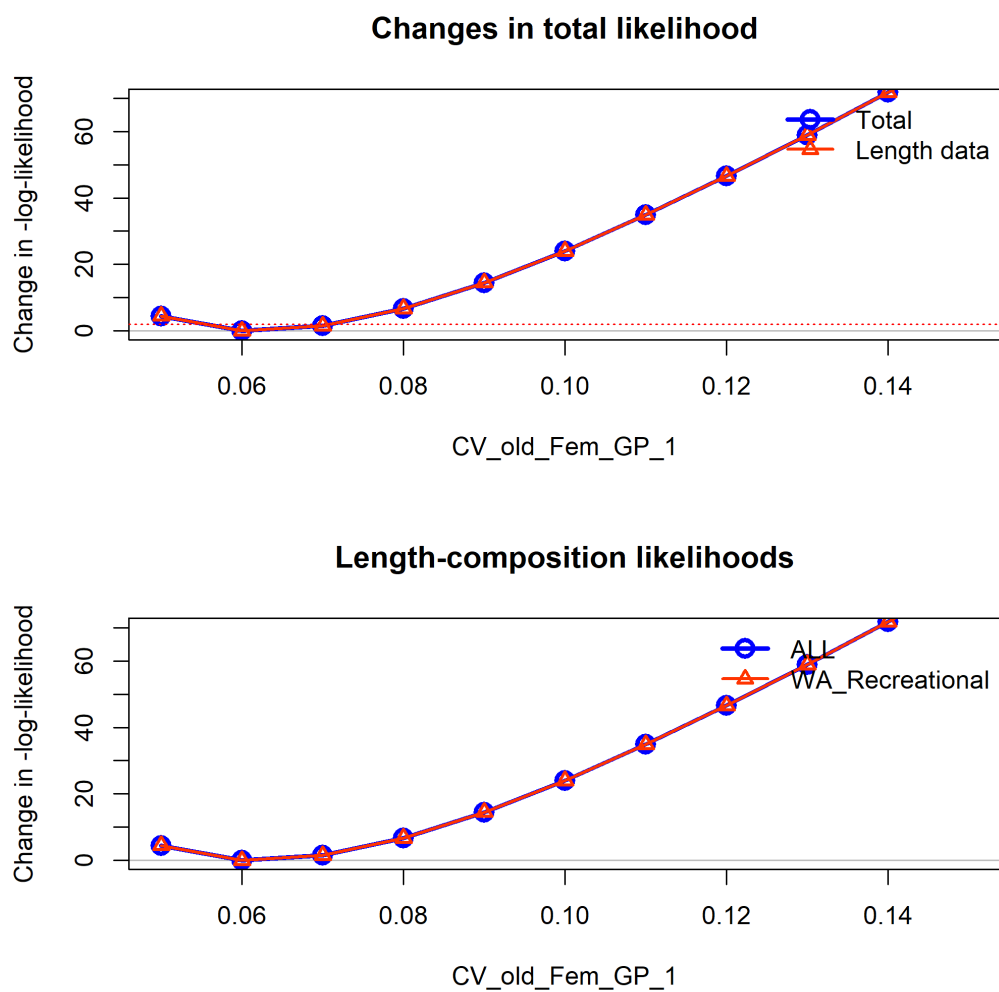


**Figure 40:** Change in the estimate of spawning output across a range of female k values.

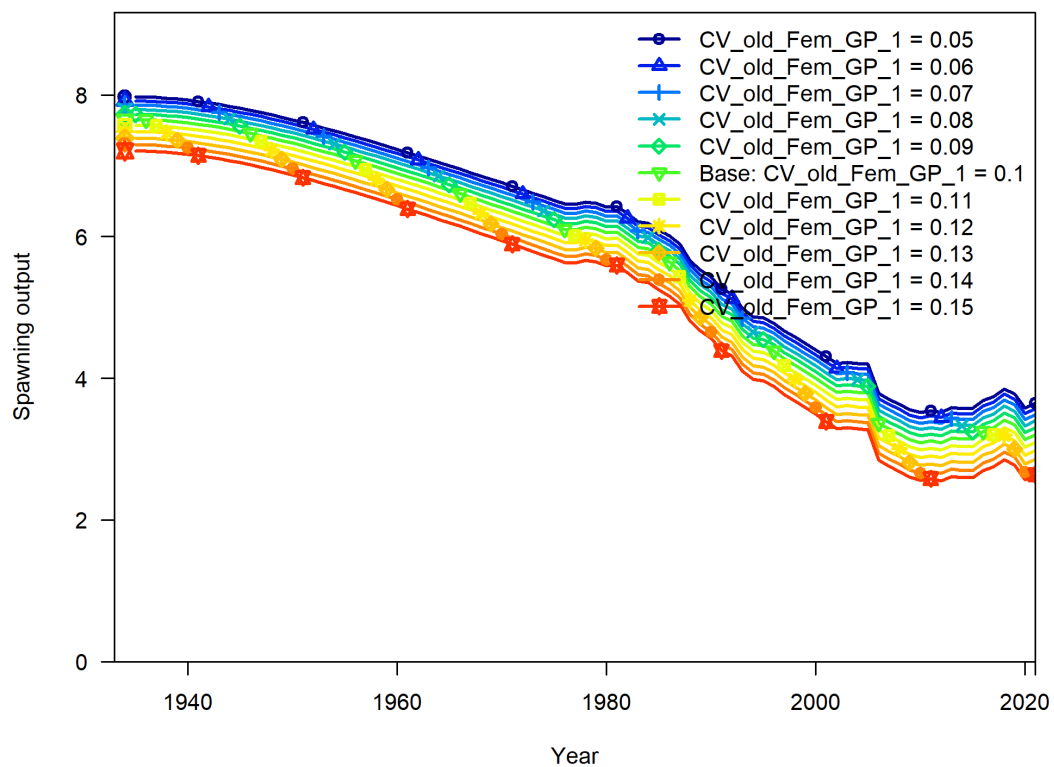




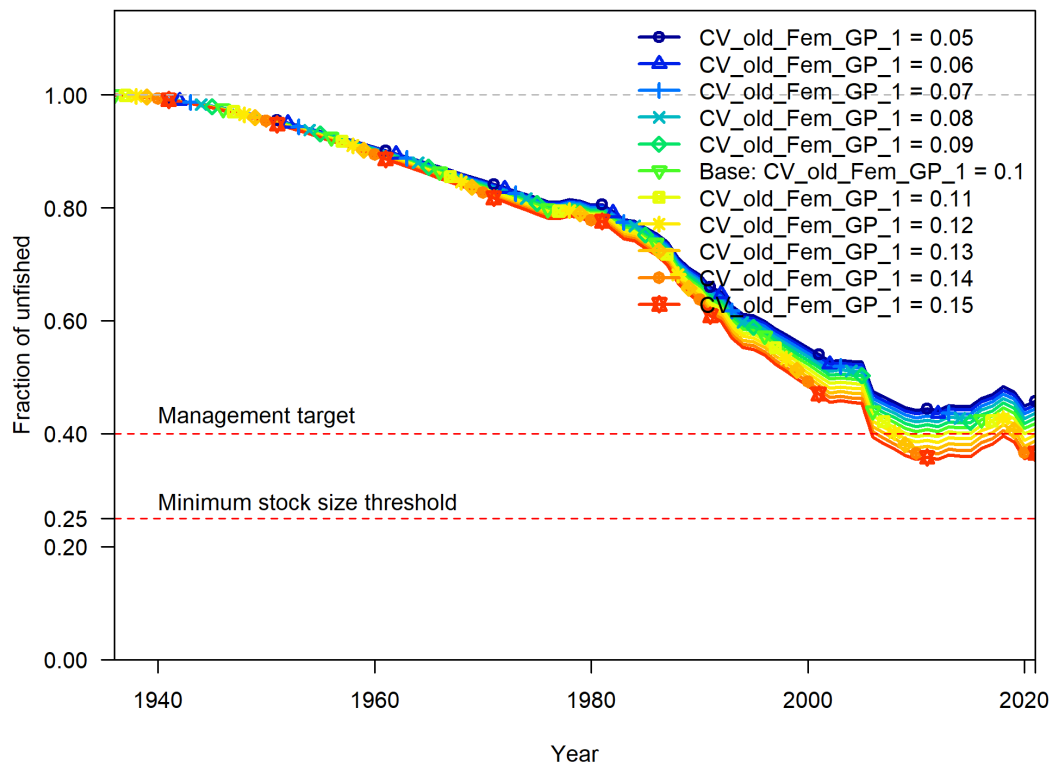
**Figure 41:** Change in the estimate of fraction unfished across a range of female  $k$  values.



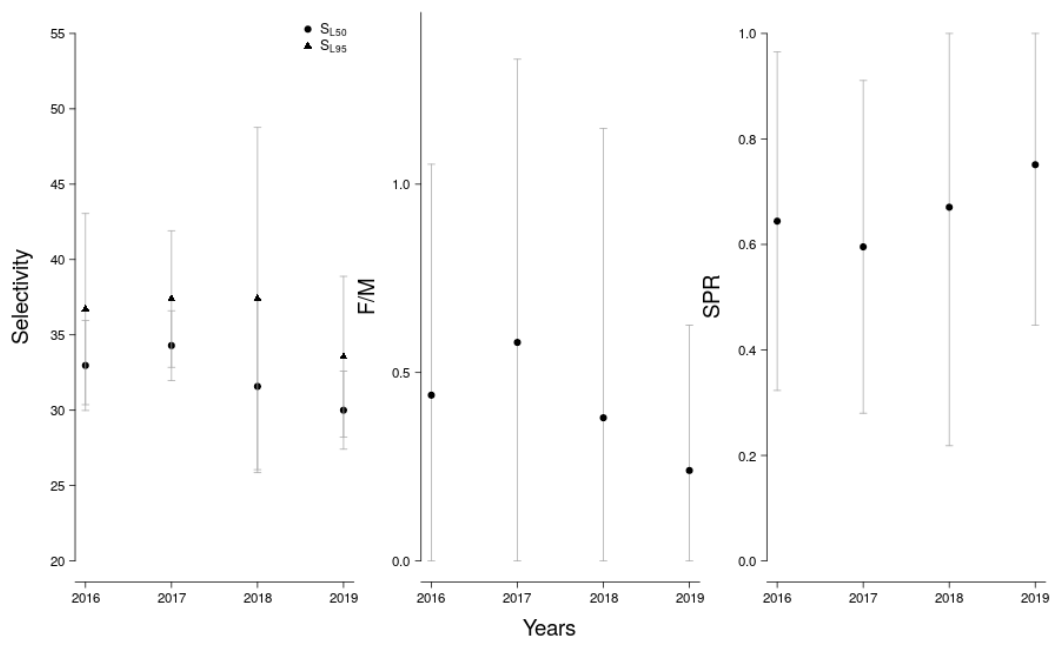
**Figure 42:** Change in the negative log-likelihood across a range of female coefficient of variation for older ages.



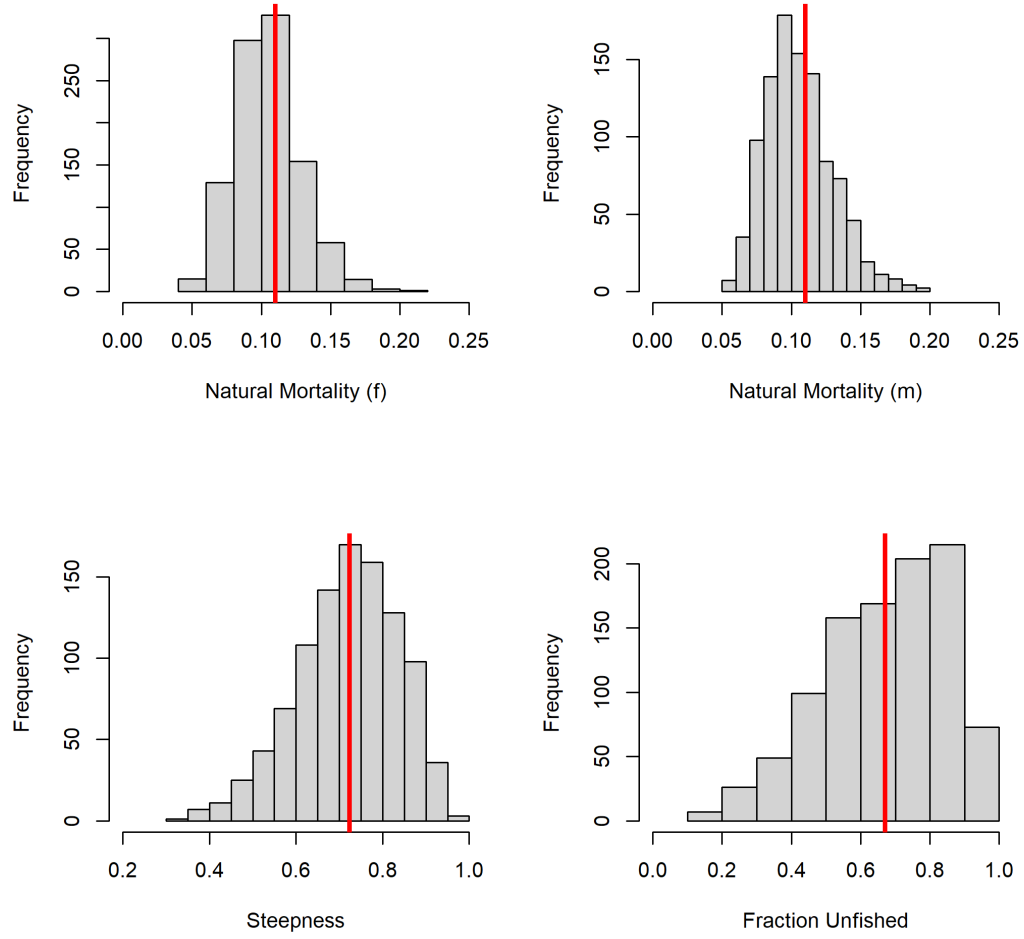
**Figure 43:** Change in the estimate of spawning output across a range of female coefficient of variation for older ages.



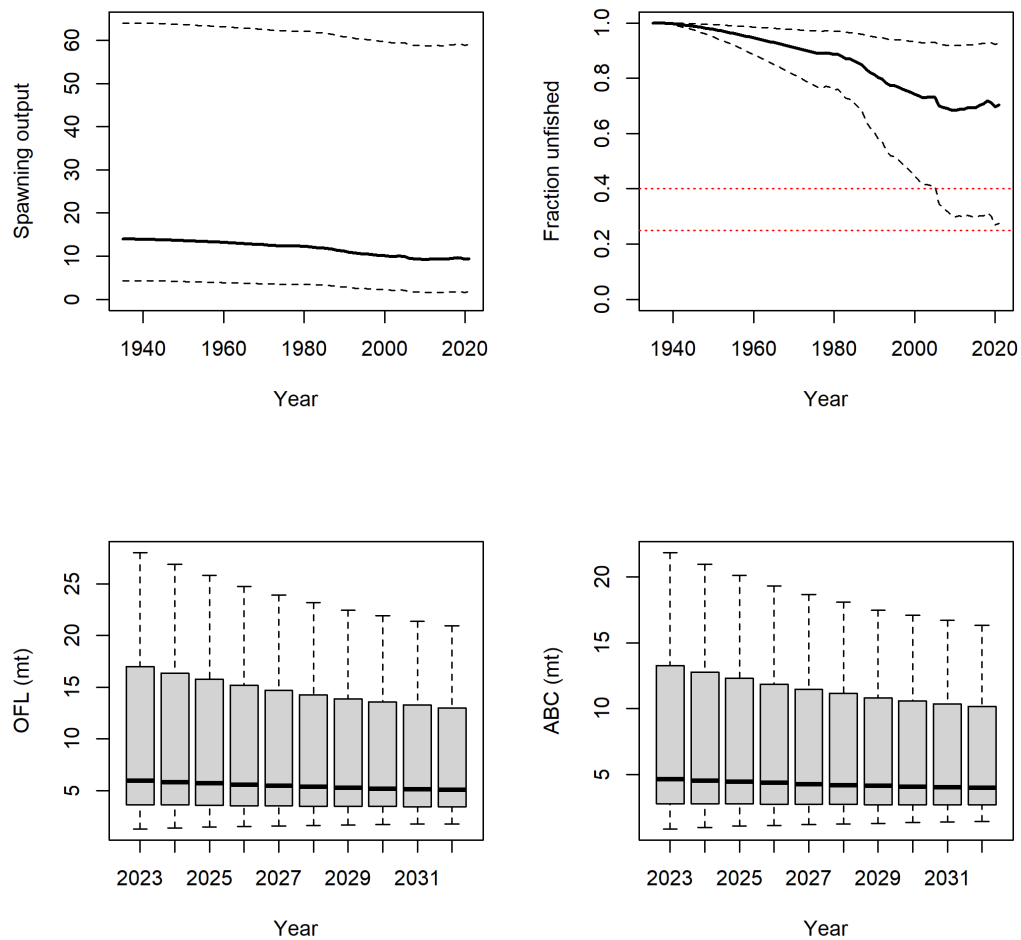
**Figure 44:** Change in the estimate of fraction unfished across a range of female coefficient of variation for older ages.



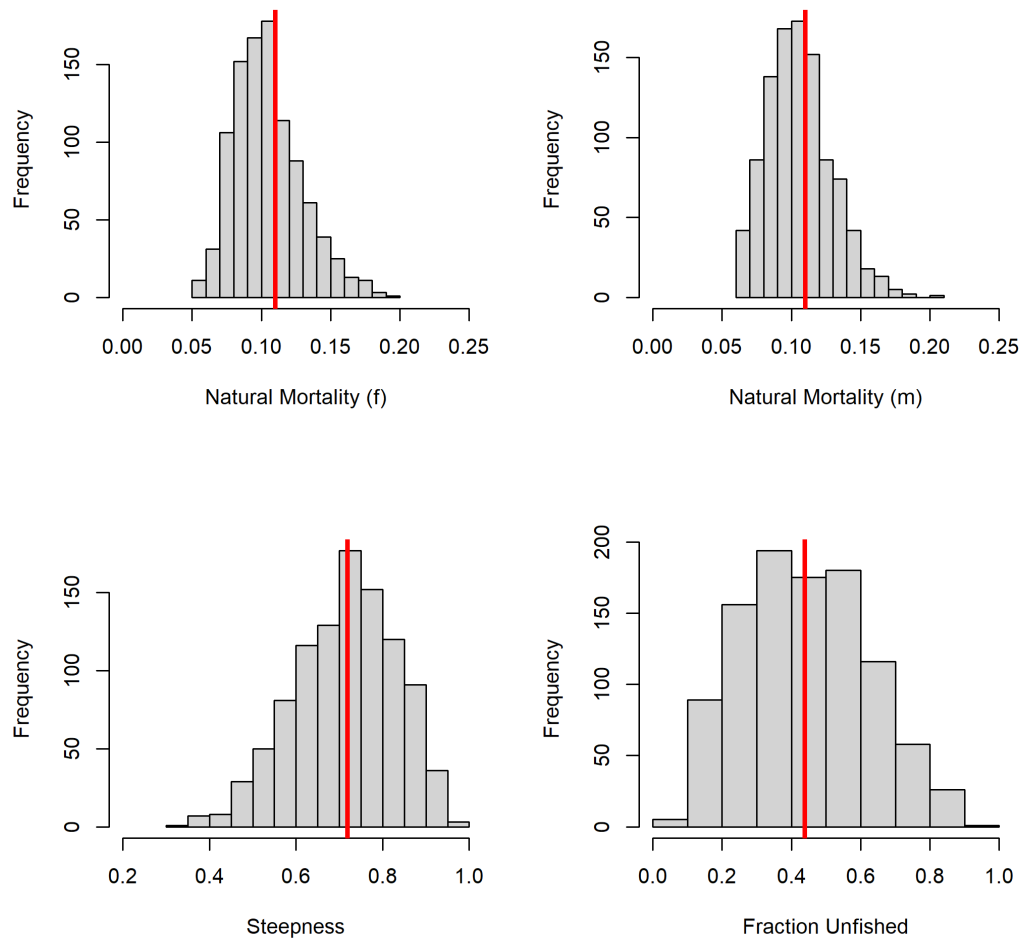
**Figure 45:** LB-SPR yearly estimates of selectivity, the ratio of fishing intensity to natural mortality (F/M), and annual spawner-per-recruit (SPR) values.



**Figure 46:** Prior distributions for parameter input for SSS.

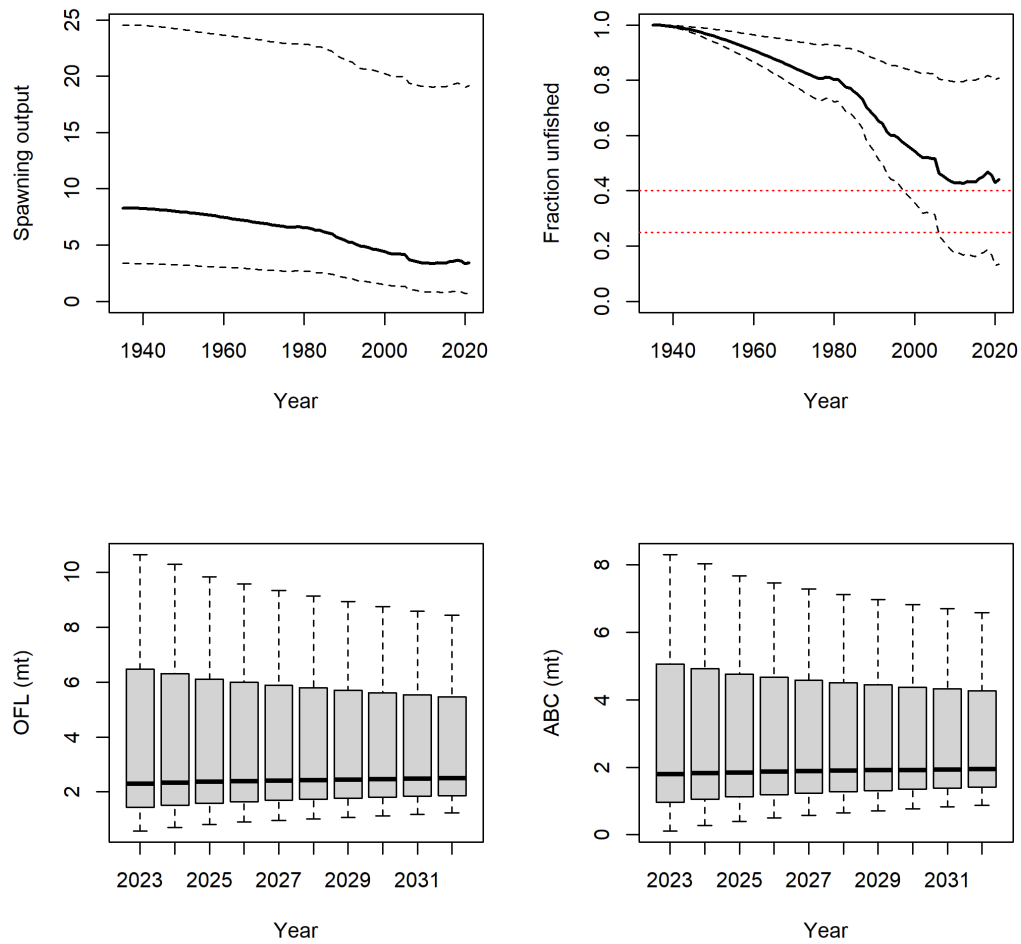


**Figure 47:** Derived quantities from the SSS run where fraction unfished was assumed to be 60 percent.

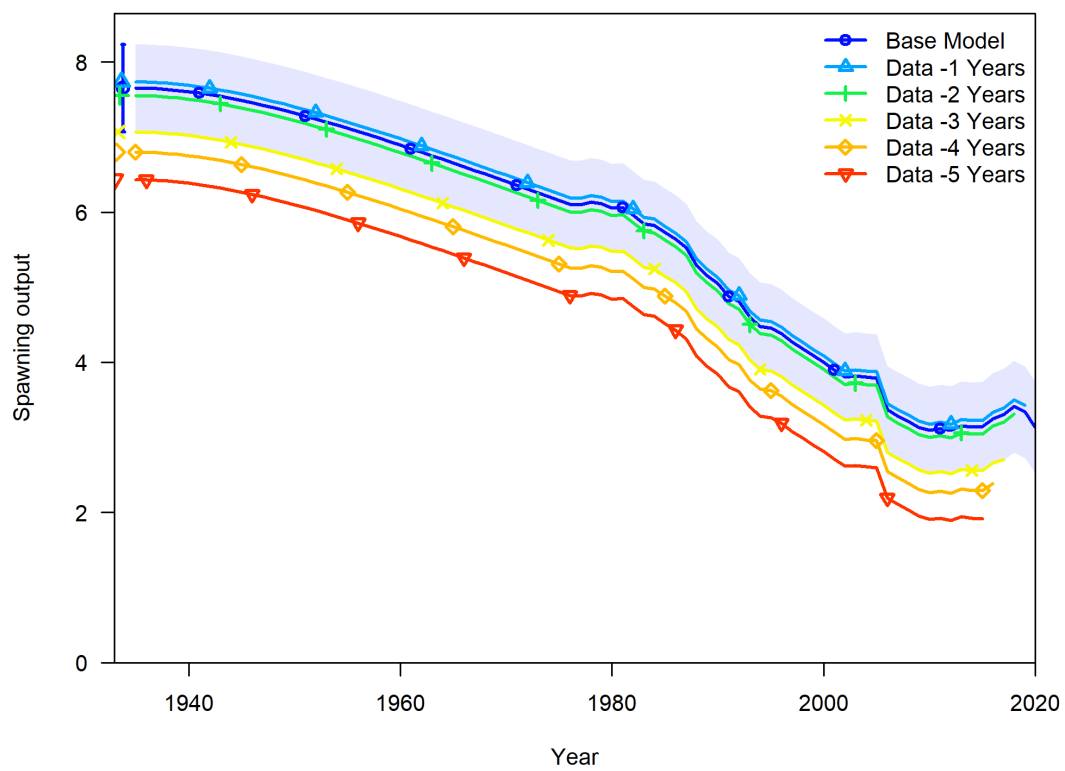


**Figure 48:** Prior distributions for parameter input for SSS.

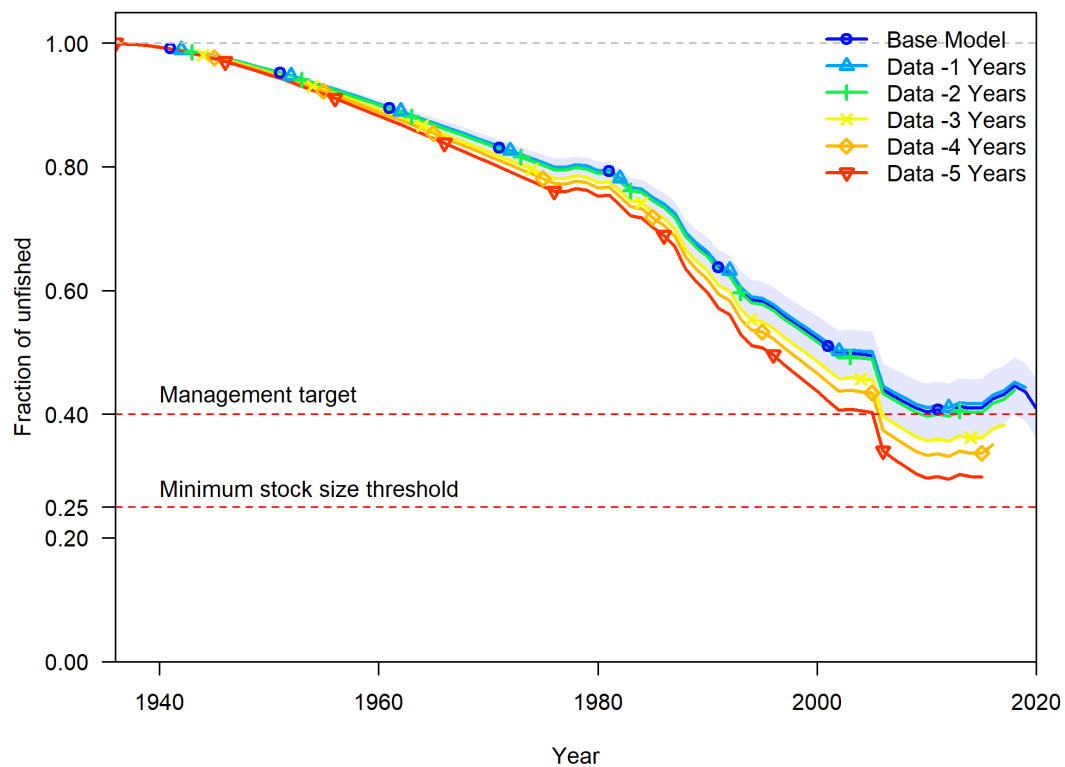




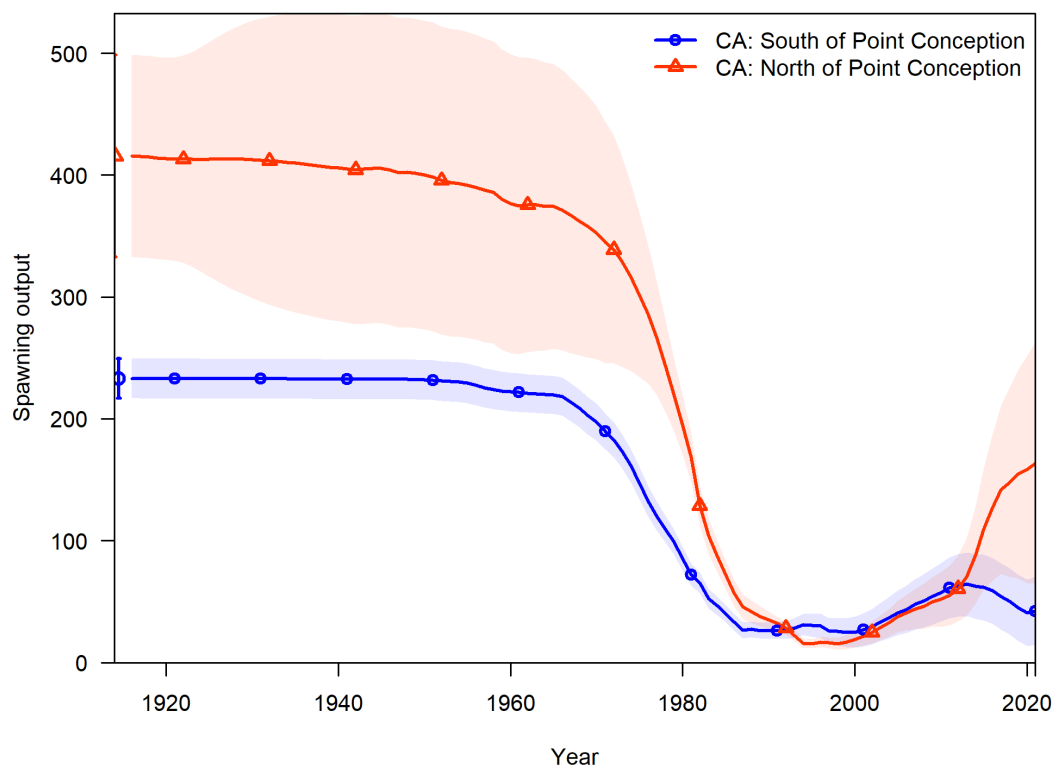
**Figure 49:** Derived quantities from the SSS run where fraction unfished was assumed to be 40 percent.



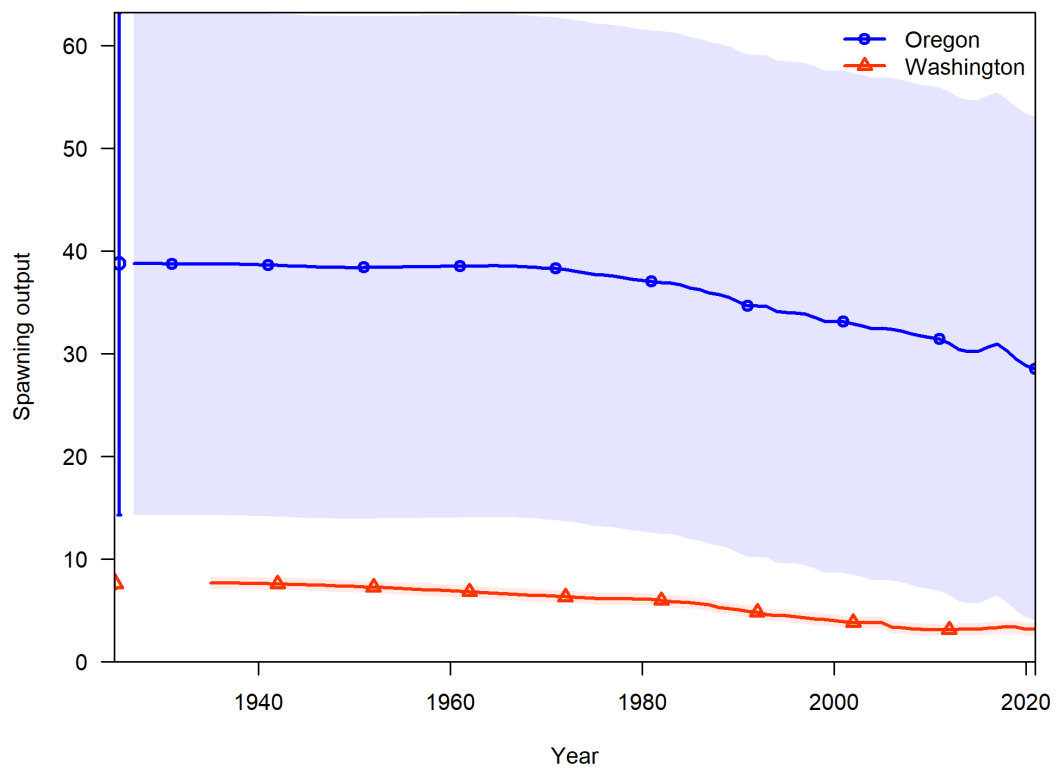
**Figure 50:** Change in the estimate of spawning output when the most recent 5 years of data area removed sequentially.



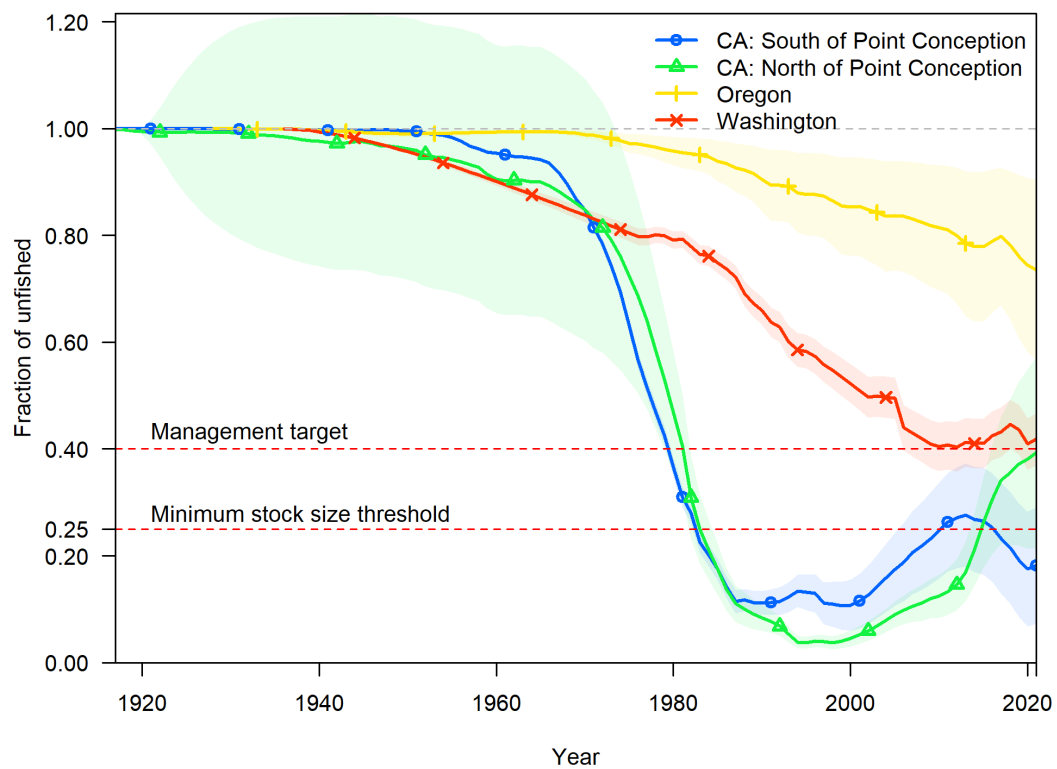
**Figure 51:** Change in the estimate of fraction unfished when the most recent 5 years of data area removed sequentially.



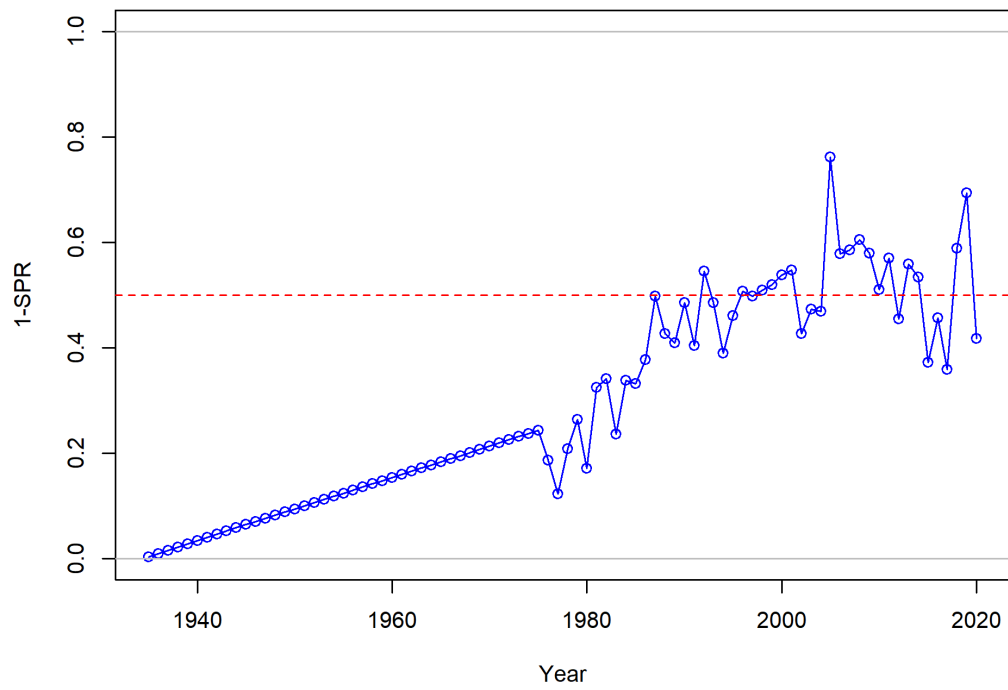
**Figure 52:** Estimated spawning output time series for the California stocks north and south of Point Conception.



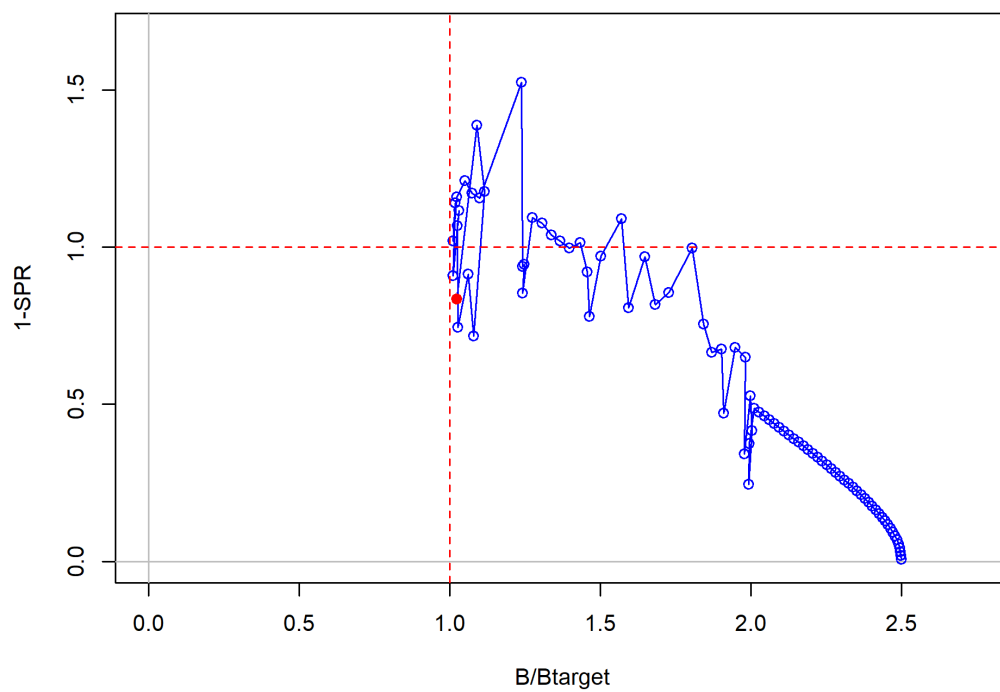
**Figure 53:** Estimated spawning output time series for the stocks off the Oregon and Washington coast.



**Figure 54:** Estimated fraction unfished time series for all West Coast stocks.

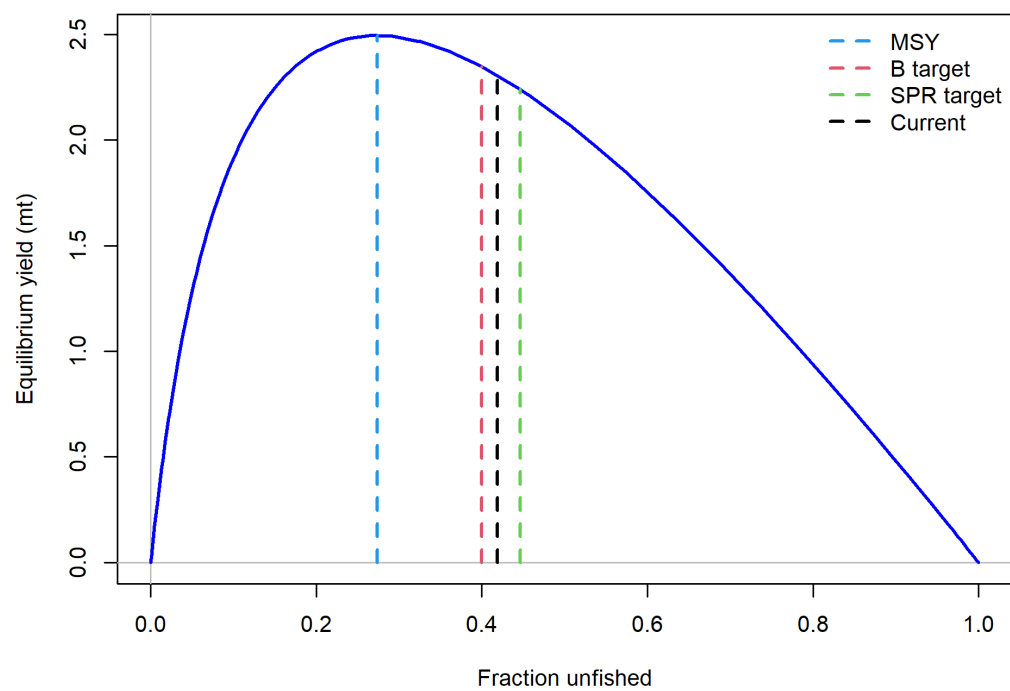


**Figure 55:** Estimated 1 - relative spawning ratio (SPR) by year.



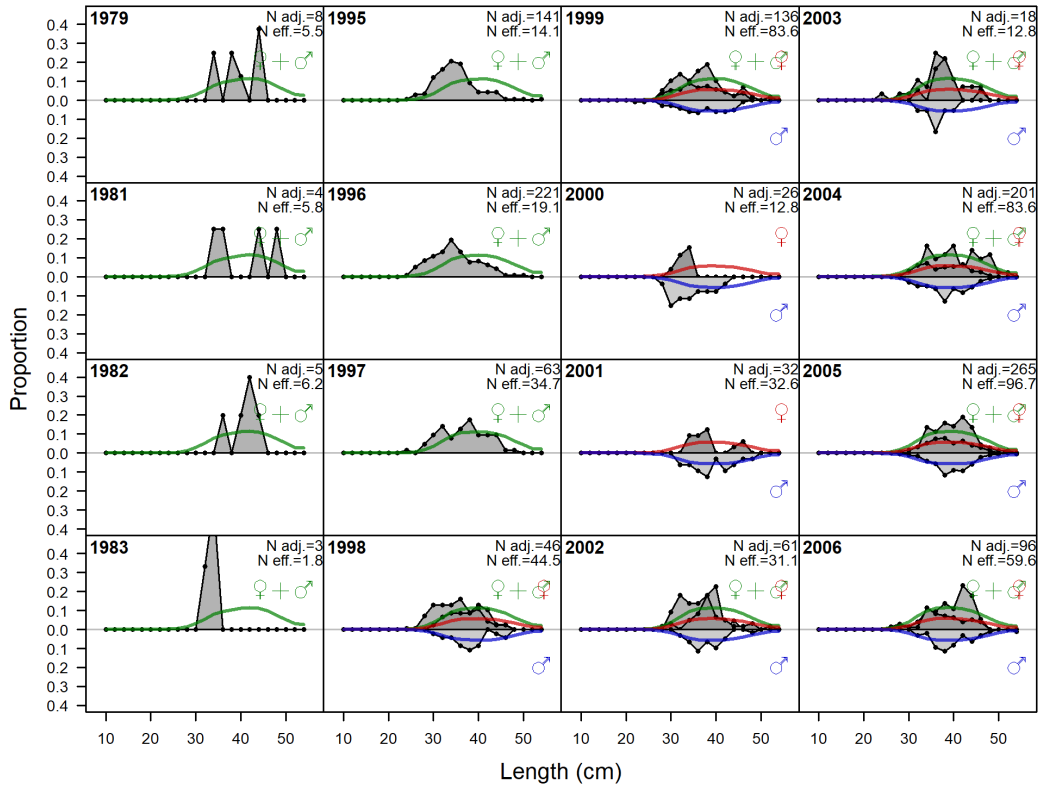
**Figure 56:** Phase plot of the relative biomass (also referred to as fraction unfished) versus the SPR ratio where each point represents the biomass ratio at the start of the year and the relative fishing intensity in that same year.



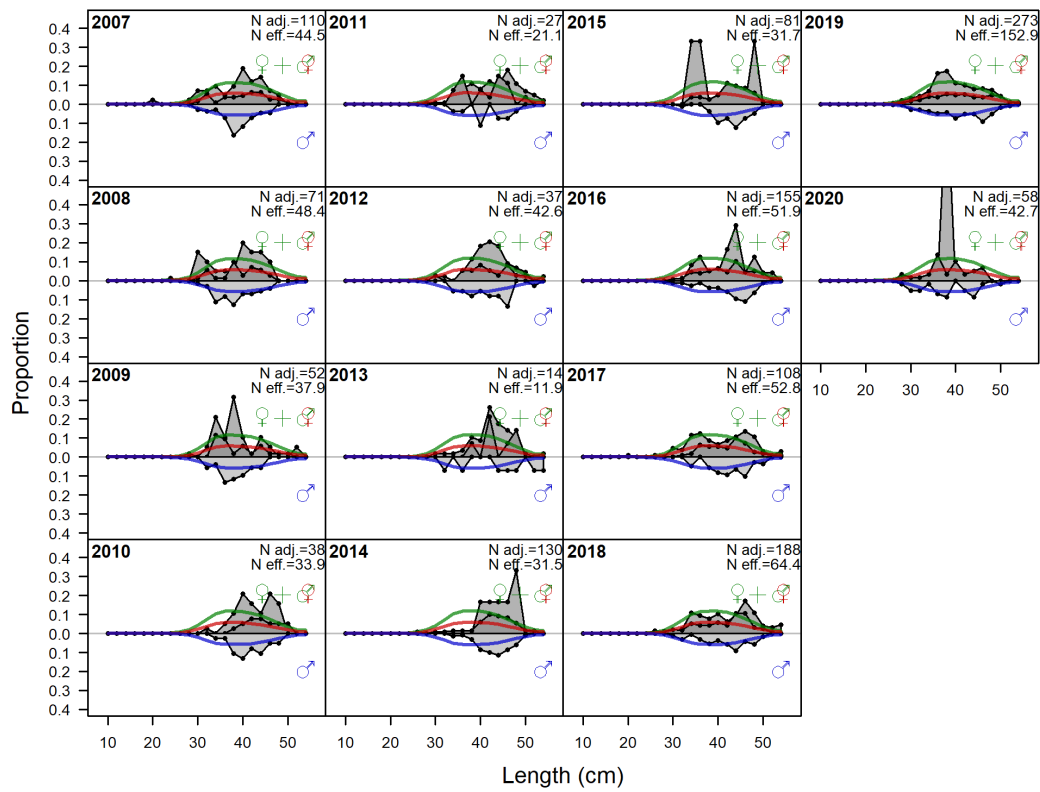


**Figure 57:** Equilibrium yield curve for the base case model. Values are based on the 2020 fishery selectivity and with steepness fixed at 0.72.

## 9 Appendix A. Detailed Fit to Length Composition Data



**Figure 58:** Length comps, whole catch, WA\_Recreational (plot 1 of 2). 'N adj.' is the input sample size after data-weighting adjustment. N eff. is the calculated effective sample size used in the McAllister-Iannelli tuning method.



**Figure 59:** Length comps, whole catch, WA\_Recreational (plot 2 of 2).